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THE AVIONICS LABORATORY PREDICTIVE OPERATIONS AND SUPPORT (ALPOS) COST MODEL . Volume 2.

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This technical report has been reviewed and is approved for publication.

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Recent DOD experience shows that a prime factor in the evaluation of alternative weapon systems for performing a particular mission is Life Cycle Cost (LCC). Since 70% of the system LCC is determined by the end of the conceptual phase, it is important that techniques to predict LCC be available during that phase. Since system definition is not complete enough in this phase to perform detailed analysis using accounting models, the major tool which can be used is parametric estimating models. This report describes a model which relates the available design parameters to LCC via various cost estimating relationships (CERs).

This document is Volume II of the Final Report which describes the mathematical and statistical techniques used to obtain the cost estimating relationships and parametric estimating relationships needed to develop the Avionic Laboratory Predictive Operations and Support (ALPOS) Cost Model. The Air Force Program Monitor was Lt Thomas G. James, Jr., System Evaluation Group (AFAL/AAA-3), Avionic Systems Engineering Branch.

PREFACE

This document is Volume II of the Final Report concerning the development of the Avionics Laboratory Predictive Operations and Support Model (ALPOS) by the Logistics Engineering Group, Westinghouse Integrated Logistics Support Division. This volume presents a discussion of the mathematical and statistical techniques used to obtain the cost-estimating relationships and parametric estimating relationships needed to develop the ALPOS Model, by means of Multiple Regression Analysis. Volume I presents a discussion of the "Design of the Experiment" and the development of the model, in addition to the model's source program listing and documentation. Volume III presents the consolidated data base used to develop the model.

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SECTION I

INTRODUCTION AND OVERVIEW

Much research has been done in Government, Business, and Industry to obtain the capability to predict future occurrences (events). The need for such a capability is demonstrated by the time, money and, in some cases, lives, which can be saved using these predictions. For instance, studies are made to predict such things as the number of red blood cells in a blood sample based upon the packed-cell volume of the blood, the average grade a student makes on a standardized test based upon their I.Q.s, the death rate of males being exposed to the environmental conditions of a coal mine for over a 10 year period, the cost of a piece of avionics equipment based upon its physical characteristics, etc.

Past experience is usually the only means of predicting the future. The approaches to prediction can take forms ranging from hard objective evidence (which is rarely the case) to pure speculation. As an example of hard objective evidence, if you drop an apple from a cliff, most people will agree with the prediction that the apple will hit the ground. Accounting type models are useful estimating tools but require a large amount of detailed information. Such things as estimating the cost of a piece of equipment early in the conceptual/preliminary design phase, however, does not usually have the luxury of hard objective evidence or such detailed information on which to base decisions. Other approaches to prediction, such as the subjective approach, relies on the opinions of qualified experts in the field of study. And then of course, there is the "crystal ball" approach.

Once an estimate is made, however, there is an obvious question, "How accurate is the estimate?" Because of certain constraints (such as time, money and scope of the study) some approaches to estimation are the only ones possible, but there is a major drawback in that the merits of future predictions usually cannot be quantified. Mathematicians and statisticians have developed (and are still developing) many techniques for estimating purposes with particular

emphasis on quantifying the reliability of estimates. A major area of statistics that has been used for over a century is that of Regression Analysis. This is the approach to estimation we take.

Inherent in the interpretation of the words prediction or estimate is the term uncertainty. It would be nice to make "exact" predictions, but this is rarely the case when dealing with a mass of statistical data. Thus, statisticians do not profess to estimate exactly, but that their predictions are "on the average" reasonably close. The basic concept of Regression Analysis is then to estimate the average value of a given variable (called the dependent variable) in terms of the known values of one or more other variables (called independent variables). Regression Analysis expresses the relationships of these variables by determining the form of a mathematical equation connecting them. In other words, there are three major questions that are asked in Regression Analysis:

- (1) Is there a relationship between the dependent and the independent variables?:
- (2) If there is a relationship, how can it be "best" expressed in the form of a mathematical equation?; and
- (3) What statistics, plots, techniques, etc., can be used to verify the accuracy of the equation obtained?

For instance, if a study is made to estimate the average weight of a female in a given university based upon her height, the procedure would be to select a "representative sample" of the females in the university, record both their heights and weights (the data) and try to fit the "best" mathematical relationship that connects weight to height. In many cases, as in estimating equipment costs, one independent variable does not provide enough information to accurately predict the dependent variable. Considering additional independent variables can, in most cases, lead to more accurate estimates, since more information should lead to better predictions.

The purpose of this study is to estimate the Operations and Maintenance (O&M) cost of avionics equipment, based upon the physical

characteristics of the equipment in addition to any current information available (such as the type of aircraft in which the equipment is used and the equipment's avionics area) early in the conceptual/preliminary design phase. The tool used to estimate avionics O&M costs is a computer model developed by the Logistics Engineering Section of Westinghouse for the Air Force Avionics Laboratory (AFAL), and is called the Avionics Laboratory Predictive Operations and Support (ALPOS) Model. The ALPOS model is highly dependent on the six estimating relationships obtained for the logistics, support and cost parameters: Maintenance Manhours per Operating Hour (MMH/OH); Mean Time Between Failure (MTBF); Mean Time Between Maintenance Actions (MTBMA); Logistic Support Costs per Operating Hour (LSC/OH); Training Cost per Operating Hour (TRAIN/OH); and the fraction Not Repairable this Station (NRTS). The approach was to collect data consisting of 21 independent variables covering a wide spectrum of avionics equipment, develop Cost-Estimating Relationships (CERs), i.e. Regression equations where the dependent variable is cost (LSC/OH, TRAIN/OH), and Parametric Estimating Relationships (PERs), Regression equations where the dependent variable is a parameter which drives cost (MTBF, MTBMA, MMH/OH, NRTS) by means of Multiple Regression Analyses. Other parameters which drive O&M cost, such as spares cost and support equipment cost, are not estimated using regressions, since there are many other subjective variables affecting these parameters that cannot easily be quantified. The relationships obtained for MTBMA and NRTS, however, are used in conjunction with an Expected Back Order (EBO) criteria to estimate the quantity of spares and hence spares cost. The interested reader is referred to Vol.I of this report for a look at "The Design of the Experiment" and the development of the ALPOS model, in addition to the approaches used to estimate spares costs and support equipment costs. This volume is mainly devoted to the Multiple Regression Analysis techniques used to obtain the estimating relationships for MMH/OH, MTBF, MTBMA, LSC/OH, TRAIN/OH and NRTS.

Since the six parameters considered are major drivers of Operations and Maintenance Cost, much emphasis has been placed on finding the most up to date approach to the subject of Regression Analysis.

The major reference noted throughout this report is a book written in 1971 by C. Daniel and F. S. Wood entitled Fitting Equations to Data [1], out of which evolved a most powerful computer program called "The Linear Least-Squares Curve-Fitting Program" (LLSCFP). As will be seen, the sophistication of the approach and techniques used in [1] is far beyond that of any standard statistics books and many advanced textbooks on Regression Analysis. The innovative use of "interior" statistics, "Indicator" variables and computerized plots are extremely helpful in leading a qualified statistician in the direction of obtaining the "best" estimating relationships that can be obtained from a given set of multifactor data.

The proposals presented in [1] have been successfully discussed in seminars at many distinguished worldwide universities as well as the Bell Telephone Laboratories and the National Cancer Institute. The LLSCFP has also been the most sought after program in both the SHARE (IBM) and VIM (CDC) libraries of computer programs, and has also been converted to run in East Germany and Russia. These techniques have been applied in a wide range of areas including studies by government agencies of variables for pollution control, searches for influential variables which cause cancer, studies to estimate hospital costs, studies in the conservation of energy and the evaluation of moon rocks at the Johnson Space Center. In addition a Bureau of Labor Statistics study has shown that the coefficients estimated by the LLSCFP are accurate to 15 digits. It is felt then that the proposals and techniques presented in [1] are the "state of the art" in Regression Analysis.

A word of caution however, is in order, in that the LLSCFP is not idiot proof and the cost analyst must remember that Regression Analysis is highly dependent on the "goodness" of the data and maybe to a greater extent on the assumed functional form of the equation. For if the assumed functional form is incorrect then the statistics will be misleading, giving the wrong values of the coefficients to be estimated, making uninfluential variables seem influential and possibly even dropping the most influential variables. Many examples

Fitting Equations to Data, Computer Analysis of Multifactor Data for Scientist and Engineers, C. Daniel and F. S. Wood with the assistance of J. W. Gorman, Wiley, (1971).

in Regression Analysis have an assured form of the estimating relationship based on a previous study or on technical knowledge of the process studied. However, there has been no previous study devoted to developing CERs for avionics equipment in as many as 21 independent variables, nor is enough known about avionics equipment that will lend to technical knowledge of the correct functional form. We must not stop here, but, should simultaneously consider all variables which are "assumed" to have an influential effect on the dependent variable, and let statistics and techniques lead the analyst in the direction of obtaining the equations that yield the best possible predictions. Many functional forms can be quite complicated, and for a given range of interest, transformations (such as, the square, the natural logarithm, the exponential, the square root, etc.) are often used to estimate these cases. To assist in obtaining the "best" possible equations, three forms or transformations of the independent variables (namely the variable, its square and its natural logarithm) and two forms of the dependent variable (the dependent variable and its natural logarithm) are used in this report.

It is to be emphasized that the independent variables are not considered one at a time or in pairs or any other grouping, but that they have all been considered simultaneously to determine their compound effect on the parameters to be estimated. It can be easily shown that a dependent variable can be highly correlated to one variable and no apparent correlation exists between another variable, but the compound effect of both variables (or many variables) has a significant effect on the dependent variable. Hence the practice of using scatter diagrams of the dependent variable versus each independent variable should not be used in determining the form of the equation when multiple variables are considered.

Thus, in this study such complicated functional forms as:

$$y = b_0 + b_1 x_1 + b_2 x_2^2 + b_3 \ln x_3 + \dots$$

and

$$y = b_0 e^{b_1 x_1} e^{b_2 x_2^2} x_3^{b_3} \dots$$

are considered as means of estimating advanced equipment costs. Here y stands for the dependent variable, x_i 's the independent variables, b_i 's are the constants, and e the exponential function.

To assist in verifying the accuracy of the equations obtained, there are over thirty statistics, five types of plots, several techniques and different tabular arrangements of the data that are available in the computer printouts of the LLSCFP. This document includes a brief discussion of the concepts of Regression Analysis including the statistics, plots and techniques used to estimate advanced equipment costs. Also given, as an example, are the procedures and approaches utilized to obtain the parametric estimating relationship for the support parameter Mean Time Between Maintenance Actions (MTBMA).

SECTION II

THE METHOD OF LEAST-SQUARES

The form of the equations considered throughout this report can be written (or transformed) into the linear equation in $(\kappa + 1)$ - unknowns

$$y = \beta_0 + \beta_1 x_1 + \dots + \beta_{\nu} x_{\nu}$$
 (1)

where y is the dependent variable, x_1, \ldots, x_{κ} and the κ - independent variables, β_0 (the constant) and κ - coefficient $\beta_1, \ldots, \beta_{\kappa}$ make up the unknown (κ + 1) population parameters. Also it is assumed that there are N observations (pieces of equipment) in the sample indexed by j. Thus y_j represents the (observed) jth observation of the dependent variable and x_{ij} the jth observation of the ith independent variable. Regression Analysis requires that the analyst find statistics $b_0, b_1, \ldots, b_{\kappa}$ which "best" approximates the unknown (κ + 1) population parameters (where we have taken a sample from the population of all avionics equipment), and whose fitted equation

$$Y = b_0 + b_1 x_1 + \dots + b_K x_K$$
 (2)

gives the "best" possible prediction. The method most widely used by statisticians to accomplish this is called the method of least-squares, which says:

"Find the values of the constants in the assumed equation that minimize the sum of the squared deviations of the observed values from those estimated by the equation."

In other words, minimize $Q = \sum_{j=1}^{N} (y_j - Y_j)^2$, where Y_j is the estimate of the jth observation of the dependent variable obtained by (2). Once the estimates b_0 , b_1 ,..., b_k are found, substituting the values of the independent variables in (2) yields the estimate of the the dependent variable Y. We thus find ourselves in an area of statistics called "Inductive Statistics" which uses the concepts of "Statistical Inference" to make generalizations (or estimates) of population parameters based upon a given sample of the population,

and to quantify the reliability of the estimates obtained. In order to make these generalizations, however, the data must satisfy certain assumptions.

SECTION III

ASSUMPTIONS OF THE METHOD OF LEAST-SQUARES

There are four major assumptions which the data must satisfy in order to use the techniques of least-squares estimation. They are:

- Al. The data is "good" data.
- A2. The correct form of the equation has been chosen. e.g., $y = \beta_0 + \beta_1 x_1 + ... + \beta_{\kappa} x_{\kappa}$
- A3. The independent variables are constant, non-random variables, measured without error.
- A4. All error is in the observations of the dependent variable y_i , i.e.

 $y_j = \beta_0 + \beta_1 x_1 + \ldots + \beta_K x_K + e_j$ where e_j represents random error. Morever, the e_j are normally distributed independent random variables with mean zero and constant, though unknown, variance $\sigma^2(y)$. If all the above four assumptions hold or "approximately" hold, then the least-squares approach will give the best estimates of the coefficients in the relationships. Past experiences, however, indicate that slight departure from the assumptions of normality and equal variances has little effect on the results.

Since these assumptions are the basis for the method of leastsquares estimation and hence Regression Analysis, much emphasis must be placed in determining how close the data fits the assumptions.

SECTION IV

ONE INDEPENDENT VARIABLE

Very often in practice a relationship connecting two variables (one independent and one dependent) is desired. The equation most widely used is the linear equation (in two unknowns β_0 and $\beta_1)$,

$$y = \beta_0 + \beta_1 x_1 .$$

If all pairs of values of x_1 and y, when plotted in a scatter diagram on ordinary graph paper, fall on or near a straight line, equation (3) is the correct form of the relationship to be used. According to the least-squares cirteria, we must use the data to calculate statistics b_0 and b_1 which estimate the parameters β_0 and β_1 , and whose fitted equation can be expressed by

$$y = b_0 + b_1 x_1 (4)$$

In addition, the statistics \mathfrak{b}_0 and \mathfrak{b}_1 must be chosen so as to minimize Q where:

$$Q = \sum_{j=1}^{N} (y_j - Y_j)^2 = \sum_{j=1}^{N} (y_j - b_0 - b_1 x_{1j})^2 .$$
 (5)

By the techniques of differential calculus, the way to find b_0 and b_1 which minimize Q is to take partial derivatives of Q with respect to both b_0 and b_1 , set the results equal to 0 and solve the equations for b_0 and b_1 . Thus, taking partial derivatives we obtain

$$\sum_{j=1}^{N} (y_{j} - b_{0} - b_{1}x_{1j}) = 0$$

$$\sum_{j=1}^{N} (y_{j} - b_{0} - b_{1}x_{1j}) (x_{1j}) = 0 .$$
(6)

and

Solving the first equation of (4) for b_0 and substituting the results into the second equation yields the following linear least-squares estimates:

 $b_0 = \overline{y} - b_1 \overline{x}_1$

and

$$b_{1} = \frac{\sum_{j=1}^{N} (x_{1j} - \bar{x}_{1}) (y_{j} - \bar{y})}{\sum_{j=1}^{N} (x_{1j} - \bar{x}_{1})^{2}},$$
(7)

where \bar{y} and \bar{x}_1 represent the arithmetic mean of the dependent and independent variables respectively.

Scatter diagrams, however, do not always give an indication of an linear relationship, but show some evidence of curvature. For instance, the graph might indicate that the form of the equation is a parabola, i.e.,

$$y = \beta_0 + \beta_1 x_1 + \beta_1 x_1^2$$
 (8)

If a linear equation is fitted to data that is better represented by a curvilinear equation, then assumption A2 is volidated and the results can be spurious.

The use of a special type of graph paper (for which either one or both of the variables are calibrated logarithmically) called semilog or log-log graph paper, is also helpful in choosing other forms of the relationship.

If in a scatter diagram plotted on semi-log paper, the observations fall near a straight line, then the exponential curve,

$$y = \beta_0 e^{\beta_1 x_1} , \qquad (9)$$

is the appropriate choice. If a straight line is obtained on loglog paper, the geometric curve,

$$y = \beta_0 x_1^{\beta_1} , \qquad (10)$$

is appropriate. Taking the natural logarithm of both sides of (9) and (10) yields

$$lny = ln \beta_0 + \beta_1 x_1$$
(11)

 $1ny = 1n \beta_0 + \beta_1 1nx_1$

respectively. Using the simple transformation $y' = \ln y$, $x' = \ln x_1$ and $\beta_0' = \ln \beta_0$, equation (11) is transformed into

$$y' = \beta_0' + \beta_1 x_1$$

and

and

$$y' = \beta_0' + \beta_1 x_1'$$
, (12)

which are both similar in form to equation (3), and whose coefficients can be estimated by (5). Thus, any functional forms (e.g. (8) or (9)) which can be linearized by simple transformations fall under the realm of least-square estimation techniques. Many other plots of linearizable equations (see [1] and [2]) are also useful in finding the correct functional relationship to be used when only one independent variable is considered to be influential.

^{2 &}quot;Fitting Curves to Data," A. E. Horel, Chemical Business Handbook (edited by J. H. Perry), McGraw-Hill, 1954, Section 20, pp. 55-77.

SECTION V

MULTIPLE REGRESSION ANALYSIS

When two or more independent variables are considered in a regression exercise, scatter diagrams and other graphical methods are often useless when trying to determine the form of the assumed equation. For instance, if the two independent variables \mathbf{x}_1 and \mathbf{x}_2 are considered, an \mathbf{x}_1 - y scatter diagram might indicate a high linear relationship whereas the \mathbf{x}_2 - y and \mathbf{x}_1 - \mathbf{x}_2 scatter diagram may show no apparent correlation, even though the true form after equation is

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$
.

Therefore, graphical techniques are not considered as an alternative to finding the correct form of the equation to be tested (as required by assumption A2). If the correct form is not known, the analyst should try several forms of the equation, and let the statistics verify the correct form.

THE "GLOBAL" STATISTICS

As stated previously, in addition to estimating the coefficients, a means is needed to determine how "good" these estimates are. The statistics used to verify the "goodness of fit" of the relationships will be briefly defined with some general comments. The capital letters correspond to the respective names of these statistics as listed in the computer printouts of the LLSCFP.

We initially begin with a few elementary statistics which are quite helpful. They are the sums, means, maximums, minimums, ranges and standard deviations of the variables (both independent and dependent). With these statistics, the analyst can get a good indication of the distributional properites of the variables.

SUMS OF VARIABLES

The computer LLSCFP lists the sum, $\sum_{j=1}^{N} x_{ij}$, for each independent variable and $\sum_{i=1}^{N} y_{i}$ for the dependent variable.

MEANS OF VARIABLES

The arithmetic mean of each independent and dependent variable denoted by

$$\bar{x}_i = \frac{\sum_{j=1}^{N} x_{ij}}{N}$$

and

$$\overline{y} = \frac{\int_{\Sigma}^{N} y_{j}}{N}$$

are listed under this heading.

ROOT MEAN SQUARES OF VARIABLES

The root mean squares of the variables (also called the standard deviation) is a statistic that can give an indication of the spread or variation of each variable, independent and dependent, in the data and is denoted by

$$S_{x_{i}}^{2} = \frac{\int_{y=1}^{N} (x_{ij} - \bar{x}_{i})^{2}}{N-1}$$

and

$$S_y^2 = \frac{\int_{\Sigma}^{N} (y_j - \overline{y})^2}{N - 1}$$

respectively.

MAX X(I)

The maximum value of the ith independent variable.

MIN X(I)

The minimum value of the ith independent variable.

RANGE X(I)

The range of the ith independent variable, i.e., the maximum value minus the minimum value.

MAX Y

The maximum value of the dependent variable.

MIN Y

The minimum value of the dependent value.

RANGE Y

The range of the dependent variable.

Many times these elementary statistics can give a quick indication that something is wrong with the data (e.g., an impossible maximum value). Recall Assumption Al specifies that the data is "good" data. The simple statistics can be helpful in pinpointing which particular variables are causing such things as outliers (impossible values) to develop in the results.

COEFFICIENT B(I)

The least-squares estimates of a multiple regression equation are obtained in a similar fashion as those estimated for the linear-equation in one independent variable. The partial derivative of Q with respect to the constant b_0 is taken and set to zero and solved for b_0 . This result is then substituted into Q where partial derivatives of Q with respect to each of the κ - coefficients are taken and set to zero, thereby yielding a κ × κ system of equation in κ - unknowns. This system of equations is then solved by the methods of determinants to determine the desired estimates of the coefficients. The linear-least squares estimates are:

$$b_0 = \overline{y} - \sum_{i=1}^{\Sigma} b_i \overline{x}_i$$

and

$$b_{i} = \sum_{j=1}^{N} y_{j} \{c_{i1} (x_{1j} - \bar{x}_{1}) + c_{i2} (x_{2j} - \bar{x}_{2}) + ... + c_{i\kappa} (x_{\kappa j} - \bar{x}_{\kappa})\}$$

where c_{ij} is the element of the inverse matrix (obtained by the method of determinants) belonging to the ith row of the jth column. The LLSCFP

lists the constant \mathbf{b}_0 and a coefficient \mathbf{b}_i for each independent variable. The coefficient \mathbf{b}_i can be used to determine the influence of variable \mathbf{x}_i on the fitted equation.

RESIDUAL

The residual is defined as the difference between the observed value of the dependent value and the value estimated by the prediction equation, i.e., $y_j - Y_j$. Note that this simple statistic is the basis for least-squares analysis. Once the prediction equation is obtained, the residuals show how well the equation estimates the dependent variable for each observation (piece of equipment) in the data base.

FITTED Y

The statistic Y_j is called the fitted y-value of the jth observation of the dependent variable. This is the value of the dependent variable estimated by the prediction equation for each observation in the data base.

TOTAL SUM OF SQUARES

The total sum of squares in an initial step is trying to get a grip on the error of prediction. It is a measure of the total variation in the dependent variable. It is proportional to S_y^2 , the variance of y and is defined:

TOTAL SUM OF SQUARES =
$$\sum_{j=1}^{N} (y_j - \bar{y})^2$$

The total sum of the squares can be partitioned into two useful sums, the sum of the squares due to the fitted equation and the residual sum of the squares, i.e.

$$\sum_{j=1}^{N} (y_{j} - \bar{y})^{2} = \sum_{j=1}^{N} (Y_{j} - \bar{y})^{2} + \sum_{j=1}^{N} (y_{i} - Y_{j})^{2}$$

SUM OF SQUARES DUE TO THE FITTED EQUATION

The sum of the squares due to the fitted equation,

SSFE =
$$\sum_{j=1}^{N} (Y_j - \overline{y})^2$$

is that part of the total variation in the dependent variable y that can be attributed to the fitted equation. A large SSFE indicates that the equation used is accounting for most of the variation.

RESIDUAL SUM OF SQUARES

The residual sum of the squares is that part of the total variation that cannot be attributed to the fitted equation (such things as experimental error, chance, function bias, i.e., having the wrong form of the equation, or other biases), and is defined as

RESIDUAL SUM OF SQUARES =
$$\sum_{j=1}^{N} (y_j - Y_j)^2$$

RESIDUAL MEAN SQUARE

Recall Assumption A4 states that the y-observations have the same constant though unknown variance $\sigma^2(y)$, and hence, we need a means of estimating this variance of prediction. One estimate of $\sigma^2(y)$ called the residual mean square (or variance), denoted by $S^2(y)$, is defined as

RESIDUAL MEAN SQUARE =
$$\frac{\sum_{j=1}^{N} (y_j - Y_j)^2}{N - \kappa - 1}$$

where N - \times - 1 is the RESIDUAL DEGREES OF FREEDOM. The degrees of freedom of a statistic is the number of independent bits of observations minus the number of parameters estimated in the calculation of the statistic.

RESIDUAL ROOT MEAN SQUARE

The square root of the residual mean square is called the RESIDUAL ROOT MEAN SQUARE (or standard deviation). This can be interpreted in a similar way as the standard deviation of the prediction equation. The residual root mean square is also called the Standard Error of Estimate.

STANDARD ERROR OF THE COEFFICIENT

Since each of the coefficients obtained by the method of least-squares are only estimates, the accuracy and importance of each estimate must be shown. From the equation which calculates b_i , it can be shown that the variance of the b_i , denoted by $S^2(b_i)$ can be written in the form

$$S^{2}(b_{i}) = S^{2}(y) C_{ii}$$

where $C_{i\,i}$ is the ith diagonal element of the inverse matrix. The square root of the variance of b_i , $S(b_i)$, is called the standard error of the coefficient and is calculated by the formula

S.E.COEF. =
$$S(y) (C_{ii})^{\frac{1}{2}}$$
.

The standard error of the coefficient gives an indication of the "inherent" precision of the coefficient estimated. If a coefficient is large, say 1,000 with a standard error of .1, we could say that the coefficient is estimated with great precision. However, a coefficient of say .2 with a standard error of .1 is obviously not as precise an estimate. Therefore, a means of determining the relative accuracy of each coefficient is needed.

T-VALUE

The statistic used to measure the relative accuracy of each of the coefficients is called the T-VALUE (denoted by t_i) and is defined by

T-VALUE =
$$\frac{b_i}{S(b_i)} = \frac{COEF B(I)}{S.E.COEF.}$$

The LLSCFP lists a T-VALUE for each coefficient. The larger the coefficients t_i - value (as compared with the other t-values), the more chance variable \mathbf{x}_i will be in the final fitted equation.

RELATIVE INFLUENCE OF EACH INDEPENDENT VARIABLE

A statistic called the relative influence of x_i , describes the fraction of the total change in Y that can be accounted for by the accompanying total change in the ith independent variable and is defined as

REL.INF.X(I) =
$$\frac{b_i w_i}{w_y}$$
,

where $b_i = COEF B(I)$ is the coefficient, $w_i = RANGE X(I)$ is the range of the independent variable and w_y is the range of the dependent variable.

MULTIPLE CORRELATION COEFFICIENT SQUARED

The total sum of the squares, sum of squares due to the fitted equation and residual sum of the squares will have different values for different equations. The most widely used statistic which gives a relative measure of the "goodness of fit" of the assumed equation is the multiple correlation coefficient squared (denoted by R_{γ}^2) where

$$MULT.CORREL.COEF.SQUARED = \frac{SSFE}{TOTAL SUM OF SQUARES}$$

The multiple correlation coefficient squared (also called the coefficient of determination) is defined as that fraction of the total sum of the squares that can be attributed to the fitted equation. If $R_y^2 = 1$ we are fortunate to have a "perfect" fit and if $R_y^2 = 0$ the fitted equation does not fit the data at all. In most cases the multiple correlation coefficient squared will fall between the values of 0 and 1, and here interpretation is necessary.

For instance, if a straight line is fitted to a pair of data points when the correct form of the equation should be exponential (i.e., a logged dependent variable), a low R_y^2 does not indicate that there is "no" relationship between the data, but that the relationship used does not adequately represent the data. Thus the multiple correlation coefficient squared measures the degree of the relationship relative to the equation used. Scmetimes, however, a large R_y^2 such as .90 can occur when the wrong form of the equation is chosen (examples will be given). Therefore a specific value of R_y^2 that indicates a "good" fit is not given. Although the multiple correlation coefficient squared is the most widely used statistic that measures the accuracy of the relationships, it means very little to this analyst when considered alone. It should be considered in conjunction with all other statistics, graphical representations of goodness of fit, and techniques used to determine the stability of the equations.

F-VALUE

A statistic used to determine the "significance" of R_y^2 is the statistic called the F-VALUE. The F-VALUE is usually used to judge the equivalence of two independent estimates of variance. It can be easily be shown that

$$R_{y}^{2} = \frac{\int_{\Sigma}^{N} (Y_{j} - \bar{y})^{2}}{\int_{j=1}^{N} (Y_{j} - \bar{y})^{2}} = 1 - \frac{\int_{\Sigma}^{N} (y_{j} - Y_{j})^{2}}{\int_{j=1}^{N} (y_{j} - \bar{y})^{2}}$$

and therefore indicates that the two estimates of variance

$$\sum_{j=1}^{N} (\gamma_j - \bar{y})^2$$

and

$$\frac{\sum_{j=1}^{N} (y_j - Y_j)^2}{N - \kappa - 1}$$

should be helpful in determining the significance of R $_y^2$, where κ and N - κ - 1 are degrees of freedom respectively.

Hence, the F-VALUE is defined by the variance ratio,

$$F-VALUE = \frac{\frac{\sum_{j=1}^{N} (Y_{j} - \bar{y})^{2}}{\kappa}}{\frac{\sum_{j=1}^{N} (Y_{j} - Y_{j})^{2}}{N - \kappa - 1}} = \left(\frac{N - \kappa - 1}{\kappa}\right) \frac{\sum_{j=1}^{N} (Y_{j} - \bar{y})^{2}}{\sum_{j=1}^{N} (Y_{j} - Y_{j})^{2}}$$

which indicates that the larger SSFE is with respect to RESIDUAL SUM OF SQUARES, the larger the F-VALUE. An equivalent form of the F-VALUE is

F-VALUE =
$$\left(\frac{N - \kappa - 1}{\kappa}\right) = \frac{R_y^2}{(1 - R_y^2)}$$

If R_y^2 is close to 1 then $1 - R_y^2$ is close to 0 and the F-VALUE is large. Of course the question must be answered, "How large is large enough?" This question can be answered by performing the statistical hypothesis test called the F-Test.

In performing a statistical test, a level of significance is usually assumed, denoted by the Greek letter α . The level of significance is the amount of risk that the analyst is willing to take in rejecting a true hypothesis. The values of α which are usually assumed are .05 (The test would be called "probably significant" and further experimentation may be in order.) and .01 (The test is called "highly significant."). Throughout this report α = .01 is the assumed level of significance.

The F-Test compares the F-VALUE with values from an F-Table (in most standard statistics books) with κ and N - κ - 1 degrees of freedom, to give a joint test of the hypothesis that:

"all the coefficients of the fitted equation are 0" (indicating a bad fit).

against the alternative that

"the equation as a whole produces a significant reduction in the total sum of square" (indicating a good fit).

Suppose for a given fit the F-VALUE = 10.8 where there are N = 15 observations and κ = 3 independent variables (α = .01). From an F-table, a value (based on κ = 3 numerator degrees of freedom and N - κ - 1 = 11 denominator degrees) of 6.22 can be extracted. Since 10.8 is greater than 6.22 the fit is considered significant. The larger the F-VALUE is with respect to its associated tabular value, the more significant the fit.

SIMPLE "LINEAR" CORRELATION COEFFICIENT

A statistic (similar to the multiple correlation coefficient squared) which gives an indication of how any two variables are linearly related is the simple "linear" correlation coefficient

$$r_{12} = \begin{pmatrix} \frac{N}{\Sigma} & (x_{1j} - \bar{x}_{1}) & (x_{2j} - \bar{x}_{2}) \\ \frac{j=1}{N} & \frac{N}{\Sigma} & (x_{1j} - \bar{x}_{1})^{2} & \sum_{j=1}^{N} & (x_{2j} - \bar{x}_{2})^{2} \\ j=1 & j=1 \end{pmatrix}$$

which is a measure of the <u>linear</u> interdependence between any two variables. The simple linear correlation coefficient is a number between -l and +l, where $\mathbf{r}_{12}=\mathbf{l}$ indicates positive linear correlation, $\mathbf{r}_{12}=-\mathbf{l}$ indicates a negative linear correlation, and $\mathbf{r}_{12}=0$ indicates that there is no linear relationship between the two variables. Again, note that $\mathbf{r}_{12}=0$ does not indicate that there is "no" correlation between the variables, but that no <u>linear</u> correlation exists.

R(I) SQRD

There are obviously many problems in which more than two independent variables are assumed to be influential, and it is very informative to see how all the independent variables are linearly related to each other. A statistic (which is a generalization of r_{12}) that gives an indication of how the ith independent variable is linearly related to all the other $(\kappa - 1)$ - independent variables, is the squared multiple "linear" correlation coefficient R_1^2 . Here with variable x_1 as the dependent variable and the remaining $(\kappa - 1)$ independent variables, we fit a simple linear equation and calculate the multiple correlation coefficient squared which is R_1^2 . Thus R_1^2 measures the degree of linear dependence of x_1 or the other x_1 , $i' \neq i$, where $R_1^2 = 1$ indicates that a strong linear relationship between x_1 and the other x_1 ,'s. If only two independent variables are considered then

$$R_1^2 = r_{12}^2 = r_{21}^2 = R_2^2$$

It can be shown that $\mathbf{C}_{\mbox{ii}}$ (the element of the inverse matrix in the ith row and jth column) may be written as

$$c_{ii} = \frac{1}{\sum_{j=1}^{N} (x_{ij} - \bar{x}_{i})^{2} (1 - R_{i}^{2})}$$

This leads to a relationship between the standard error of the coefficient b_i , $s(b_i)$ (S.E.COEF) and the squared multiple linear correlation coefficient, R_i^2 (R(I)SQUARED), namely

$$s(b_{i}) = \frac{s(y)}{\sqrt{\sum_{j=1}^{N} (x_{ij} - \bar{x}_{i})^{2} (1 - R_{i}^{2})}}$$

With the form of the equation chosen, s(y) and $\sum_{j=1}^{N} (x_{ij} - \bar{x}_i)^2$ are constant and determined by the data, and R_i^2 therefore determines the size of $s(b_i)$. If R_i^2 is close to 1, then $1 - R_i^2$ is close to 0 which increases the size of the standard error of the coefficient, $s(b_i)$. A large $s(b_i)$ will result in a small t_i -value, thus possibly dropping x_i from the set of influential variables to be used in the final prediction equations. Some independent variables are simply uninfluential in their effect on the dependent variable and will be dropped (by using a technique called the Cp-search technique), while other independent variables may be so highly correlated (linearly) to the remaining independent variables that their effects on the dependent variable can be explained by the remaining variables.

THE D-STATISTIC

Many times the statistics and plots will indicate some form of lack of fit of the assumed equation. This lack of fit may be caused by having the wrong form of the equation, sometimes called function bias. For instance, we may be fitting a straight line to a set of paired data when actually a parabola is the correct form, i.e., a squared term is needed in the equation. If the statistics and plots indicate that curvature in a particular variable \mathbf{x}_i is needed, \mathbf{x}_i^2 can be used as an additional independent variable. Usually $(\mathbf{x}_i - \bar{\mathbf{x}}_i)$ and $(\mathbf{x}_i - \bar{\mathbf{x}}_i)^2$ can be used as independent variables to reduce the high correlation between \mathbf{x}_i and \mathbf{x}_i^2 , resulting in an unwanted large \mathbf{R}_i^2 . Sometimes, however, the mean $\bar{\mathbf{x}}$ does not sufficiently reduce the high correlation between a variable and its square. A

statistic d_i (called the d-statistic) due to O. Dykstra [1], which requires that the covariance between $(x_i - \bar{x})$ and $(x_i - d_i)^2$ to be zero, can be used to reduce such high correlations. This reduces to solving

$$\sum_{j=1}^{N} (x_{ij} - \bar{x}_{i}) (x_{ij} - d_{i})^{2} = 0$$

for d,, which yields

$$d_{i} = \frac{\sum_{j=1}^{N} x_{ij}^{2} (x_{ij} - \bar{x})}{\sum_{j=1}^{N} (x_{ij} - \bar{x}_{i})^{2}}$$

In this case we use $(x_i - \bar{x})$ and $(x_i - d_i)^2$ as the independent variables. If \bar{x}_i is used instead of d_i , a large R_i^2 for either the variable or its square may occur, thereby possibly dropping the variable and/or its square as being uninfluential variables, when in fact both may be influential variables. Through this report, the d_i -statistic is used if there is an indication of curvature in the relationships.

THE CP-STATISTIC, $P = \kappa + 1$

In many cases when multiple variables are considered, not enough previous work has been done in the area of study to be sure that all the independent variables are influential, but that possibly a subset collection of the variables fit the data "better" or as "good" as the initial set. If there are $\kappa = 18$ independent variables, then there are $2^{18} = 262,144$ possible combinations of variables whose equation must be compared.

A major innovative statistic due to C. Mallows ([1] and [3]) called the ${\rm C_p}$ -statistic, is used as a measure of "goodness of fit" to compare all the possible $2^{\rm K}$ combinations of equations for the "set" of equations which best fits the data. The ${\rm C_p}$ -statistic represents the "total squared error (random squared error and bias squared error)" and is defined as

[&]quot;Choosing Variables in a Linear Regression: A Graphical Aid,"
C. L. Mallows, presented at the Central Regional Meeting of the
Institute of Mathematical Statistics, Manhattan, Kansas, May 7-9, 1964.

$C_p = \frac{\text{RESIDUAL SUM OF SQUARES}}{\text{RESIDUAL MEAN SQUARE}} - N + 2p$.

A C_p -statistic is calculated for each equation. Note that for every variable dropped, the C_p -statistic can decrease by at most 2 units. For the derivation of C_p see [i] and [3].

PRECISION OF X(I)

This statistic is a new statistic not now in the statistical cannon. A paper to be published in the near future by F. S. Wood will introduce it and, hence will not be discussed here.

THE "INTERIOR" or "LOCAL" STATISTICS

All the previous statistics discussed fall under the heading of "global" statistics in that they are statistics of the entire set of data. The 'global' statistics are helpful in determining how the indepen-... dent variables influence the fitted equations, but they do not describe how the observations (the interior of the data) in multifactor space affect the fit. Four innovative "interior" statistics which have been developed (see[1]) can assist in such things as detecting outliers; indicating observations which may influence the form of the equation (possibly introducing curvature); detecting those observations which have the largest effect on the assumed equation, finding those observations which are taken approximately under the same x;-conditions (called nearby neighbors) and in testing the validity of the "global" statistics. These nearby neighbors are used to obtain a less biased estimate of our error of prediction, $\sigma^2(y)$. The interior statistics defined are weighted by the b; - values so as to reduce the effects of uninfluential factors.

WSS DISTANCE (Weighted Squared Standardized Distance)

An observation whose points are at the extreme ends of the independent variables are usually far from the "centroid" of all the observations of the data. For instance, a piece of equipment may have a much larger (or smaller) weight than all the rest of the data. A statistic called the Weighted Squared Standardized Distance (WSS DISTANCE), defined by

WSS DISTANCE =
$$\sum_{i=1}^{K} \frac{b_i (x_{ij} - \bar{x}_i)}{s(y)}^2$$

is useful in detecting those observations which are far from the centroid of all the observations. These observations (with large WSS DISTANCE) may indicate that outliers are present or possibly that curvature is needed in the equation.

COMPONENT EFFECTS, C

The statistic Cii, where

$$C_{ij} = b_i (x_{ij} - \bar{x}_i)$$
,

may be defined as the component effects of $\mathbf{x_i}$ (the ith independent variable) on $\mathbf{y_j}$ (the fitted value of observation j). A nice tabular arrangement (See 'COMPONENTS EFFECTS" TABLE) of the components effects assists the analyst in determining the influence that each particular observation has on the fitted equation. This is helpful in the analysis of trade studies.

WSSD

In some cases the observations may have nearly the same $\mathbf{x_i}$ - values (for instance similar pieces of equipment) and can be considered as being "close" to each other in multidimensional factor space (nearby neighbors). The statistic

$$WSSD_{jj} = \sum_{i=1}^{K} \frac{b_{i} (x_{ij} - x_{ij})^{2}}{S(y)} = \frac{1}{S^{2}(y)} \sum_{i=1}^{K} (C_{ij} - C_{ij})^{2}$$

measures the (squared) distance in "effect space" between two observations j and j'.

This statistic is also helpful in detection of what is known as "Nested Data." If the analyst determines that his data is nested, additional things must be considered to find the correct form of the equations (See [1]).

CUMULATIVE STANDARD DEVIATION ESTIMATED FROM NEAR NEIGHBORS

Recall that as a "global" estimate of random error of prediction, the residual mean square (variance) or the residual root mean square (standard deviation) was used. Sometimes the inner characteristics of the data may indicate that these are not very good estimates. The statistic WSSD identifies those near neighbors which are used to obtain a less biased running estimate, S_n , of S(y) called the standard deviation estimated from near neighbors where

$$s_n = \frac{.886 \left(\sum_{n=0}^{\infty} \Delta_n d\right)}{n}$$

Here $\Delta_n d$ is the absolute value of the differences of the residuals of the neighboring observations, and the value .886 = 1/1.28 is used since the expected value of the range for pairs of independent observations from a normal distribution is 1.128. If the residual root mean square is close to the successive estimates of cumulative standard deviation S_n , there is then no evidence of lack of fit of the proposed equation.

STATISTICAL PRINTOUTS AND TABLES

The statistics are printed out in an orderly manner which the analyst can use to further evaluate the fit of an equation. A partial print-out of the statistics on the equation obtained for the fit of the dependent variable MAINTENANCE MANHOURS PER OPERATING HOUR, where Y1 = MMH/OH, is shown in Figures 1 to 4. Figure 1 shows many of the global statistics, including the coefficients b_i , $S(b_i)$, t_i and the relative influence of x_i for each independent variable x_i . In addition R_y^2 , the F-VALUE and the residual root mean square are displayed.

OBSERVATIONS ORDERED BY COMPUTER INPUT AND BY RESIDUALS

As shown in Figure 2, under the heading "ORDERED BY COMPUTER INPUT," the residuals are listed in the order in which the observations were given to the computer. The Work Unit Code (WUC) of each piece of

FIGURE 1

IND.VAR(I)

LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

Mark COEF.	2.300 = 2			+ 8(7)×130 + 8(12)0 + 8(12) + 8(14)	### ##################################	17) X21050	# 8 (18) x 16 + 8 (18) x 16 (14) x 140 S 0 + 8 (18) L h x 8 PER OPERATI	# 8(3) + 8(1)x3 + 8(2)x4 + 8(3)x7 + 8(4)x13 + 8(5)x16 + 8(6)x8N + 8(7)x184 + 8(8)x144 + 8(9)x154 + 8(10)x16 + 8(11)x14 + 8(12)x9052 + 1(13)x12053 + 8(14)x14053 + 8(19)x15050 + 8(12)x16054 + 8(17)x21055 + 8(18)LNS + 8(19)LNXB # MHYDH (MAINTENANCE MANHOURS PER OPERATING HOUR)			
2.320-82 2.3	2.320-02 2.3	IND.VAR(I)	NAN	COEF. R(I)	S.E. COEF.	T-VALUE	R(1) SORD	HIN X(I)	HAX X(1)	RANGE X (1)	REL. INF. X(I)
2.600.02 2 2.600.02 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	2.000.002 2 2.000.002 2 2 0 0 1110 0 0 0 110 0 0 0 110 0 0 0		× 3	-4.512210-82	2.320-82	6.1	8.2678	-2.1340-81	5.3260-81	7.4680-81	8.87
2. 250.002 2. 250.002	2.550-02 3.550-02 3.550-02 4.0 5.550-02	. ~	9×	-7.747450-82	2.690-02	5.0	0.4116	-2.3170-61	5.7830-01	8.1000-61	41.9
3.510-04 1.220-32 1.220-33 1.220-34 1.220-	1.070.024 1.020.024 1.020.024 1.020.024 1.020.025 1.020.		**	-6.499060-92	2.530-02	2.6	0.3296	-2.8570-81	6.9430-01	8.1000-01	8.12
1.400-04	1.400-04 4 0.0577 4 1.4000 02 1.4000	•	X 1.3	1.622920-03	3,510-84	4.0	0.7724	9.0	1.0000 02	1,0000 62	8.35
1,20-25 5,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,40-64 6,4	1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-25 1.20-26 1.2	•	X 1 8	-3.346750-84	1.400-04	2.4	0.5777	6.6	1.9400 82	1,0000 92	6.67
6.500-04 7.6 9.9972 -3.3480 01 1.2970 02 1.775	5.500-024 7.8 5.500-02 1.2070 01 1.2070 02 1.2	•	1 6 ×	-6.613690-95	1,220-25	3.6	0.9634		6.7620 P3		1.19
3,400-84	2.400-04		K 1 3 H	4.628710-03	6.610-04	7.0	0.9672				1.75
2.400-04	2.40	•	XIAN	1.872780-03	3.490-94	3.6	0.9117	-6.2700 01			9.41
4,707-24 4,707-24 4,507-24 4,507-20 49 1,107-10 20 1,1	4,707-24 4, 707-	•	×13×	1,200000-03	3,440-04	3.5	0.9817	-1.6980 61			6.26
5.30-04 2.3 6.633 4.930 80 6.10 0 1 6.1	5.30-04 4 2.3 6.6335 4.930 80 5.10 0 71 0.10 0 1	19	×16×	1.543680-83	4.070-04	3.8	P.7162	-3.4330 98			9.34
2,400-79 7,8 0.8466 6,1810 84 2,504.0 97 2,1810 07 2,204.0 97 2,20	2,400-89 7,8 8 8,846 8,1810 84 2,5040 97 2,8040 07 2,204	=	x214	1.249280-03	5,390-04	2.3	6.6835	-4.839D 98			6.17
5.210-85 2.5 8.552 1.8470 88 1.1950 84 1.1950	2, 20 - 70 - 70 - 70 - 70 - 70 - 70 - 70 -	12	0506×	1.701100-08	2.450-09	7.0	9.8466	6.1910 B4			00.0
0.10-25 5.4 7.374 4.0920-82 2.5510 83 2.5510 8	0.10-26 5.4 6.3764 4.0900-n2 2.3610 93 2.3610	13	K10050	-1,292830-45	5.210-86	2.5	P.8362	1.9400 98	1,1950 84	1,1950 84	8.34
7.900-05 4.5 0.4176 91 2.4010 93 2.6100 93 0.400 91 2.4010 93 0.400 91 2.4010 93 0.400 91 2.4010 93 0.400 91 2.4010 93 0.400 91 2.400 92 0.400 91 2.400 92 0.400 91 2.400 92 0.400 91 2.400 92 0.400 92 0.400 93 0	7.900-95 4.5 6.4176 91 2.4100 93 2.4100 93 2.4000 93 2.4	-	K14050	-3,335810-85	6.140-86	3.4	8.5764	4.8920-82			9.19
0.4129 1.7500 03 2.5170 03 7.6730 02 0.6730 02 0.6730 02 0.6730 02 0.6730 03 0.76730 03 0.6730 03 0.76730 03 0.6730	0.4129 1.7550 83 2.5170 83 7.6730 82 2.50.093 2.5170 83 7.6730 82 2.50.093	13	x15050	3,568720-85	7.930-06	4.5	0.4176	5.4760 81			6.22
2,700-63 2.1 0,4433 6,2800 01 1,1610 03 1,7000 03 1,7000 03 1,2000	2,700-63 2.1 0.4433 6.2800 01 1.1610 03 1.5000 03 0.000 03 0.0000	91	x16030	-6.386870-85	3.910-03	2.1	R. 4329	1.7500 83		7,6730 82	9.14
0.0000 0.	1.870-82 4.1 0.9669 3.4810 88 B.8120 88 6.9730 89 6.9730	11	X21050	5,781580-65	2,760-45	2.1	0.4833	6,2880 81			9.14
0-82 1.970-82 2.5 0.9872 1.8230-61 5.1370 88 4.9750 88 63 63 63 63 63 63 63 63 63 63 63 63 63	0.922 1.970-02 2.5 0.9872 1.8230-61 5.1370 68 4.9750 68 63 63 63 63 63 63 63 63 63 63 63 63 63	9.	CXXO	7,478670-82	1.830-02	1.7	0.9669	3.4910 99	9.0120 88	5,6110 68	8.92
0-4N	REPRESENT OF THE REPRES	0.1	LNXIB	-4.981970-82	1.970-92	2.5	8.9672	1.6230-01	5,1570 88	4.9750 88	9.54
- • •	60 00 00 00 00 00 00 00 00 00 00 00 00 0	NO. OF 0836	RVATION	63	20						
	25.5 25.00.00 25.00.00 25.00.00 25.00.00 25.00.00 25.00.00 25.00 2		3390	100000000000000000000000000000000000000							
RE AREO	RE 6.63912463 6.63990842 6.3999633 8.3999533 8.8ED .	F-VALUE	200								
AREO	6.000027 6.0000842 6.0000842 6.00008333	RESIDUAL PO	DT MEAN		8.83915483						
AREO	6.020033 6.020033 AREO .0005	RESIDUAL ME	4N SQUA	.AE	9.8889827						
	6.000000000000000000000000000000000000	RESIDUAL SU	IN OF SC	UARES	9.83989842						
	s 9 6 6 6 .	TOTAL SUM G	F SOUAR	123	6.39298338						
	REGULACO X (1) PRECISION	MULT, CORRE	L. COE	. SOUARED	. 9865						
	ACICIATO (X/X) PARCICIATO										

LINEAR LEAST-SOUARES CURVE PITTING PROGRAM

	MASS DISTANCE	COMPUTER OBS. Y	INPUTATION	RESIDUAL	085V.	088, 4	ITTED Y	ORDERE	SEG
	10.		6.631	8.933	19	9.168	4.107		- 0
			8.267	969.60	• ;	. 9	20.00		
	•		0.199	949.6	200	200	200		•
	23,		.6.63	8.868	12	1000	933		80
	8		6.873	20.00	**		00.00		100
			100.00		17	6.211	0.176		•
	82.		200	900	25	6.456	8.422		•
	***		8.923	6.000	4.0	9.99	610.0		
			0,042	.0.608	54	6.977	000		
			8.332	210.0	13	8,003	013.4		::
			-0.016	0.621	31	9.500			-
		8.88	8.463	198.401	=		200.0		
	*2	110.0	8.027	-8.816	63	6.673	404		
	::		8.176	9,935	69	0.123	001.		
			9.2.6	-0.933	29	430.6	804.0		
	• • • • • • • • • • • • • • • • • • • •		87.0	-3.921	96	0.190	4.607		1
			000	6	33	8.105	8 8 9 5		
	142.	5			56		-9.867		0.
	63.	8.00	108.8	670.0	, •	010	8.023		20
	527.	957.6	8.425	200			-8.007		21
	64.	0.077	8.836	179.9			200		22
	75.	0.0.0	0.823	10.01	20				23
	30.	8.878	9.666	8.894	2 :	0.00	400		2
	29.	9.224	8.179	446.6	0	8.638			
	227	9.100	9.186	-0.919	000	9.917	214.4		
			6.849	-8.918	56	8.978	000.8		2 6
			9.188	0.010	•	0.003	80.0	S	2
		000	25.0.0	8,957	-	8.634	160.0	200	,
			864.6	6.6	•	9.934	8. P.33	200.6	
	•		0.07	400.00	97	0.00	8.049		5
	• • • • • • • • • • • • • • • • • • • •	200	9 987	623.6-	13	8.853	8.865	86.	2
				500	62	0.652	4.054	86.8	2
	•			606	7	6.065	8.867	8.81	ń
	::		9.929	-8.883	38	0.826	9.050	88.8	,
	•		150.0	-8.826	99	6.649	20.00	•	,
		6	819.61	3.926	5	876.6	100.00	200	5
		6	0.057	569.60	70	0.867	174.0		3
	•		776	6.4.6	3.0	200.0	6.007	•	2
		100	6		43	6.237	6.242	•	2
			000		36	9.811	0.016	•	=
		21.0		500	61	B. 834	P. 642		•
		0.0			37	6.171	8.179		•
	.00	100.00			25	400.0	P.105	•	•
		165.0	2000		25	9.00	0,823	.8.	•
	. 20	205	0000		7	8.884	9.915		•
			27.0		61	0.011	6.82)		•
					67	8.174	6.196	8-	•
31	. 2	100	410		30	0.062	6.079		•
	13.		10.0		;	8.912	9.02		•
	. 21	100.00	200		58	8.831	0.04		•
	15.	200			28	9.168	9.18		•
	:			6 6	0.7	6.015	6.036	6 -0.021	80
		200				8.849	6.87		•
					•	8.995	6.05		60
	•				21	6.885	P. 031	1 -0.026	•
			. 6		30	800.0	8.63		n
			6	196.6	86	8.8.8	8.87		0
	•			66.6	1.0	0.176	B.28		•
	50.	200			•	9.826	8.86		•
	27.	5				6.931	9.067		•
		200		200.00		6.831	6.100	•	•
		50	9 (200	**	0.437	8.987	•	•
			•	7.6					•
		47176	9.150	0.01					۰

LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

rioni 8.83	5
TIONI	031111
OLM MODEL RUN I DEP VAR II VI RESIDUAL ROOT MEAN SOURRE OF FITTED EQUATION: 0.83	
SEP VAR 11 Y1	
GEN MODEL RUN I DEP VAR II YI	1

SEO	-		, .	•	•	•	•	•	•	9	=	12	13	=			-	: :				: :	22	2	77	52	56	27	88	50	80	31	32	33	70	23	36	37	99	90	4	7	42	43	;	•	9	17	=	9	20	5	25	23	3	93	90	37	80	30	9	6	8	
0884	0		21	47	25	5	•	=	•	23	36	32	43	25	a				;	3	; -	• •	3		35	9.	8	42	50	4.8	4	•	65	25	63	•	13	26	~	7	*	•	6	30	9	35	99	33	28	•	57	6	67	10	89	17	20	27	58	5	•	0	4	80
			٠.		_			•	_	_	~	~	~	2	2					, ,			,					•	•		•			90			^									•	•	•	•	•	_	_	_	_	_				•	•	•	_	•	•
FITTED						6		6	5		6.	6.0	6	6	6	6	6	6	6	6	6	•		•		8			6	0.0	6	8	6	•			6	8	9.0	6	0.0	6	6		0.0	6	6	8.					-		:								2.	
AFSIDUALS	8.91		 	00.0	8.00	60.00	.0.	9.85	10.01	10.0		8.82	8.99	9.65	6.0	00.0	. 6						99.0			10.9	10.6	8.82	9.82	2	9.61	8.98	68.80		8.63	6.63	6.89	46.0	8.93		8.62	66.8	0.01	9.05	8.93	6.63	10.0	80.6	9.00	8.64	8.63	8.62	6.83	9.12		70.0	5.0	8.68	0.07	6.63	0.05	6.83	1.0	- 12
961																																																																
1350	9.2			20.00		96.39	10.0	32.62	67	40.6	38.17	74.74	78.37	30.27	31.47	24.23	14.53	00	67.24						32.10	1.3	184.75	95.53	107.01	91.77	88.48	7.51	55.94	47.28	14.77	9.23	25.20	149.71	7.93	42,38	38.02	3.09	63	112.08	9	23.98	21.29	16.86	6.11				0		13.63			100001	118.17	61.93	123.98	10.03	2.7.	20.00
DEL RESIDUALS	8.93					8.6				•	50.0	9.82	8.92		8.89	6	50.00	6					2.5		6.60	8.83	8.92	8.92	8.91	96.8	8.92	9.92	16.0	10.0	0.04	00.0	6.67	6.03	68.8	6.83	0.05	6.94				9.91	80.6	0.05	6	60.63	20.4				10.0	24.4			200	200			2:	20.00
0937	27			10	0	0 0	35	8	6	38	20	25	63	•	3.0	•		37		19	2		1		3	35	12	-	12	•	3	30	52	27	37	13	69	5	62	=	-	~	;	2	8	•	67	6	2	=	n	•	5 0		6 :	:			93	:	-:		5;	
0834	37		:	2	-	5	:	0	,	•	80	28	33	52	•	•	99	99		-	ē						^	:	94	30	67	99	36	11	11	25	41	67	•	~	2	2	13	•	96	~	27	90	•	• ;		2 .	6	6 6	10	9:	::		`;				•	
4880	0.0	6	 	0.0	0.20	8	1.10	1.04	6.	2.17	5.19	5.19	2.98	3.51	3.60	3.78	6.4	. 36	79.7	56.	6					9.0	5.95	9.00	6.02	6.11	6.18	6.83	7.29	7.45	7.45	7.49	7.51	7.51	7.51	2.03	1.96	8.29	9.51	6.59	8.73		0.10	0.0		4.			0.00						10.01	20.01	200		::	
370 064	9.96	6	 		69.69		60.0	6.63	6.63	60.00	60.0	9.64	9.94	0.64	8.94	10.0	9.00	6.93		9.94	70.6					9.6		9.9	9.04	8.64	10.6	9.94	9.84	9.94		***	9.94	74.0	9.04	10.0	9.84	9.64	9.34	9.94	9.0	9.94																		
.0												12	2	:	5	91	11	1.0		50	21		,,			52	50	27	50	50	86	31	32	33	*	35	36	37	38	30	04	=	45	2	:	•	9			2:	90		70	20										

								THE LT	CONTROL OF	FAFE					
	1006		DEP VAR	18 41											
COMPONENT	A DE TO SE	EFFECT OF E	ACH VAR	BY THEIR RELATIVE		EACH OBSERVATION (IN UNITS	SERVATION	ITS OF Y)	RED BY IN	BY INFLUENCE	OF HOST	INFLUEN	INFLUENTIAL VARIABLES	A81.E3	
	0834.	*	VARIABLES												
			0 (V	9 5	2 -	a n	••	•	13	1.0	a	9	1.4	::	91
		X10H	1 6 X	O K X	OSOSX	L	X14H	X13	X18DSG	Х16н	X13H	X15050	X140SQ	X21H	X16050
	47	9.0	82.0	4 6			-6.12	.8.81	-8.13	19.6-	-8.02	8.91	-8.83	20.07	6.
~	25	600		2.1		88	6.92	.0.01	.0.01	.6.81	. 8. 92	.0.	6		
•	69	9.79	5 -	6			69.12	6	6						44.4
•		00.00		. 90						10.01		2.0	20.	.6.	
	0	8.63	200		000		8.95	-8.81	P. B2	-8.81	-0.05	9.01	8.63	10.01	80.0-
'n	58	9.54	. 8.33	9.16	0.07		8.82	10.00	8.82	. 8.	.8.92	9.9	6		60 64
•	45	9.50		. e.	66.61	000	6	6							20.0
	;	00.	8	8	0.01		*	10.0	90.0		20.0		. 8. 83		.8.
	;	-8.81	2 0	8 6	0 6	6.6	10.0	8.84	8.92	-8.81	-8.85	8.88	8.84	-0.98	88.84
•	50	6.13	-0.19		50.6-		9.87	.0.01	9.82	.0.61	-8.82	6	60		
0.	17	8 -			88.89										
					5 6		8.87		9.85	-0.01	. 9. 03	0.61	-0.94	-6.81	-8.88
9	27	9.12			-0.05	.0.03	8.82	.8.81	6.02	.0.01	-8.82	6.9	8.03		66
:	37	9 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6	8 8	0 6	6 6	20.0	6							•	
:	;	00.0			8.8		30.0	10.0	90.0	10.	-8.85	9.01	9.63	-8.01	68.69
7					6.6		-8.84	-9.91	0.85	18.6-	8.93	18.81	8.83	-8.91	. 68
2	69	9.90				.9.64	9.92	8.81	8.83		. 82	6	6		
:	61	2 2	9 9		8.0	9.91									
	;	9.89	40.0		0		70.0	8.0	8.85	.0.01	-0.05	e.	6.63	-9.01	-0.00
2	63	6 6		0.0	78.0	40.0	0.07	.8.81	0.02		-0.82	0.01	40.0-	-8.01	80.0
•	67	8.35			96.6		9.32	6	0.0		6				
11	29	000	0.00		000	6.6			•	•	•		6.63	10.61	80.
		9.00			. 8. 8.				8.83	18.81	10.0-	-0.01	0.01	-8.91	96.9-
	=	4 6	10.0		90.00	-8.93	.8.83	10.0-	9.93	-8.81	9.93	-0.97	9.93	-8.01	.0.00
10	40	8.93	6			50.0	-0.12	9.13	0.82	-0.01	8.82	6			
50	99	000			9.9	10.01								2	20.01
;		80.6	6.05	.0.0	3.00	. 6. 6.	3000	10.0	9.0		-8.85	8.81	8.83	10.01	-6.88
:	,				50.00	.0.03	0.95	.0.91	9.05	-0.01	-0.05	8.91	0.03	.0.01	-6.00
55	*	0.32	00.0	8.83			8.82		68.8	6	00				
53	34	8 8	6 6 6	00.00	000	10.01		. ;					2.0		
		-8.81	80.80	16.0	. 83			61.5	24.0		-0.05		0.00	00.0	99.0-
:	~			0.00	10.0	. 62	0.05	10.0-	0.01	-0.01	10.0	.8.06	8.03	.0.00	88.8-
23	•	9.91			.0.03	-0.65	9.84		6			. 0			
56	98	8 8	6.6	60.0	00.0	80.9	. :						10.0		88.8
	:	80.6	6				20.0			-8.01	-0.05		80.03	-0.91	84.6-
,	26		0 0	6.6	40.0	20.00	6.82		. 8.	18.61	-0,52	8.91	8.93	.0.01	.8.
28	•	00.0		98.	-0.07		.0.12	6	6				:		•
		9.95	6.		. 9	.0.0			•		70.0			9.07	00.0-

30

88.6		-0.80	-0.30	94.8-	98.4	8.98	94.9-		8. 8	98.8-	-8.86	80.0-	98.60	86.4	.0.0	-0.09	-8.00	64.6-	60.9-	-6.33	-6.06			88.6-	80.6	86.9-	-8.88	. 8.88	86.81	.8.88	88.85	8.81	.0.00	88.60	.0.0	
6.		-8.81	8.85	10.0-	00.0	8.83	-9.01		.0.01	-0.01	8.80	8.81	10.01	-8.81	9.02	8.82	.8.61	. 9.	10.01	8.98	10.0-	19.01	-6.91	.8.81	-9.01	-8.61	-6.08	.0.00	9.91	. 69	.8.81	16.91	18.81	19.6-		
48.0	-0.93	6.93	8.63	8.83	8.83	8.83	40.0			9.91	18.8-	. 8.	-8.93	6.03	8.93	8.84	.8.83	. 8.83	48.61	. 8. 83	.4.03	40.6	9.95	.8.83	-0.83	. 8.	. 9. 91	-8.84	48.61	-8.84	-8.8-	18.81	18.81	-0.03	.0.0	
. 9.	6	10.0	10.0	9.01	8.8	8	6		.0	.8.03	8.91		10.0	10.0	80.6	. 8.	. 8	8.83	9.91	8.83	89.63	-8.83	14.9	8.83	8.93	.8.83	10.8-	9.91	10.0	16.9		0.01			9.91	
. 9. 92	-8.85	8.97	.9.62	.0.02	.6.92	.8.82	6		-8.85		-6.92	.9.01	-3.83	-8.82	-9.92	94.6	-0.02	9.10	-8.82	9.10	6.19	9.96	-8.02	9.18	9.19	. 8.	19.81	-8.85	-8.83	-8.82	-8.82	-8.92	-8.85	-8.85	-8.05	
.0.01	.0.61	.8.81	. 6	.8.61	9.01	6			-8.91	10.0-	-0.01	19.0-	9.15	10.0-	.0.01	.0.01	18.01		-0.91	10.0-		-9.81	10.9-	18.81	.0.01	-0.01	10.01	.0.01	-0.01	-8.91	-9.91	-9.91	-0.81	18.01	.0.01	
10.0	. 6	10.0	9.91	0.01	6.6	6			6.8	88.80	18.81	10.01	19.91	. 9.		. 8.	18.81	19.0-	18.81	9.61	6.61	-9.82	-8.82	-8.82	-0.02	-4.82	-9.82	-8.82	-8.92	- 8.82	-9.82	-8.82	.0.03	.0.03	-0.03	
8.	. 8.	10.0-	-8.81		6			19.61	8.15	.0.01	-8.81	.8.91	10.0-	.9.91	9.97	.8.81	.8.91	. 8. 6.	. 6.	-9.81	-0.61	.0.	9.00	-9.81	6.61	-8.91	-8.91	.0.	.8.91	. 6.	. 9.	. 8. 81	.0.01	.0.01	.8,91	
8.	18.87	-8.87	9.82	6.85	8			18.8	-P.12	9.84	18.8	8.95	-0.12	8.82	.0.63	90.6	6.87	-8.12	10.0	-9.12	-0.12	.8.	-9.83	-8.12	-0.12	8.97	98.0	8.87	9.87	4.87	9.87	9.87	8.87	90.6	9.87	
. 6.92	-0.85	20.62	66		6.6			6.6	6.6	9.01	9.62	9.0	9.9	9.61	9.9	9.92	6.0	96	88	6.6	9.95	9.62	9.9	18.8	9 6	9 9 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8			66	6.6	6.6	66	9.0	8.0. 8.1.	9.0	99.0
86	6.	99	66	000	86.6			20.03	600		8.8		e e e	6.6	6.6	88.8	8.8	66	66	20.0		6.6	8.8	9.00	88.	8.8	6.0	6.6	6.6	9.90	9.07	9.0	9.91	9.91	6.6	8.81
6.0			8 6	66	6.6		2 6	6 6		20.	88.8	9.99	88.	16.9	6.6	66	68	6.6	66	66.6	86	6.6	20	46.	6.6	10.01	6.6	6 6	66						-9.83	•
6.6	0.37		6.6			3 2	2 2 2	6 6	56	20.0	66	000	9.93	6.0	6 6	66	66.6	0.0	50	6.0	66	66	5.5	888	5 6										6.6	
			6.6	• •			8 6	6.6		8 8 8	80	13.9	20.00	8 6	6 6	6.6	50.00		0 6	6 6	8.5	6 6	9.99	9.99	6.0	60	8	18.81	16.81	8.61	6.0	8.00	8	9.00	9.99	-9.61
22	33		:	2 :	B (20		25	5	-	;	a	3.6	•	: :	: :	: 2	; ;	, ,	: :		: 9		16	6		13	98	15			52	61		20	
			; ;		2	*	25	90	37			. 6	-			, ,		, .													2		5	25	63	

equipment is given in the first column for identification purposes. The third column shows the WSS DISTANCE for each observation. Here those observations far from the centroid of all observations can be easily spotted (observation 47 and 22). Under the heading "ORDERED BY RESIDUALS," the residuals are listed in the order of the magnitude of the residuals. This gives an indication of which observations are fitted the best (or worst). As can be seen, observations 46 and 13 are fitted best and 61 is fitted worst.

STANDARD DEVIATION ESTIMATED FROM RESIDUALS OF NEIGHBORING OBSERVATIONS

The cumulative estimates, S_n , of the standard deviation are printed in the second column of Figure 3. The WSSD_{jj}, of the observations in columns 4 and 5 are printed in the third column. Also the observations are ordered by their increasing fitted y values. At the top of Figure 3 is the residual root mean square = .03. The cumulative standard deviation column indicates that the standard deviation estimated from near neighbors is approximately .03, hence there is no evidence of lack of fit.

"COMPONENT EFFECTS" TABLE

The component effect, C_{ij}, of each variable on each observation is printed in tabular form (Figure 4) where the variables are ordered by their decreasing relative influences in columns, and the observations are ordered by their decreasing effects on the most influential variables in rows. Here the analyst can see which particular observations are most influential in their effects on the fitted equation. In addition this table can be used to determine the importance of high correlation among the independent variables.

STATISTICAL PLOTS

As with any endeavor dealing with Deductive Reasoning, the conclusions are dependent on the validity of the assumptions. Thus the analyst must have some means of verifying the degree to which the assumptions are satisfied. In addition to the number of statistics and statistical tables, there are five types of computerized plots that can be used to determine how close the data and fitted equations satisfy the assumptions. These plots give the analyst much insight

into the fit that the statistics alone cannot. The plots are used to determine (1) whether the assumptions of the method of least-squares are "nearly" satisfied, (2) just how well (or bad) the equation fits the data, and (3) to obtain further insights into the distributional properties of the data and how these properties affect the fit.

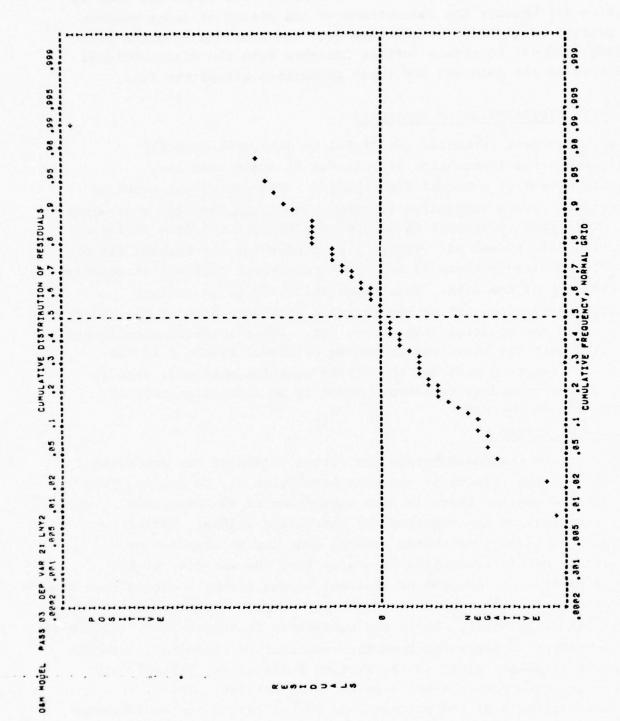
CUMULATIVE DISTRIBUTION OF RESIDUALS

When <code>k</code> independent variables are fitted to data with normally distributed error (Assumption A3), it can be shown that the residuals also have a Normal distribution. Therefore, the graph of the residual versus cumulative frequency should be "nearly" a straight line. This plot is helpful in determining whether the data satisfies Assumptions A1, A2 and A4. Figure 5 is a plot for the initial fit of ln(MTBF). Obviously there is an observation whose residual is separated from the rest of the data. This observation may be an outlier (violating Assumption A1) or may indicate that some form of curvature is needed in the equation (Assumption A2). After investigation it was determined that the point was indeed an outlier. Figure 6 is the cumulative frequency plot for the fitted equation obtained, with ln (MTBF) as the dependent variable. There is no indication here of deficiencies in the fit.

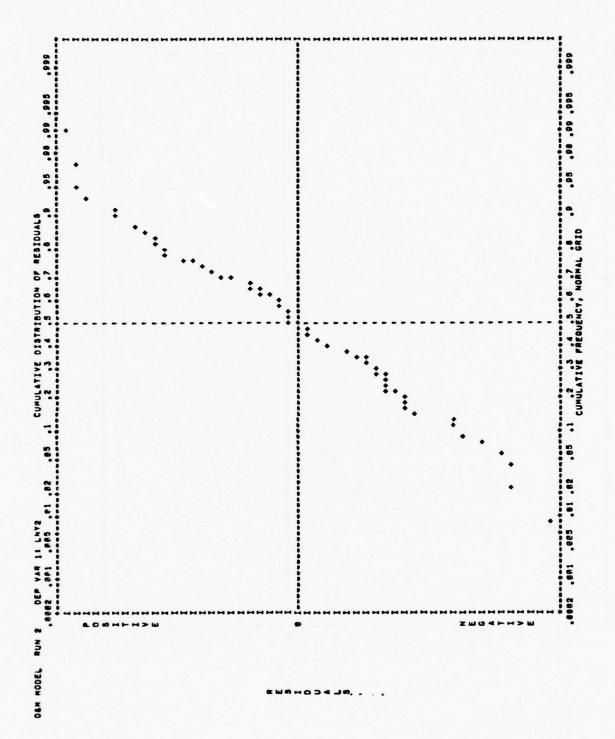
RESIDUALS VS FITTED Y

The plot of the residuals versus the fitted values of the dependent variable is also helpful in checking Assumption Al, A2 and A4. This plot may show whether there is some dependence of the magnitude of the residuals on the magnitude of the fitted values. Daniel and Wood [1], gives four common defects that may be revealed by such plots. Recall Assumption A3 states that the variance of the error is constant. The plot of residual versus fitted Y should then show an equal scatter about the 0-residual line. Figure 7 is a plot for the initial fit of ln(MTBF). As in the cumulative frequency plot, (Figure 5) one observation is separated from the remainder of the data. Both the cumulative frequency plot and the plot of residual vs. fitted y are necessary to determine whether a point is an outlier. Again, it this observation is at the extreme ends of the ranges of the dependent variable, curvature may be the solution. Figure 8 is a plot of









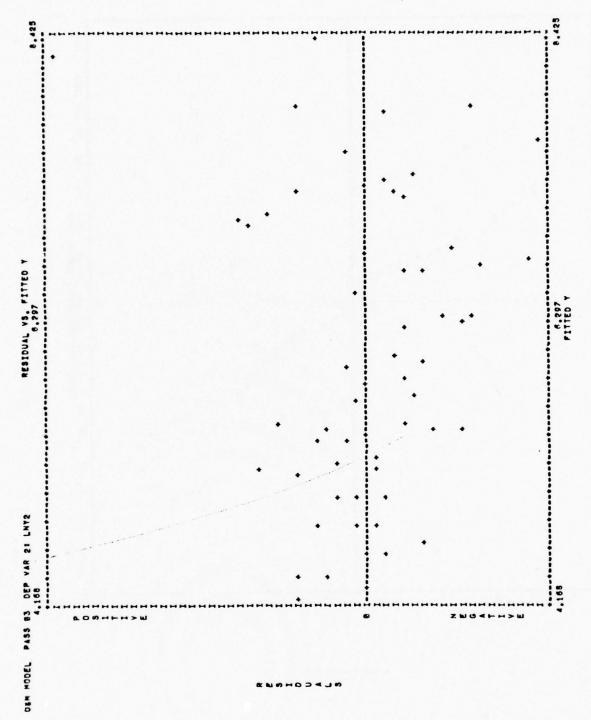
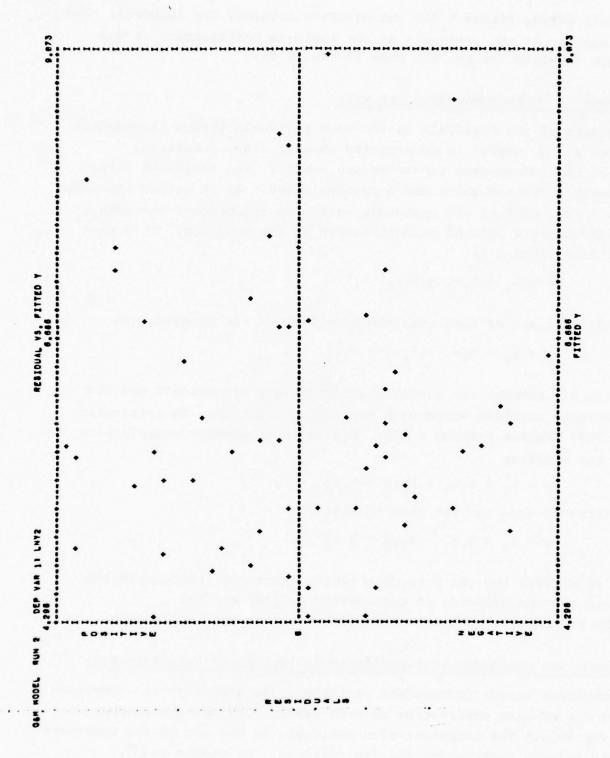


FIGURE 7



residuals versus fitted Y for the equation obtained for ln(MTBF). The equal scatter of the residuals do not indicate deficiencies in the equation obtained (as was the case in Figure 6).

RESIDUALS VS INDEPENDENT VARIABLE X(I)

The pattern of the residuals in the plot residuals versus independent variable \mathbf{x}_i is useful in determining whether other functional forms of the independent variables are needed. The residuals should be equally scattered about the 0-residual line. As an actual example, Figure 9 is a plot of the residuals versus an independent variable \mathbf{x}_1 where obviously a squared term is needed in the equation. This plot was obtained when a fit

$$y = b_0 + b_1 x_1 + b_2 x_2$$

was made to a set of data when the true form of the equation was

$$y = b_0 + b_1 x_1 + b_2 x_2 + b_1 x_1^2$$

For this fit however the 'global' statistics were significant and did not indicate anything wrong with the fitted equation. In particular $R_{\underline{y}}^2$ = .9047 and the F-VALUE = 228. Figure 10 is another example plot where the equation

$$y = b_0 + b_1 x_1 + b_2 x_2^2 + b_3 x_1^3$$

was fitted to data and the true form was

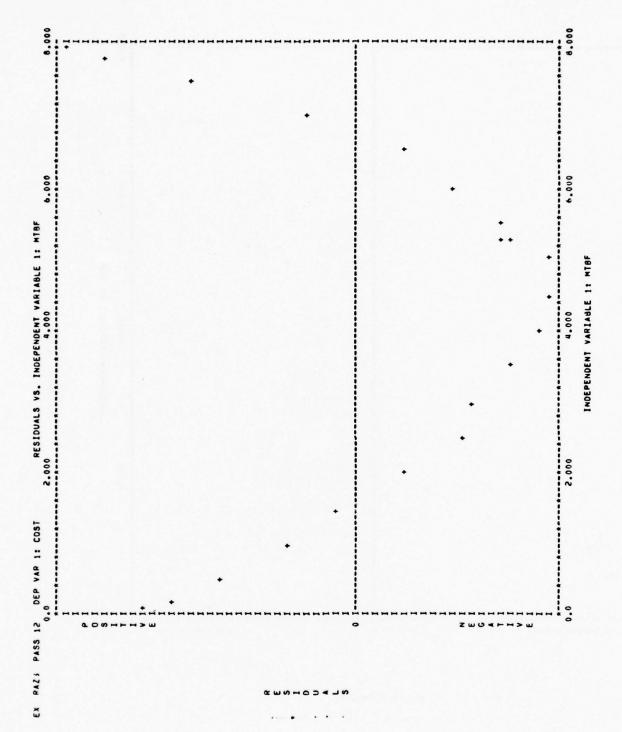
$$y = b_0 + b_1 x_1 + b_2 x_2^2 + b_3 x_1^4$$

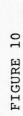
Here, R_y^2 = .9963 and the F-VALUE = 3267. These two simple examples indicate why the practice of considering only R_y^2 and the F-VALUE as measures of 'goodness of fit" is not statistically sound.

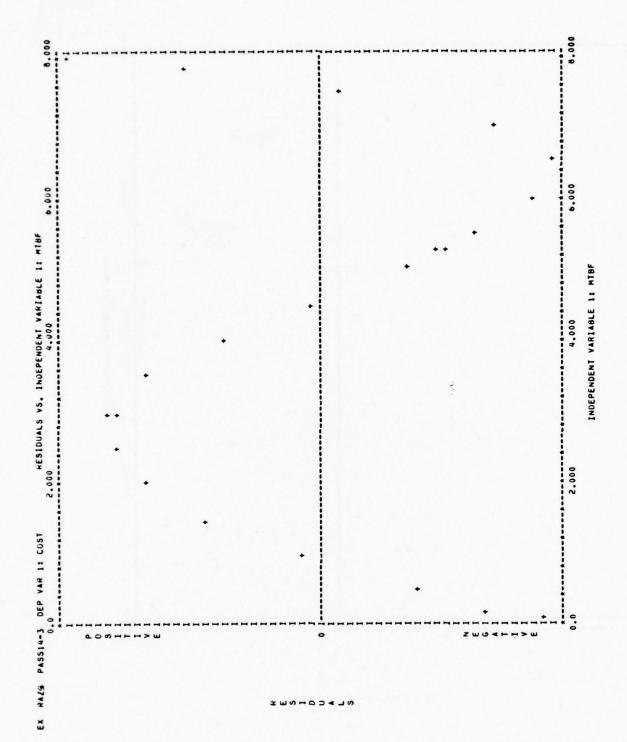
COMPONENT AND COMPONENT-PLUS-RESIDUALS VS INDEPENDENT VARIABLE X(I)

The component versus independent variable x_i is a plot of the component effect C_{ij} of each observation on each variable versus the independent variable. The component-plus-residuals is the sum of the component effects of each observation and its residual. As stated in [1],









"Component-plus residual plots are used as an aid

- (1) to choose the appropriate form of the equation,
- (2) to observe the distribution of the observations over the range of each independent variable and
- (3) to estimate the influence of each observation on each component of the equation."

Observations at the extreme ends of the ranges of the independent variables usually control the estimates of the statistics. The component-plus-residuals plots can be used (with indicator variables and the C_p -search technique) to determine if these extreme points are compatible with the remainder of the data. If it is determined that these extreme values are not compatible with the rest of the data, then either curvature should be introduced in that independent variable or other subjective information (introduced by indicator variables) about the points in question should be considered.

Figure 11 is a plot for the training cost equation obtained, where the independent variable is the % power supply. As can be seen from the graph, only one observation extends the range of % power supply by over 3,000%. It was later found out that introducing curvature in the % power supply had a significant impact on the fit. The residuals should be equally scattered about the component line.

CP VS P

The C_p -plot (developed by Mallows [3]), is a plot of the C_p -statistic for an equation versus P where P = κ + 1. For those equations with negligible bias, the C_p -statistic will fall near line C_p = p. Obviously the analyst would like to choose the equations with smallest total squared errors (C_p) and with the least amount of bias. Figure 12 is an example of a C_p -plot for the NRTS equation obtained where it can be seen that the C_p -statistic 1 is on the line C_p = p which indicates that there is no evidence of function bias. The C_p -statistic 2 is about the same as 1 but is above the line C_p = p and indicates that more is present in equation 2 than equation 1.

STATISTICAL, TECHNIQUES

There are two techniques utilized that are helpful in finding the subset collection of variables which best fits the data and in determining the stability of the equations obtained.

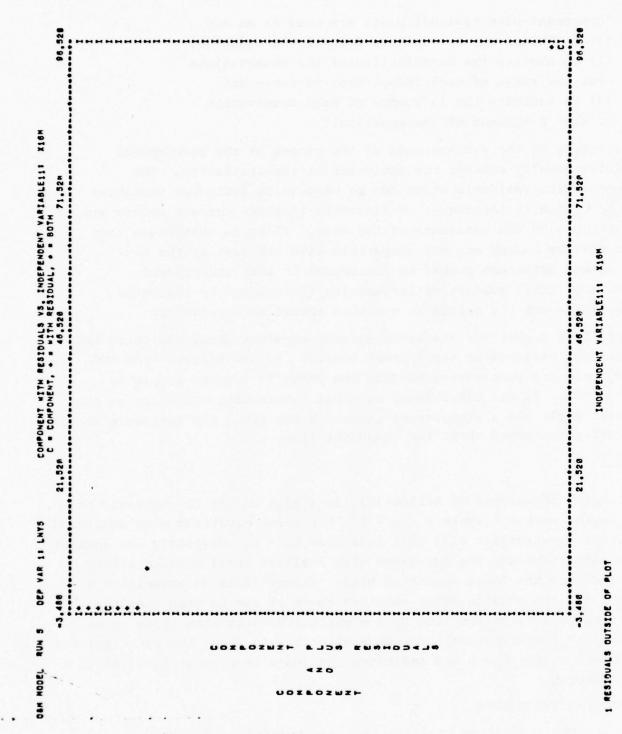
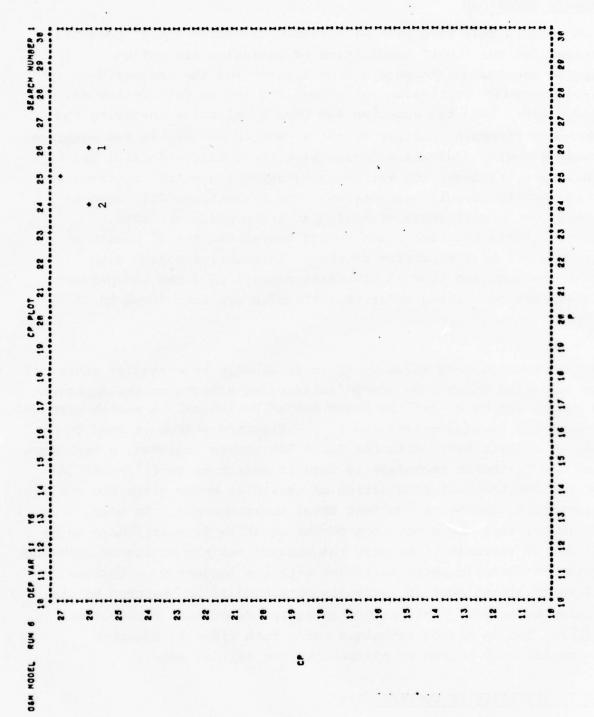


FIGURE 11





CP-SEARCH TECHNIQUE

Two approaches that have been widely used to search the 2^K possible equations for the "best" combination of variables are called "Stepwise Regression (forward and backward)" and the "F-test." Forward stepwise regression introduces the independent variables, one at a time, into the equation and uses a criterion involving the t_i values to determine whether or not x_i should be left in the equation. Backward stepwise regression begins with the complete initial set of variables and drops the variables by using a similar criterion as that of forward stepwise regression. The F-test is widely used to determine the significance of adding an independent variable. Obviously, these two techniques do not search all the 2^K possible equations, but only a portion of them. Moreover, a search with these techniques can lead to different results when the independent variables are correlated or if the variables are introduced in different orders.

With the equation form assumed, there is usually some smaller subset of these variables which have "very" influential effects on the dependent This subset can be called the Basic Set of Variables. A search proposed by Daniel and Wood [1] called the $t_{\rm K,i}$ - directed search is used to determine a Basic Set. With the Basic Set always included, a technique called the $\rm C_p$ -search technique is used to search up to 2^{18} = 262,144 equations for the best combination of variables which gives the smallest $\rm C_p$ -statistics, and hence smallest total squared error. In some cases (e.g., when the wrong form of the equation is used) there is no Basic Set of Variables, and here the analyst has the option of choosing a basic set (usually those variables with the largest t_i - values) until there are at most 18 variables remaining to be searched by the $\rm C_p$ -search technique. This method is called Fractional Replication (See [1]). The $\rm C_p$ -search technique using both types of searches have consistently helped to narrow down the initial set.

CROSS VERIFICATION OF COEFFICIENTS

Once a presumably final equation is obtained, the analyst must determine the stability of the obtained equations coefficients. There

may be a few observations in the data base (such as those with large WSS DISTANCE and large residuals) that are not compatible with the rest of the data and may be controlling the estimates of the fitted coefficients. A way to determine stability is to drop those observations, run another regression and determine the effects on the least-squares estimates of the coefficients. This technique is called cross verification of coefficients with a second sample of data and provides a rigorous test of the data, the model and the fitted coefficients.

Component-plus-residual plots of the second sample of data (where residuals are calculated using the initial coefficients) may point out those observations which may indicate that other forms of curvature are needed.

SECTION VI

AN EXAMPLE,

MEAN TIME BETWEEN MAINTENANCE ACTIONS (MTBMA)

The data (a year's worth of data) was chosen from existing data systems to determine the CER's and PER's used to estimate O&M cost and is shown in Appendix A. There are 64 pieces of avionics equipment (observations), called LRU's, on which the study is based. Each observation can be identified by its observation number and Work Unit Code (WUC). Also associated with each is a total of 27 variables, of which 21 are independent variables and 6 are dependent variables. The independent variables are of two types, quantitative and qualitative. The usual types of variables in a regression exercise are quantitative (i.e., variables that may take on values over a given range) such as weight or other physical characteristics of the equipment. Many times additional (qualitative) information is available, such as certain characteristics of the equipment or a certain class in which the equipment belongs, which should not be discarded, but should be introduced into the regression, since more information should lead to a better fit. "Indicator" variables (variables which take on the value of 0 or 1) are used to introduce qualitative information into the regressions. A "1" indicates that the observation is in a certain class and a "0" indicates that it is not.

The type of aircraft in which a piece of equipment is used and the equipment avionics areas are the two qualitative classes used in

this study. There are three types of aircraft: Fighters, Bombers and Cargo and three areas of avionics: Navigation, Sensory, and Communications. Table 1 shows the observation numbers of each piece of equipment and the class to which it belongs. For instance, observation 9 is a piece of navigations equipment that is used in a fighter. Since there is no sensory equipment used in cargo type aircraft, no observations are present there. The numbers in parenthesis indicates the quantity of observations in each category. Thus 18 LRUs are used in bombers and 16 LRUs are sensory type equipment. The numbers in the corners of the inner rectangles indicate the number of observations which fall in the respective interactive classes. There are 4 observations which are used in fighters in addition to being communications equipment. Table 2 lists the names of all the variables and their associated variable names used in the computer printouts of the regressions. Also, listed are the units in *which the variables are expressed. The quantitative independent variables and the dependent variables are defined in Volume I of this report. The indicator variables (qualitative independent variables) need some further clarification.

BOMBER and CARGO are indicator variables used to represent equipment that is used in bomber and cargo type aircrafts respectively. and COMM are indicator variables indicating that the avionics areas of the equipment are sensory and communications respectively. It is noted that there is no indicator variable for fighter aircraft or navigation equipment. Using indicator variables in this fashion, that is having a "baseline" of each class or category, can be very informa-Without loss of generality, fighters are chosen as the "baseline" for the aircraft types and navigation equipment are chosen as the "baseline" for the avionics area. We can then find those members of a certain class that are significantly different from the baseline. As an example, if the indicator variable BOMBER is significant enough to be in the final equations for MTBMA, this may be interpreted to mean that the MTBMA for equipment used in bombers is statistically significantly different from the MTBMA for equipment used in fighters (the baseline). Conversely if COMM does not remain, this indicates that the MTBMA for communications equipment behaves in a similar

Qualitative Categories

TABLE 1.

	I	FIGHTER (29)	RS		MBERS 18)		CARGO (17)	
NAVIGATIONS	1	6	11	15	21	27	33	38
(35)	2	7	12	17	22	28	34	39
	3	8	13	18	24	29	35	
	4	9		19	25	31	36	
	5	10		20	26	32	37	arts
			13		10			12
SENSORY	40	45	54	56				-
(16)	41	46	55	51				
	42	47		52				
	43	48		53				
	44	49				Dens.		
			12		4			0
COMMUNICATIONS	56			60		64		
(13)	57			61		65		
	58			62		66		
	59			63		67		
						68		
			4		4			5

TABLE 2.

Variables Used in the Regressions

Independent Variables

"Indicator"

 $X1M = (BOMBER - \overline{BOMBER})$

 $X2M = (CARGO - \overline{CARGO})$

 $X3M = (SENS - \overline{SENS})$

 $X4M = (COMM - \overline{COMM})$

 $X5 = X1 \times X3$

 $X6 = X1 \times X4$

 $X7 = X2 \times X4$

Quantitative

X8 = Unit price

X9 = Volume

X10 = Weight

X11 = Components Count

X12 = Components density

X13 = % Digital

X14 = % Analog

X15 = % Electro-mechanical

X16 = % Power supply

X17 = % Transmitter

X18 = % Solid state

X19 = Power Dissipation

X20 = Utilization factor

X21 = % BIT/FIT

Dependent Variables

Y1 = Maintenance Manhours per Operating Hour (MMH/OH)

Y2 = Mean Time Between Failure (MTBF)

Y3 = Mean Time Between Maintenance Actions (MTBMA)

Y4 = Logistics Support Cost per Operating Hour (LSC/OH)

Y5 = Training Cost per Operating Hour (TRAIN/OH)

Y6 = Not Repairable This Station (NRTS)

manner as that of navigation equipment (the baseline). There also exists the possibility that communications equipment used in cargo type aircraft might have a significantly different effect on the dependent variable from that of navigations equipment used in fighters.

Three indicator variables (which are products of the four initial indicator variables) that can be used to determine the effects of such interactions are: BOMBER x SENS, BOMBER x COMM and CARGO x COMM. In the regressions however, the 7 indicator variables used are the ones shown in Table 2, where the bar above the variables indicate the mean. This is done in order to reduce the sometimes high correlation between the indicator variables and their products. If any of the three interactive variables X5, X6 and X7 are significant, then the analysis of the MTBMA example (above paragraph) will have quite a different interpretation. If interactions prove to be significant (as was the case in all regressions equations obtained in this study), the interpretation of how the levels with which the two classes (aircraft type, avionics area) compare with their baseline should not be used, but the analyst should find a means of interpreting which specific interactions are significantly different from which others.

Returning to Appendix A, we see three lines of information associated with each observation. The first line lists the 7 indicator variables, the second the (quantitative) independent variables and the third line are the 6 dependent variables. Thus observation 2 is a piece of navigation equipment used in a fighter with weight = 36.60, % solid state = 73. and MTBMA = 274.

Initially 85 LRUs were considered for the study. Many of the observations were dropped from the analysis because of the lack of data or the difficulty in obtaining the necessary data. Other observations, such as equipment which had not been in the Air Force inventory long enough to experience "good" data, were discarded so as not to introduce bias in the results. In addition a couple of the observations had missing dependent variable data and were omitted for that particular fit. Once all the data has been collected, there should be a panel of qualified experts on the studied equipment to determine the validity

of each data element in the data base. This was done to a limited extent in this study (because of time constraints).

Many times the statistics, plots, tables and techniques indicate that some observations do not behave like the remainder of the data. Besides other possible subjective variables, curvature may be causing this instability. In addition to the variables shown in Table 3, two transformations of the independent variable (the square and natural logarithm) and a transformation of the dependent variable (natural logarithm) are introduced into the regressions when curvature is indicated. The natural logarithm transformation is considered for those variables whose range is contained in the positive real numbers.

Since the LLSCFP allows six card columns to identify the names of the variables, alphanumeric representations consisting of six letters or less are used in computer printouts. Using X8 as an example, the transformed independent variables are of the form

$$X8M = (X8 - \overline{X8})$$

$$X8DSQ = (X8 - d8)^2$$
,

and LNX8 = lnX8

where the bar indicates the mean, d8 is the d-statistic of variable X8 and ln is the natural logarithm. Table 3 shows the d-statistic and mean for each of the variables used in the regressions.

Before beginning any regressions, the data must be critically analyzed for outliers (impossible values) and for what is known as "Nested Data." The data is said to be "Nested" if some of the observations have all or nearly all the same or approximate x_i -values. Obviously outliers would have a significant impact on the fitted coefficients, thereby yielding the incorrect relationships. If the equations are fitted without checking to determine whether or not the data are nested, the wrong factors may be significant. The analysis of "Nested Data" was first introduced into the statistical literature by Daniel and Wood ([1], Chapter 8).

Since there is human intervention in the development, collection and analysis of data, outliers might not immediately become apparent.

TABLE 3

D-Statistic and Mean

24	1				
	мін/он			MTBF	
I	MEAN(I)	DSTAT(I)	I	MEAN(I)	OSTAT(I)
8	27173.650	133726.400	8	27536.109	133828.600
9	1438.048	3307.065	9	1483,725	3337.016
10	34.981	64.404	10	36,221	65.088
11	909,889	2948.044	11	940.016	2965.510
12	0.924	2.489	12	0.924	2.489
13	7,556	43.028	13	7.677	43.038
14	62,683	49.183	14	61,677	48,615
15	16.000	46.378	15	16,258	46.403
16	3,429	49.827	16	3.484	49.829
17	10.492	40.395	17	11,065	40.258
18	61.297	51.914	18	61.044	52.254
19	368,968	723.684	19	382.419	729.178
20	1.657	1.681	50	1,645	1.679
21	4,825	26.957	21	4.694	27.148
	MTBMA			LSC/OH	
1	MEAN(I)	DSTAT(I)	ı	MEAN(I)	DSTAT(I)
		00.21(1)			
			8	26943.220	133606.300
8	27536.100	133828.679	9	1457.984	3311.530
9	1483.725	3337.016	10	35,570	64.314
10	36.221	65.488	11	927.746	2956.742
11	940.016	2965.510	12	0.933	2.501
12	0.924	2.489	13	7.556	43.028
13	7.677	43.038	14	63.349	48.895
14	61.677	48.615	15	14,937	46,991
15	16.258	46.403	16	3.429	49.827
16	3.484	49.829	17	10.889	40.172
17	11.065	40.258	18	61.138	51.898
18	51.044	52.254	19	374.159	722.249
19	382.419	729.178	20	1.639	1.681
20	1.545	1.679	21	4,555	27.288
51	4,694	27.148			

TABLE 3

D-Statistic and Mean (Con't.)

	TRAIN/	ОН		NRTS	
I	MEAN(I)	DSTAT(I)	I	MEAN(I)	DSTAT(I)
8	26959.100	133922.500	8	27288,080	133763.300
9	1445.145	3305.984	9	1475.339	3321.294
10	34 955	54.418	10	36.005	64,594
11	889,468	2982.834	11	936,758	2961.209
12	0.904	2.522	12	0.938	2.505
13	7.677	43.038	13	7,677	43.038
14	62.516	49.242	14	63,468	49,468
15	15.823	46.514	15	14.468	45,651
16	3.484	49.828	16	3.484	49.828
17	10.661	40.474	17	11,065	40.258
18	61.108	51.989	18	62,108	52.151
19	372.097	723.719	19	378.529	723.518
20	1.646	1.680	20	1.645	1.679
21	4,839	26,952	21	4,694	27.149

However, the computerized plots of an initial regression may point them out. Figures 13-17 show summary computer printouts of an initial run (Pass 03) after fitting the equation

$$y = b_0 + \sum_{i=1}^{21} b_i x_i$$

where Y = ln MTBMA and x_i = Xi, i = 1, 2, ..., 21, are the 21 independent variables. The statistics (Figure 13) are not significant and the stars, *, indicate error is too large to be printed out in the space provided. The cumulative standard deviation table (Figure 15) shows that there is a serious lack of fit. The cumulative distribution plot (Figure 16) and the fitted value plot (Figure 17) indicate that there is a single oversized residual (possibly an outlier). Since there is also the possibility of curvature, a fit is made with Y = ln Y3 = ln MTBMA with printouts shown in Figure 18-22. The statistics using the natural logarithm are greatly improved but Figures 21 and 22 also (although not as profound as before) indicate that an oversized residual is present.

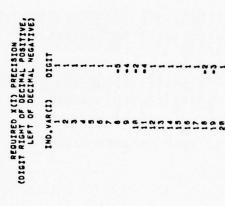
In Figures 14 and 19, under the heading "ORDERED BY RESIDUALS", we immediately see that observation 36 with WUC 7171A is the culprit. In addition Figure 13 shows that the maximum value of the dependent variable, MTBMA is 31,490 hours, and hence there is an observation (piece of equipment) in data base with such a large MTBMA. On investigating the data collection system (AFM 66-1 "6-LOG") from which the MTBMAs were extracted, it was found that observation 36 with MTBMA = 31,490 was a subassembly of an LRU (called an SRU) and not a LRU. Since the study is based on LRUs, leaving this particular observation in would bias the results, and therefore another observation with WUC 71716 was used in its place.

In addition to impossible values and discrepancies in the data systems, outliers may be caused by simple keypunch errors. If those errors are profound, the computerized plots should reflect the discrepancies. In many of the regressions considered throughout this report, the plots were helpful in detecting those observations which

FIGURE 13

LINEAR LEAST-SOULARES CURVE PITTING PROGRAM

2.750 63 1.970 93 1.970 93 1.970 93 1.970 93 1.970 93 1.970 93 1.970 93 1.970 93 1.270		311	COEF.9(T)	3.E. COEF.	T-VALUE	R(1) 50RD	MIN X(1)	MAX X(1)	RANGE X(1)	REL, INF. X (1)
X		× 1 ×	8.324450 B			2480				
X X X X X X X X X X X X X X X X X X X	~	121	3.818310				2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	18-000-1	1.0000 90	
X	•	H CH	1.773290 8					14-0202	88 0844.	6.12
X	•	***	A 275.070 a				14-0000	1.0000	1.6700 00	8
X			0.01.01.01		2.	0.75	-2.14BD-P1	12-0266.	1.6880 86	0.05
X		,	1.000000		•	8.34E4	-2.9330-91	5.3870-01	7.4280-01	6.6
X			2.977.050 6			9.3871	-2.1650-01	5.7350-91	7.9000-01	9.91
X			4.356000 0		1.2	0.4712	-2.1650-01	5.7350-01	7.5000-01	-
	•	E ×	8.719880-P	_	8.8	9.7129	1.5590 92	3.2400 05	3.2380 05	
	6		7.517030-0	_	0.0	8.8543	3.2200 91	2000	A 1740 94	
	-		. 1.4820 P		1.3	8.0108	00 0000	1 7370 83		
X X X X X X X X X X	:		-8.937540-P			9.649	6 0000	2000	7 6200 91	
	12		-3.616930 B		4.	8.5792	4.9230-93	6 6740 99		
	2		-2.297210 A		4.0	8.00A2	6			
	-		-2.174640 A		4.0	1000				200
	13		-2.178420 B		7.6	000	. 6	000		
18 X17 -2, 2074AC 22 5,660 P2 8.4 8.9975 8.4 1.0000 P2 1	9		-2.58P950 A			0000	. 6			
19 X19 9.757120 PM 2.290 R1 6.4 6.539 PM 1.1 8.5659 PM 1.1			-2 987860 P			000				>
	1.8		9. 850120 0		. 6			200000		
18 X21 9.46210 82 2.430 83 8.3 8.20 7 1.1310 83 1.030 83	61		2.243640 9		-			200000		8.83
	20		6 306220 a				0000	100000		8.12
A CASETALTION S						1490.0	2000000	Se CHAC.	2. 2890 000	46.8
OF ORSERVATIONS FIND, VARIBLES UNAL DEGREES OF FREEDOM UE ROOT MEAN SOUARE UNAL ROOM SOUARE WAS S		134	9.402110 8		:	9.5878	0.0	6.1880 81	6,1000 61	9.18
IND. VARIABLES UNAL DEGREES OF FREEDOM UNEL POOT HEAN SQUARE UNAL MEN OF MEN SOURRE	J. OF DASER	VATIONS		62						
:	1. OF IND.	VARIABLE	5	2.2						
:	SIDUAL DEG	REES OF	FREEDOM	45						
:	VALUE									
:	SIDUAL POD	S NEAN	GUARE	4305 10319194						
	SIOUAL MEA	SUUARE								
	SIDUAL SUR	OF SOUA	23	***********						



VAR 11

PASS 03 DEP

THE CHARLES AND THE CONTROLES OF A DECIDENT AND A STATE OF A STA $\begin{array}{c} V_{1} = V_{1} + V_{2} + V_{3} + V_{4} +$ NCE SY

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LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

STANDARD DEVIATION ESTIMATED FROM RESIDUALS OF NEIGHBORING OBSERVATIONS (OBSERVATIONS 1 TO 4 APART IN PITTED Y ORDER). RESIDUAL ROOT HEAN SQUARE OF FITTED EQUATION: 4395,19 DEM MODEL PASS 83 DEP VAR 11 YS

64.0	0.0			-	000	DEL RESIDU	50		0884	SEO
382.27	0.	37	9 6	4.14	200	727.89	-2850.38		•	-
1.97	0.0	91	0		44.23	19.54	-2081.99		45	~
6.03	60.03	37	27		36.78	20.10	-2042.34		97	•
2.87	8.89	5.5	67			2000	1861-		~	•
4.47	40.0	20	52		4.33	517 41	00.00		75	n
500	62.6	a 0	^	1265.54	42.63				2:	0
	00.0	35	36		38.59	227.33	-1118 32		0 0	
	8.12	22	21		3.23	428.31	-1964 72			0 0
		c	25		23.11	148.39	-1000			
	200	80	60		22.21	169.12	-832.41		6	9 .
			28		9.0	16.99	-733.80			::
		0	•		0.23	19.15	-733 AG			
200	1.5	0	28		9.17	68.70				2 :
	1.5	,	22		6.0	41.49	.624		?:	::
		9	28		2.98	35.32	. K2 & 2 %			2
		69	26		18.84	83.68	2405 73			
	0.19	61	37		16.18	340.08	2000		200	-
	6.1.8	63	27		13.82	171 08				•
.14	0.22	25	4		00	20.00			22	2
.85	0.23	:	•			10.751	-185.95		54	8
.47	8.23	69	•			02.041	-104.55		7	2
.35	0.25	52	13			00.00	9.71		•	55
.73	6.30	20				131.93	193.78		•	23
89.	6.39	;	28		00.00	384.49	119.77		38	54
. 95	60.0	6	25			88.92	274.86		31	25
. 83	8.47					1 99.18	455.98		-	58
. 0.5	9.62	-				538.98	728.59		6	27
.62	9.74	. 0			100	739.58	838,68		47	28
.42	6.63	•			3.43	785.13	904.78		24	50
.23	1.13	2			21.0	150.06	989.33		58	30
16.	1.21	5				221.43	983.85		00	31
**	1.26				500	166.35	1924.79		99	35
.95	1.49	22				858.92	1940.24		50	33
.75	1.40					155.27	1711.36		28	3.4
34						316,11	1834,69		;	
			3 6			1626.18	1833,33		•	
7.5	-					1811.19	1893.98		21	33
6.8	1.41	67	9			9.74	2037.18		26	38
69	1.44	58				39,33	2438.23		37	30
41.28	1.45	64	35			25.55	2042.76		27	9
35	1.64	29	62				2178.92		•	7
.58	1.69	30	99				2192.08		43	4
•	1.78	•	80				06.0022		**	5
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24	18.0	5	5				2575.01		200	,
	• • • • •	- 1	20				2752.98			
	27.7	0.	;				2947.99		2 6	
26	25.50	0 4	5 .				3229.51			
3.			5 6				3275.98			
90	7.5		ç:				3334.96		. 4	
20	200		2 5				3454.18	• .	34	3
31	2.62						3565,67		33	
1535,80	2.65			00.000	49.33		3635.91		52	20
25	2.65	000					3968,98		93	24
8.8	2.72	99					4667,61	-	30	38
28	2.85	8					4872.39		53	30
28	2.98	47	63				4802.39		51	69
									2000	
-	2.98	*	-				20100		32	-

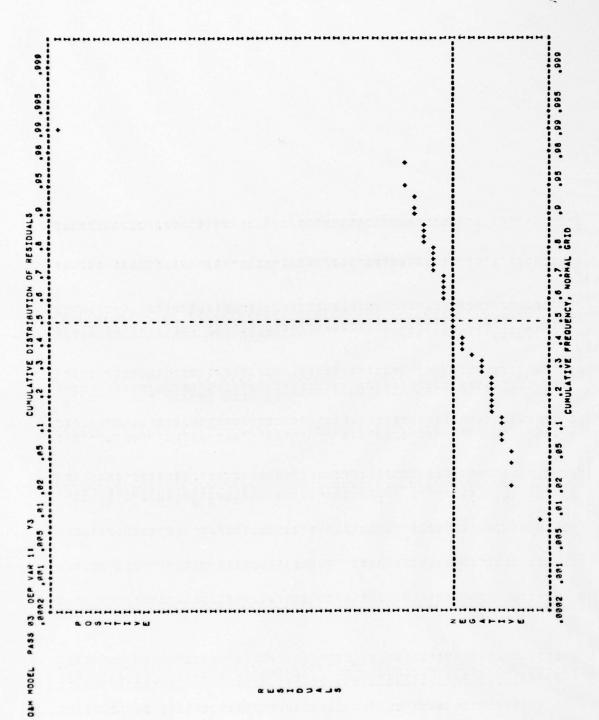
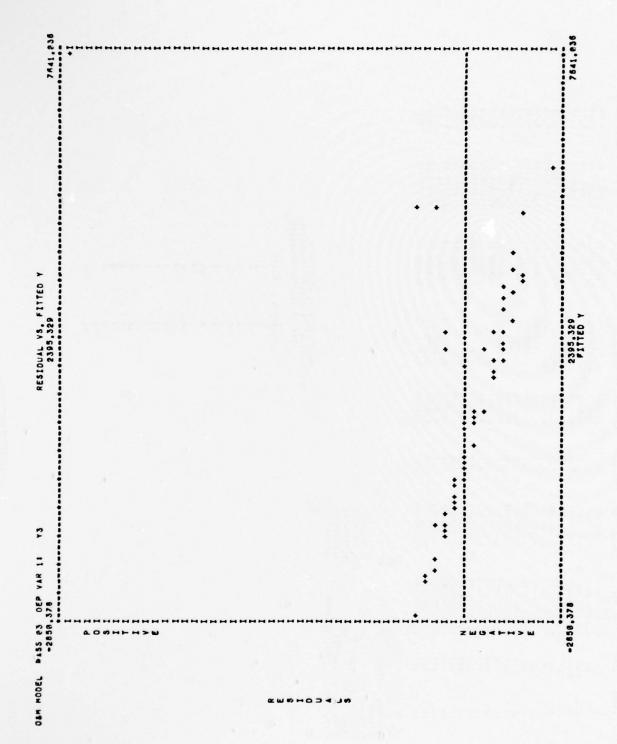


FIGURE 16





PROGRAM
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IND.VAR(I)	MAN	COEF,8(1)	S.E. COFF.	T-VALUE	R(I)SORD	HIN X(I)	MAX X(I)	RANGE X(I)	REL, INF. x(1)
6		7.793620 48			-				
-	×1×	5.858370-01	5.540-01	-:-	8.7850	-2.7400-01	7.2600-01	1. 9980 48	
. ~	XOX	8,833530-21	6.580-01	1.3	8.4338	-2.7400-01	7.2590-91	1.6960 98	4
	X3H	7.255430-01	4.230-01	1.7	9.5800	-2.58v0-01	7.4200-01	1.0000 00	9.11
•	1 4 X	-7.656A90-21	3.890-01	2.8	A. 4288	-2.1000-01	7.9000-01	1.2000 00	0.12
	S X	6.779220-01	7.730-01	0.0	8.3464	-2.0330-01	5.3870-01	7.4230-91	6.6
	×	1.743210-31	8.200-01	2.0	R. 3871	-2.1650-01	5.7350-01	7.9920-91	9.92
. ~	× 7	-7.315130-01	7. 690-71	0.	8.4012	-2.1650-01	5,7350-21	7.9000-01	6
. •	×	2.858460-05	3.670-08	80	9.7129	1.5800 02	3,2400 05	3.2380 85	9.14
0	, o	2.903130-04	1.930-04	1.5	P. A 5 4 9	3. DAND BI	8.2000 03	8.1720 93	9.36
-	× 1.8	-4.144290-92	1.330-02	3.1	00:0:0	1.2000 00	1.7370 P2	1.7250 02	64.1
:=	x 1.1	-2.385990-94	1.540-94	5.1	9.6499	9. 0880 98	7.6380 83	7.6290 03	9.28
	x 1.2	7.253570-02	1.750-71	4.	P. 5792	4.9230-03	6.5740 23	6.6490 08	78.0
: :	x 1.3	-1.820370-02	1.250-01	9.1	A. 0982	8.8	1.0000 02		8.28
:	¥1.4	-0.377110-N3	1.210-01	1.0	0.9993	8.6	1.0000 02	1.0000 92	41.6
	x 1.5	-A. 480350-94	1.220-41	6	1656.0	0.0	1.0000 92	1.0000 62	
91	x 16	-2.073120-02	1.220-01	8.8	9 66 6	6.6	1.9990 92		0.32
	x17	5.427920-04	1.210-01	0.	8.9975	6.6	1.9880 92		6.6
	x 1.8	2.775790-03	6.910-93	6	6.659.9	6.0			40.0
	61x	-6.610770-04	4.400-04	5.1	6.5659	7.880 98	1.6480 83		P. 17
60	x2x	-1.941200-01	5.210-01	0.5	9.8297	3.0000-01			58.8
51	x 2.1	2,934040-02	1.480-02	2.8	9.5878	6.	6.1000 01	6.1890 91	9.27
O. OF DASE	RVATION	•	62						
NO. OF IND. VARIABLES	VARIAB	LES	21						
RESIDUAL DEGREES OF FREEDOM	GREES O	F FREEDOM							
F-VALUE			6.9						
RESIDUAL ROOT MEAN SQUARE	TOT MEAN	SOUARE	94294949						
RESIDUAL MEAN SQUARE	AUGS VA	R	A. A8915375						
RESIDUAL SUM OF SQUARES	JH OF 50	UARES	35,56514993						
OTAL SUM C	SE SOUAR	50	109,92202018						

FIGURE 18

DEP VAR 21 LNY3

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LINEAR LEAST-SOUARES CURVE FITTING PRUGRAM

9.04	APART IN PITTED Y ORDER).
RESIDUAL ROOT HEAN SQUARE OF FITTED EQUATION: 8.94	STANDARD DEVIATION ESTIMATED FROM RESIDUALS OF NEIGHBORING ORSERVATIONS (OBSERVATIONS 1 TO 4 APART IN PITTED Y ORDER).
DEM MODEL PASS AS DEP VAR 21 LNYS	STANDARD DEVIATION ESTIMATED FROM RESIDUALS OF

SEO.	-	8	,	4	0	۰,	. 40	٥	1.0	= :	~ -	4	2	16	2:	0	50	21	25	2 4	52	28	27	e e	2 20	31	35	200	33	36	90	39	6 :		3	:		6 7	•	•	e :	200	33	2.0	32	90	6		63		20
0884	1.8	11	90	67		22		60	17	£ ;	6 6	. 4		28	9 6		28	•	":	22	~	38	.	* 6			= :		33	. 62	- 66	1.0	S, C			23		; «	•	28	52		. 4	32	90	25	20	9	32	8	10
FITTED Y	4.81	4.87	4.14	4.23	4.4	24.4	4.47	4.47	4.51	4.52	0.4	. 4	4.76	4.78		0 0	5.95	5.83	20.00	20.00	5.17	3.26	5.34	00.0	9 60	5.57	90.0	90.0	5.73	5.77	20.00	5.97	6.04	2.0	9.36	6.37	6.37			6.82					7.99	7.12	7.17	7.52	7.50	8.40	8.48
ESTOUALS	-:	-		•	4	1.32	. ~	0.42	6.53	9.34	54.6		4.4	80.6	0.71		1.32	1.11	9.47	2.3	68.6	08.0	98.6	74.9	4 4	P. 42	9.48	2.25	6	04.6	****	6.79	1.33			2.47	1.95	2.19		1.78	9.12	9.50	1.63	1.68	9.16			3.61	1.21	60.0	0.02
. 95		~			12.04	200	. 0	1.14	9.17	6.54	8 6	3.5	5.5		12.11	ė v	9.0		1.78	200	7.81	8.06		n i	2.87	9.73	12.39	2.77	3.58	1.70	2.37	2.98	1.84		2.11	1.72	1.00	1.10	1.24	9.00	1.42	7	1.79	8.83	1.72	2.67	2.47	2.73	9.76	6.6	7.92
S	6	8.42	~	٠.		6.50	-	^	1.07	S . N	2.0			86.0				9.31	6,1	200	:	~	1.39		4.6		10.0			9.19	9 60	2.63		5.6	. 0	~	•	96.		0.41	9.36	- 0	1.17	2.19	9.18	9.42	4.14	2.6.2	::	9.17	'n
088V.	51	38	69	27	67			52	9	9	n •	. 5	. ~	0	en c	~ ~	. K.	50	8	2 ~	56	7	6	25	2 5	53	-:		39	9 2	20	36	5	4 -	*	4	32	=:			23	- 5	4	•	17	-:		9 0	8	• ;	:
085v.	ç	57	61	37	9		. 40	20	53	63	2 6	4 4	. •	-	~ .	0 0		•	~ 0		•	•	22	c ·	3	?	6.0	• 5	32	8		30	5.	2 4	3	13	21	9 3	63	2	54	• =	. 6	;	9	6	C 0		3.0	=:	2
1880	8.8	6.6			0	20.0			00.0	9.30	4 6		3.23	9.22	0.24	2.0	4		0			17.6	8.74		200		0.79			0.0		66.6	6		26.1	1.84	1.04		88	1.38	1.98	200	1.1.	1.13	-:-	1.14			1.24	1.37	1.35
910 0EV	6.19	~	~	2	?	20.00	. 4			-	500						. 0	0.	0.	200		0.		•	2 6		0.	1.02	1.04	00		1.24	1.23	20.	1.83	1.01	1.01		1.61	1.36	50.5	00.0	96.0	00.0	60.6	200	10.0		00.0	90.0	•
		~	•	4	en i	0 1		0	6	=	2 :	2 4		16	1.		200	21	25	200	52	53	27	800	200	31	35	200	35	90	90	6	6	- 2	2	;		6 4		9	6	20	53	*	25	90	2	0 0	60	5	29

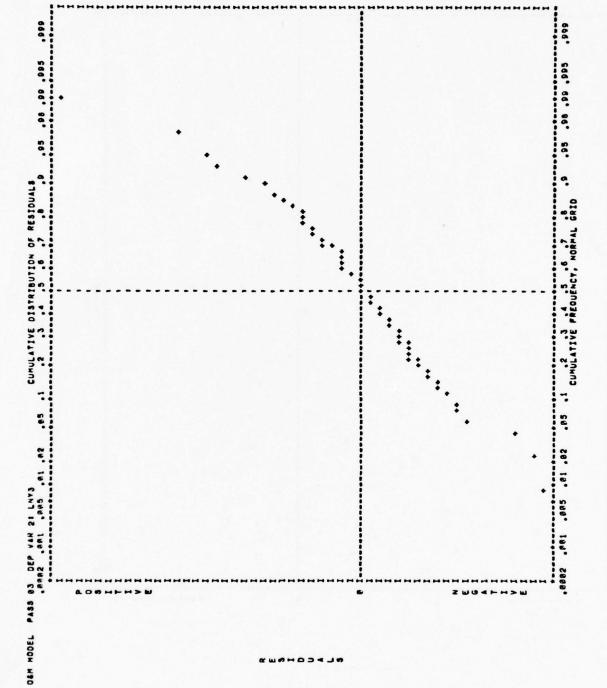


FIGURE 21

606.

« w « » D D < J »

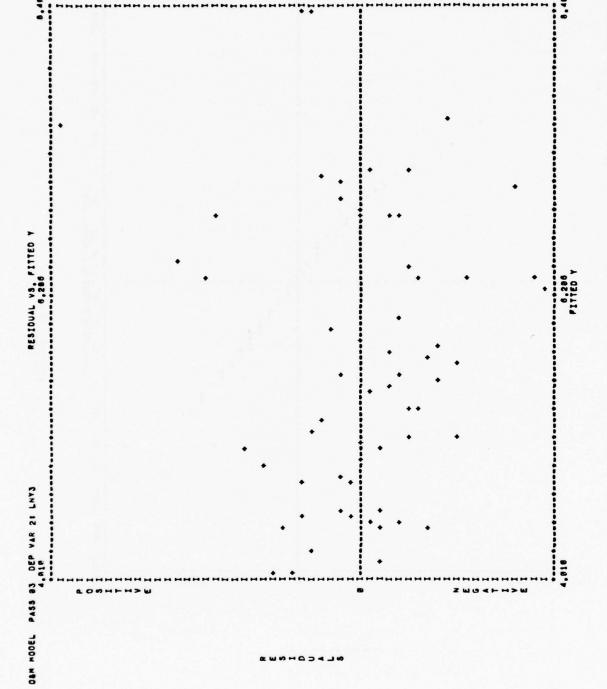


FIGURE 22

had not been in the Air Force inventory long enough to experience "good" data. These observations were dropped from the analysis in order to reduce bias.

As a means of determining whether the data are "nested", a computer program was written to sort the data table by each independent variable, $\mathbf{x_i}$. For instance, Table 4 is a printout of the data ranked by the unit price. As can be seen there are only a few observations with the same $\mathbf{x_i}$ values, and hence no evidence of serious nesting exists. If it is determined that the data are nested, two fittings must be made: one on the nested data "within plots" and one "among plots". (See [1], Chapter 8).

After the data has been critically examined, and alternate LRUs considered where necessary, we are now ready to begin the regressions. As before, a fit is made (Pass 3) where Y3 is the dependent variable (see Figures 23-27). Note that the statistics in Figure 23 have been greatly improved over that of Figure 13, which indicates the power of an outlier in the data. The fitted values plot (Figure 27), however, shows a strange trend in the residuals, that possibly the natural logarithm function can straighten out. Figures 28-32 show the results of fitting the same independent variable but with ln Y3 as the dependent variable. Here the statistics are slightly better than those of Figure 23, but the cumulative distribution plot shows an approximate straight line and the trend in the fitted values plot has disappeared.

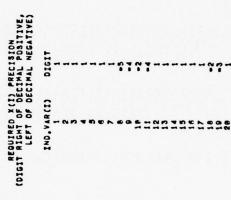
The anxious analyst might feel at this point that the use of the natural logarithm transformation in the dependent variable is the appropriate form to use the equation. It might be, however, that other transformation (curvature) in the independent variables are needed in the equation to determine the correct form of the dependent variable. Consequently, both forms of the dependent variable should be fitted simultaneously, analyzing the statistics, plots and tables at each stage and introducing curvature whenever the statistics indicate. In developing several of the fitted equations such as MMH/OH, LSC/OH, and TRAIN/OH, the statistics did not definitely

TABLE 4
RANKED BY INDEPENDENT VARIABLE X(8)

0.6.	4 >	ANGLE	6	3774.	6878	2521.	1000	1367	272	330.	3542.	1486	113.	382	9.8	1781	1359	70.	133.			186	250		233.	173	121		161	78.	22.	200	346.	124	127	139	163	21.7	284	736.			636	837	101	274	0	00.0	127	85	152	100
	x(21)	BITFIT	68.83	6.4	00.0	1.92	24.7	96.0		8.30	60.0	2.0		6.63	8.58	5 5	. 6	8	6.00	20.0		6	5.38	8 6	. 6	9.39	88.6		00.0	00.0		13.90	80.0	6.0		80.6	00.0		4.92	17.33		13.00	9.0	8.0	5 6	20.7	2.53		66.9	61.30	01.00	8.89
	x (29)	UFACT	1.38	1.30	200	2.38	1.30	1.29	1.23	1.23	2.33		200	2.30	1.30	1.20	200	1.20	1.23	1.39	200	1.20	1.33		300	1.38	1.30	200	2.3	2.30	200	200	1.30	2.30	200	1.23	2.30		1.20	2.30	1.30	200	6.83	1.30	2.0	2.33	9.83		1.23	2.30	2.39	2.39
	(61) x	POW018	17.	34.		23	7.	25.		160	.7.	850	289.	335	17.		7	200	1640.	200		60	60	592.	225.	389	1667.	173	212	256.	526		354	150	212	153	283.		851.	175.	152.	172	300	145	200	175		1623.	659	270.		. 88
	x(18)	SS	00.0		5 6	00.00	93.69	88.88		86.6	97.76	00.0	20.00	88.68	2.03	6 6	2 6		25.70	23.42	23.79	97.58	97.58	23.00	5 6	70.00	25.20	6.00	27.00	97.69	97.60	60.00	94.78	84. 78	5 5	97.98	84.99	86.00	56	33.00	66.59	26.45	98.69	100.30	99.00	73.99		00.00	98.70	80.00	100.00	
	x(11)	X X X	96.4	60.0	5 6	000	80.0	80.0			00.0	. 29	6.6		9.40	25.88		25.00	25.00	55.28	25.00	13.69	15.00		8 6	6	25.63		25.00	25.00	25.00	000	8	25.99	5 6	25.69	0	23.89		6	25.00	5 6	0.00	8	200	6	6	6 6	200	103.00	6.0	
	x(18)	S	60.0	8.0	6.0		0 . 0	6.			8	80.0	6 6		8.39	000			6	9.26			6	9.50	6 6		6.29	5 6		64.9	00.0	5 6		0.00	5 5	5	8		. 6	5	0.30	5 6	8	0.00	5 6	. 6	3.00	200	2 0	9.00	6.0	
	x(15)	χ.	80.0	64.	60.00		7.00	80.0			3.28	67.53	14.09		60.6	80.00			6	8.63			6	6	20.00		0.00	1.00		6	00.0			6	20.00	6.9	14.99	0		67.99	8	6 6	12.00	0.00	6 6	27.99	100.00	5 6		6.09	8.0	108.89
	x (14)	ANALOG	66.00	30.00	8.38		63.33	00.00	60.00		62.76	33,63	84.08	8 6 6	00.00	75.03	80.0	25.78	75.22	15.90	75.29	20.00	76.03	75.08	37.70	20.00	75.00	66.66	22.22	75.33	75.00	74.98	84.92	75.98	200	75.09	86 20	75.99		33.00	15:00	6.0	98	100.00	33.68	73.79		0	20.00	6	0.0	5 6
2023	x(13)	DIGTAL	80.0	3.30	00.0			30	1 55.0			2.50	65.0		80.8	0.30	3.35	5 6	8 6	6.00	2.23		. 6	6			6.53	5		6	6	50		00			00.0	00.0	5 6	6.0	6.00	100.00		0	55.00	5 5	2.20	9.50	87.99	8	100.00	88.8
100	x(12)	CDENS	0	1.27	. 0		2.0	.73	6.0	28.		0.24	9.50	85.0		53	0.07	61.1		64.0	9.69		0.57	0.71	6.67	2 4	8	2.11	5 5	1.22	1.22	9.47		0.71	65.0		6	5.69	8.59	4	90.0	200	1.72	61.	2.18	5 6	0	0.00	2 0	0.50	3,35	. E
2	x(11)	CCOUNT	78.	38	•	•		21.4	120	328	. 7 4	88	756.	177	924	924	50.	128	121	1153	1153.	1153	1236.	1186.	1615.	126	321	1013.	600	1674	1674.	689	946	790	412.	700	410	982	000	61.	561.	1319.	250	1457	2743.	2176	53	932	32.	1558	7638.	18.
	x(13)	WE19HT	68	. 23	.00	60.00	2	7.53	3.48	24.93	25.00	. 0	36.50	11.50	60.09		19.50	•	27.50	40.7	49,33	20.00	600	31.98	6.0			12.88			36.22	49.95	20.54	29.00	13.98	0 5	17.73	14.93	78.53	13.93	118.90	25.25		26.44	32.04	36.00	12.23	35.70	29.87	173.78	41.99	118.88
	(6) x	VOLUME	ď	36	222.	225		294	132.	300	. 203.	3.0	1276.	464.	1732	1734.	256.		FA78.		1683.	1649	1866	1680	242.		5478	479.	577	1367	1367.	1473.	. 767	1120	142.		443	377.	1377.	424	8200.	551.	307	1223.	1257.	1676.		1846.	866.	5334	2278.	3656.
	x (a)	UPRICE		. 28.		A35.		1208								_	_		4			_	30.00				5928	6247.	5257	7191	7191.	8412	67.50	9464	13040	13713	13721	14271.	15254	16570	14724.	19274	24542	31495	34462.	36013	43912	45759	57139	156230	238730	324707.
9	NO.			2.3	9.	93		200	52	3.8	0	33	9	58	4 .	37	21	6	20	. 50	10	67	4 6	65	6.8	24		32	£.3	7 40	52	en i	2 5	9 60	1.0	. 4	-	26	5 4 5	5 - 1	55	45	. 4	52	33	• (N 150	6.7	23	7.4	64	-64

LINEAR LEAST-SQUAMES CURVE FITTING PROGRAM

7.220 02 7.260-41 1.270 02 1.9 0.7630 -2.740-41 7.260-41 1.270 02 4.420 02 0.7 0.2740-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 1.270-41 7.260-41 7.260-41 1.270-41 7.260-41 7.260-41 1.270-41 7.260-41 7.260-41 7.260-41 7.260-41 7.260-41 7.27	IND, VAR(I)	NAME	COEF, R(I)	S.E. COEF.	T-VALUE	R(1) SQRD	HIN X(1)	MAX X(I)	RANGE X(1)	REL. INF. X(1)
7.400 52 4.800 52 4.800 52 4.800 52 1.6 7.400 72 1.6 8.700 72 1.6 8.700 72 1.6 8.700 72 1.700 72	-	*!*	1.165540 03	6.290 92	6.1	8.7658	-2.7400-01	7.2600-41	1.2000 00	41.6
4 4 8 70 8 2 4 4 8 70 8 2 1 1 0 8 2 8 8 2 8 2 8 2 8 2 8 2 8 2 8 2 8 2	~	K5H	5,475220 62	7,480 82	8.7	P. 8338	-2.74PD-R1	7.2640-01	1.8820 88	20.0
### ### ##############################	,	H.S.H	7,871970 82	4,890 92	1.0	9.5899	-2.58PD-P1	7.4200-21	1.8960 28	
0.00 0.00	•	XAX	-3,189710 82		6.0	P. 4288	-2.10PD-P1	7.9000-01	1.2200 00	
0 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	•	× S	2.646990 83	8.790 02	3.8	P. 3464	-2.0330-81	5.3870-81	7.4200-01	9.08
6.2 6.9 6.0 6.2 6.3 6.40 6.2 6.2 6.20 6.20 6.3 7.50 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.41 7.50 6.40 6.40 6.40 6.40 6.40 7.50 6.40 6.40 6.40 6.40 6.40 6.40 6.40 6.4	•	e x	3,923750 81	9.32D 02	9.6	0.3871	-2.1650-01	5.7350-81	7.9800-81	
4.160-63 6.6 6.7 1.29 1.5600 92 3.2400 65 3.24		x 3	-2.974550 02	8,960 62	8.8	0.4012	-2.1650-#1	5.7350-01	7.5000-01	. 9.3
2.200-81	•	0 X	3.278670-83	4.160-03	9.0	0.7129	1.5880 02	3.2400 65	3.2380 85	41.0
1.320 01 1.320 01 2.4 9.9190 1.2900 00 1.7370 02 1.7250 00 1.320 01 1.300 02 1.7250 00 1.300 02 1.3250 00 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 0.3 1.300 02 1.300 02 0.3 1.300 02 1.300	•	6 ×	2,472110-01	2,290-01	1.1	P. 8540	3.6000 91	8.2070 03	8.1700 03	8.27
1,860-#1 1,860-#1 6,8 9,540 9,970 PM 7,6340 93 7,6290 03 11,900 82 8,2 7,0340 PM 7,6340 PM 7,634	10	×10	-3.584770 AL	1,520 01	2.4	9 1 9 B	1.2000 00	1.7370 02	1.7250 02	8.83
0 31 1.900 62 8.2 9.3792 4.920-83 6.6740 99 6.6690 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	=	x 1 1	-1,457730-01	1.860-91	8.0	P. 6498	NA 0200 6	7.6380 83		8.15
0 31 1,420 82 8.4 8,9942 8.6 1,9950 9.1 1,9950 82 1,9950 92 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	12	x12	-4.891180 at	1,990 02	9.5	9.5792	4.9230-03	6.6740 09		44.6
0 01 1,350 P2 0.4 P.9093 P.0 P 1,7000 P2 1,700	13	X13	6,323970 31	1,420 02	9.0	8,9982	9.6	1.8600 82		8.84
0 01 1,390 02 0,5 0,5 0,900 1 1,390 02 1,590 02 1,590 00 0 1 1,300 02 1,590 00 0 1 1,300 02 1,590 00 0 1 1,300 02 1,590 00 0 1 1,300 02 1,500 00 1,2 0,500 00 1,2 0,500 00 1,2 0,500 00 1,2 0,500 00 1,2 0,500 00 1,500 00	:	×1.4	5.834790 61	1,340 82	7.0	6.000	0.0	1.0000 02		9.78
0 91 1,390 82 8,4 8,9959 8,8 1,6250 82 1,6260	13	x15	6.577430 81	1.390 82	8.8	1666.6	6.6	1.9090 82		8.88
0 41 1,340 62 8,6 6,5975 8,9 1,8990 82 1,8990 82 1,8990 82 1,8990 82 1,8990 82 1,8990 82 1,8990 82 1,8990 82 1,8990 83 1,8990	9	x 16	5,43P250 P1	1,390 02	4.0	9.9969	8.8	1.0000 02		9.73
0.00 5.560 00 1.2 0.0000 0.00 1.0000 02 1.0000 02 1.0000 02 1.0000 03 1.0000 03 1.0000 03 1.0000 03 1.0000 03 1.0000 03 1.0000 03 1.0000 0 0 1.0000 03 1.0000 0 0 0 1.0000 03 1.0000 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	-1	X17	8.023820 M1	1,340 82	9.6	8.9975	6.6	1.8990 82		1.87
0-01 5,010-01 1,0 9,5659 7,000 00 1,640 03 1,6330 0 0 3,920 02 0 0,0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	9.	x 18	-6.856420 00	5,580 00	1.2	9.6589	8.6	1.0000 02		9.89
0 01 1.600 01 1.0 0.4 0.8000-01 2.3200 00 2.0000 0 0 0 0 0 0 0 0 0 0 0 0 0	6	x18	-5,165a90-81	5,910-01	2.9	9.565P	7.6980 88	1.6490 83	1.6330 83	9.11
0 01 1.680 01 1.0 0.5878 0.0 6.1940 01 6.1960 01 1.1960	50	acx.	2,228960 02	5,920 92	7.0	0.8297	3.8880-81	2,3340 88		90.00
	2	x51	1,736950 81	1,680 81	1.0	8.3878	8.8			8.14
	0 0F 089ER	VATION		62						
	SIDUAL DEG	REES DI	FREEDOM	. 67						
2	VALUE			1.5						
	SIDUAL ROG	IT HEAN		971.42229661						
	STOUAL MEA	N SOUA		945,54481195						
FRIDUAL SUM OF SOURRES	SIDUAL SUR	1 OF SQ		*********						



VAR 11

PASS 3

HODEL

 $\begin{array}{c} (1) & (1)$

NEIGHBORING OBSERVATIONS (OBSERVATIONS 1 TO 4 APART IN FITTED Y ORDER). 1971.42 RESIDUAL ROOT MEAN SQUARE OF FITTED EQUATIONS LINEAR LEAST-SQUARES CURVE FITTING PHOGRAM 6 ESTIMATED FROM RESIDUALS 2 DEP VAR 11 STANDARD DEVIATION PASS 3

1300m #80

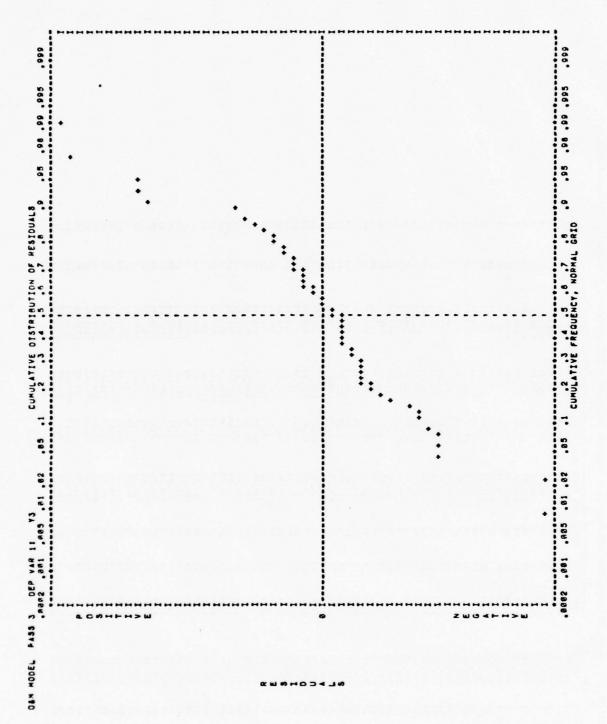


FIGURE 26



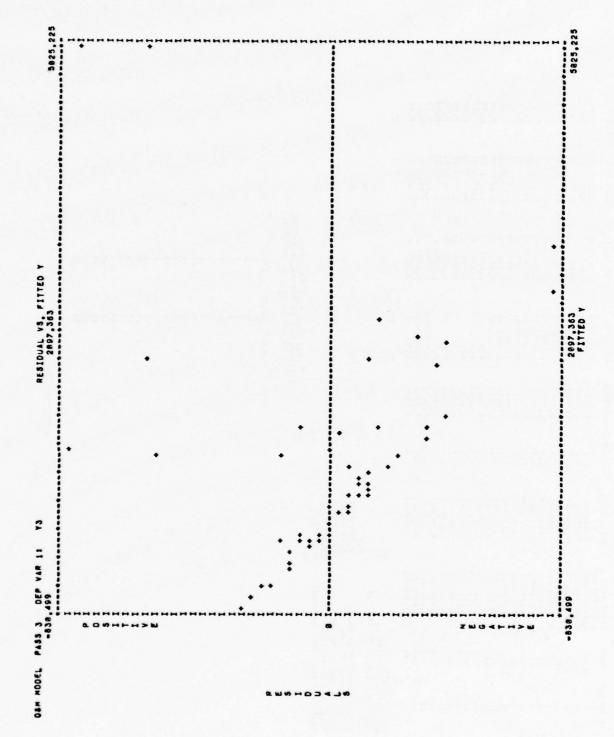


FIGURE 28

REL. INF . X (1) RANGE Y . 5.1890 RANGE X(I) 88 Y . 8,9270 HAX X(I) MIN Y . 3.8180 89 MAX LINEAR LEAST-SQUARES CURVE FITTING PROGRAM MIN X(I) R(I)SGRD S.E. COEF LNYS VAR 2: NO. OF DOSERVATIONS
NO. OF IND. VARIABLES
RESIDUAL DEGREES OF FREEOOM
FESTIONAL ROOT MEAN SOUARE
RESIDUAL SOM OF SOUARES
FOTAL SUM OF SOUARES
MULT, CORREL, COFF, SOUARES 066 PASS HODEL ...

COIGIT AIGHT OF DECISION FIGURED FOR STATIVE, LEFT OF DECIMAL PROSTITIVE, STATIVE, S

GAN MODEL PASS 3 DEP VAR 2: LN+3

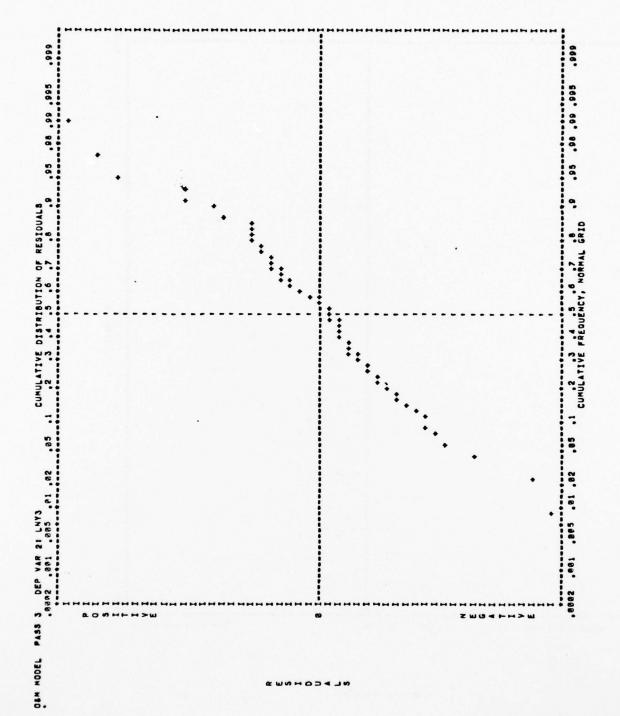
560	-	~	•	•	*	•		0 0		:=	: 2	2	-		2	1	0 0		2 .	22	53	54	52	5.0	57	58			35	33	3.4	33	60		000		7	42			. 4	13	•	9	8.		20	3 4	33	36	37	99	8	6.0	
ORDERED HESTO.	1.846	1,632	1.476	1.014	1.889	8.7.8	0.715	0.00	200	707.6	9.418	9.411	9.396	P.378	8.361	166.4	2010		240	6.236	6.212	0.176	6.133	-2.608	-6.534	900.00	20.65	0.042	- 8.652	685.6	-9.186	-0.118	-9.137		-0.169	-9.173	-A.198	-8.215	-6.249	800.00	907	-0.327	-0.377	-0.411	9.430	757.81	0000	10.507	-6.704	-0.737	-6.761	-0.779	66.6-	-1.992	1000
ITTED Y	6,326	6,201	6.741	4.785	5,156	3.925	4.437	200	8.425	6.683	4.665	8.425	6.839	4.898	4.343	101.	140	1000	800	6.672	3,556	4.689	4.836	6.808	5.654	4.772	800.	3.562	5,773	5,334	4.686	5.952	2.218		6.766	6.648	5.044	4.571	6.039	4.946	0.00	6.272	7.107	107.0	4.57	0.001	7.386	4.415	4.947	5.833	5.691	6.144	5.557	5.689	
085, Y	8.172	7.832	8.217	5.799	6,166	4.683	5.155	70.7	8.027	7.177	5.084	8.836	7,235	5.276	4.704	. 483	2000	208.4	5.612	6.997	5.767	4.784	4.936	6.848	5.656	4.765	7.2.7	5.528	5.023	5,245	4.588	3.833		20.00	6.691	6.475	4.846	4.357	9.4	4.648	5.276	5.945	6.739	2.689		20.0	0.00	3.618	4.243	300.0	4.849	9.368	5.667	4.784	
0880	•	,	20	56	•	4	2 •		150	. 6	45	93	36	9	200		5 -		. ~	15	30	57	99	4	35	75		62	9	3	17	80	2 4	2.5	=	52	3.6	-	c 4	32	8.7	:	35	200	\$ 6) es	: sc	22	27	-	7	2.	35	• :	
RESIDUAL	-0.737	0.245	-8.137	-1.095	-0.467	1.009	200-1	000	9.307	-8.163	.9.779	9.236	-9.108	9.513	1.476	100	20.00		-1.537	-0.704	0.295	1.014	-0.023	6.809	-2.411	80.00	300	-8.298	0.316	8.507	164.6	-8.751	200	-8.327	9.418	.0.052	0.758	90000	975.8	A. 502	-9.377	0.411	600.00	0000	22.0	-8.138	9.212	-9.459	-8.215	P. 942	0.715	000	66.258	401.4	
FITTED Y	5,833	5.368	5,218	5,880	5,361	5,156	102.0	6.235	5.954	6.766	6.144	6,672	4.686	4.280	6.741	61.0	1000	6.548	6.265	4.947	4.597	4.786	4 300	6.557	5.491	44.0	900	4.945	5,291	961.9	6.583	2.091	7 . 78	6.272	4.666	5.075	3.925	5.581		A 425	7.187	8.425	6.838	1.580	4 4	4.958	5.536	4.571	4.571	5.562	4.437	5.334	912	0000	
	5.896	5.612	5,981	4.784	4.893	9	8 172	2.546	6.271	6.531	5,366	6.947	4.589	4,795	7 485		4.00	6.475	4.728	4.243	4.892	5,799	4.277	5.667	8 8 8 8		7.235	4.548	5.607	1.384	7.177	2 4 4	7.143	5.945	5.0.94	5.923	. 386.	5.276	0.00	8.927	6.738	8.836	0.00	0.010	4.784	4.822	5,757	4,113	4.357	5.528	5.153	5,245	20.0		
SS DISTANCE		-	:	•	:	: ,	: .		,		-	,			. <u>.</u>		. 4			2.	18.			, i	,,	• 0	: .	. 2	•	•	•	•			•	16.	28.	. 40		14.	3.	14.	•	•••	: .:	. 2	3.	2.	2.	2.				: .	
		~ .		•	n	•	. ••	a	1.6	:	13	£ .	2	9 6	2 .	22	24	52	56	27	28	50	31	35	25		36	37	38	30			4.5	7	64	9			8	15	55	25	4 4		22	90	60	69	6	29	200	9 6	99	:	
	11829	25.55	21.00			90.4.4	72460	71653	PIFAB	71583	71040	71 A B E	1140A	13084	73CEN	73CFK	73DAH	73EBA	73E6F	71CAB	72EAA	72ECA	72889	71748	72000	72459	71716	71310	72850	7288	72938	745	74910	76119	76640	7455	TAPAB	745119	77EC?	77EE0	770CA	77086	7354	7: 240	11000	63440	65448	638AA	63544	65944	41919	63459	63448	41101	

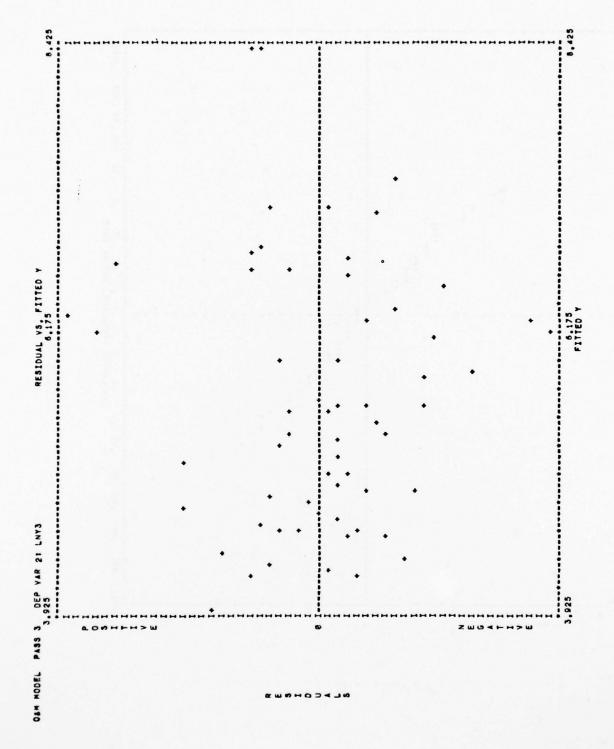
LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

	ORDER).
8.89	APART IN FITTED Y
OF FITTED EQUATIONS	COBSERVATIONS : TO 4
RESIDUAL ROOT MEAN SQUARE OF FITTED EQUATION: 0.88	IGHBORING OBSERVATIONS
	STANDARD DEVIATION ESTIMATED FROM RESIDUALS OF NEIGHBONING OBSERVATIONS (OBSERVATIONS : TO 4 APART IN FITTED Y ORDER).
OLM HODEL PASS 3 DEP VAR 2: LNY3	STANDARD DEVIATION ES

SEO.	-	~ -	,		, • 0	^	•	o,	6	-	15	2:	: :				6	50	21	22	23	24	8	0 10	200	0 0		31	35	93	3.4	33	90	37	000	9 6	7	42	3	7	4.5	9	•		9 4			33	3.6	92	36	22	80 9	9 6		
08 SV.	47	٠ د مه	0 .		20	63	91	8	20	22	90	0 1				40	37	27	80	3.4	4	•	,	200	•	0	33	0.00	62	4	24	7	33		• ;	0 -			. 0.	26	4.4	•	9 6	200		. 4	. 6	11	30	2	90	25	25	3 5	3	3 2
FITTED Y			82.		. 4.	4.44	4.57	4.57	4.68	10.4	19.4	19.4					00.	4.93	4.96	5.84	2.07	5.16	5.22	3.29	2000	200		90.00	5.56	5.58	5.59	5.68	5.65	8.83	80.0			. ~	6.21	~	~	2	2.			80.00		6.77				7.11	7.15	7.10		
FSIDUALS	1.01	9.77	40.8		17.	80.0	P.24	8.75	0.12	0.42		0.52	-:				. 4		90.0	P.15	1.06	1.15	0.45		90.0			20.0	20.00	8.16	8.62	9.76	6.73	6.36	86.8	2 6			0.12	1.21	2.17	2.38	0.33	8.72			79.	8.67	8.51		0.77	0.71	8.38			
	49.13	7.82	9.0		2.83	1.24	9.6	7.79	10.99	6	8.79	12.71	21.93	26.33	200	24.40	6.0	3.45	6.74	5.12	12.63	1.23	8.44						10.69	13.11	2.58	4.58	1.33	7.82	12.08	20.0		69	1.0	1.34	1.22	5.93	67.7	. 42		200		9.91	7.68		2.85	24,55	1.28			
DEL RESIDUALS	8.6	0.42	9.54			1.65	A.21	45.4	14.6	1.24	67.0	3.29	8.59	2.45	1.01	2.6	2.24	9.67	3,38	1,33	0.12	1.96	8.63		90.00	200			. 6	1.48	2.17	9.67	1.15	1.17	6.0		. 6	20.00	9.44	1.07	80.0	1.09	80.0	60.00	00.00		90.6	1.13	1.38	0.40	0.72	60.0	2.37			
desv.	51	20	6 6	, ,	'n	23	œ	53	5	8	54	0		62.	0.0	2 6	9 60	39	80	4 4	56	4 4	96	0	6	'n	7		30		0	4	n	0	-	200			53	51	93	40	2	6	-		4	2	67	36	52	6	53	40		3 0
0884	23	25	6	,	2 -	52	^	43	52	2	4		D			. a	90	=	56	œ	œ	^		8	6	5 0		: 0	; ;	. •	77	52	ø	63	63			43	3	35	23	30	5	7 .	2 .	9 0	2 -	35	35	54	35	79	32	5 O		::
1880	6.0	9.				6	40.0	00.0	41.0	0.23	9.27	9.00	29.62	8.73				3.91	2.93	1.99	1.84	1.89	1:1			-:		9		12	1.22	1.22	1.23	1.24	1.24	1.24		1.26	1.25	1.26	1.25	1.26	1.27	0	1.20	2		1.36	1.49	1.5	1.42	1.47		80.		1.02
STO DEV	6.19	6.28	6.53	200		65.0	6.55	0.55	40.9	6.62	9.61	10.0		• • • • • • • • • • • • • • • • • • • •			.35	1.32	1.43	1.44	1.38	1.41	1.36	15.1	1.28	1.20	22.			1.15	1.18	1,17	1.17	1.19	1.17				:-	1.1	1.11	1,12		=:					1.09	1.08	1.00	1.95	1.89			100
.04	-	~	,		n «		•	0	10	:	15	2	9 1	2	0.		0	53	21	25	23	54	25	59	57	0 0			22	23	34	35	36	37	99	3		. 6		;		46	47		9 5			2 20	34	22	28	23	80	2 6		







pointout which form of the dependent variable was appropriate until the final iterations were made.

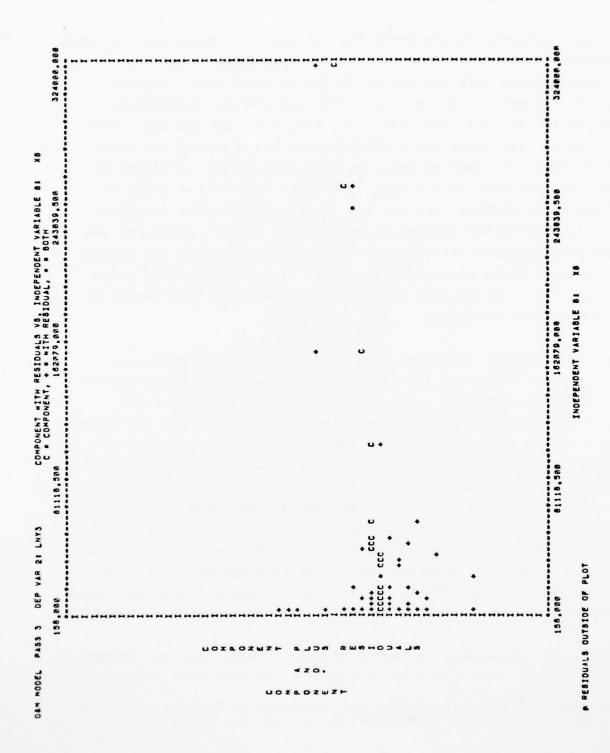
After much analysis in the MTBMA fit, it was discovered that In MTBMA is the more appropriate form of the dependent to be used. We will then proceed with the sketch of the ln MTBMA fit. Figures 33-42 are component-plus-residual plots for the ten independent variables X8, X9, X10, X11, X12, X16, X17, X19, X20 and X21. Figure 33 shows that there are 5 observations which extend the range of X8 by 300%. It must be determined whether these observations behave like the rest of the data (thereby extending the range of variable X8) or whether they are not consistant with the remainder of the data (possibly indicating curvature). Table 5 shows the ten independent together with the observations which extend the respectives ranges. These observations numbers can be obtained by using the component-plus-residual plots and the tables of data ranked by each independent variable.

Indicator variables in conjunction with the $\rm C_p$ -search technique can be used to determine the effects of such extended observations (see [4]). The approach is to define indicator variables, X22 through X31, denoting those observations which extend the variables shown in Table 5, then multiply these indicator variables by each of their respective independent variables, and use the $\rm C_p$ -search technique to determine if any of these interactions are significant. If any of these products prove to be significant, then curvature will be introduced in the variables, since we have assumed that 7 initial indicator variables sufficiently describes all qualitative information about the observations. This is displayed in Figure 43, where for instance X22 is an indicator variable indicating observations 49, 31, 46, 47 and 34 and X22X8 is the product of X22 and X8.

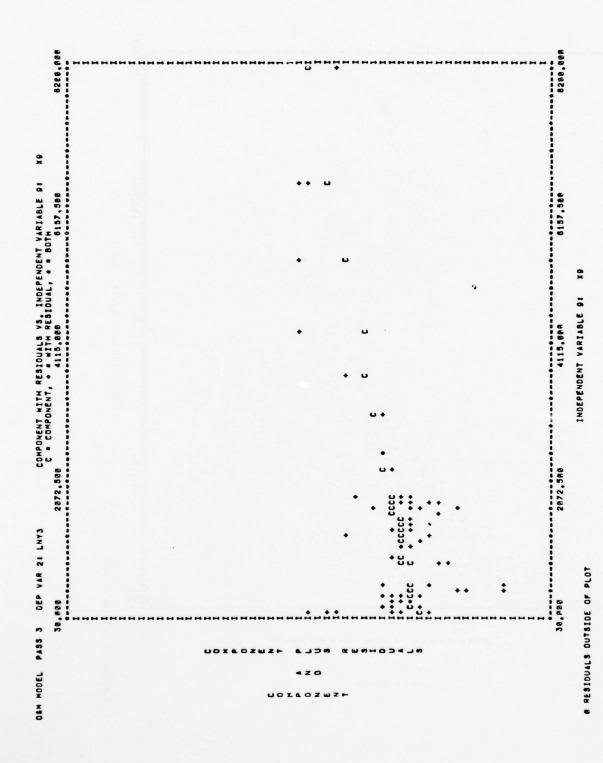
There are 31 independent variables shown, of which only one belongs to the "basic set," and leaves 30 variables to be searched. The approach

[&]quot;The Use of Individual Effects and Residuals in Fitting Equations to Data," F. S. Wood, Technometric, 15, No. 4, (1973), pp. 677-695.

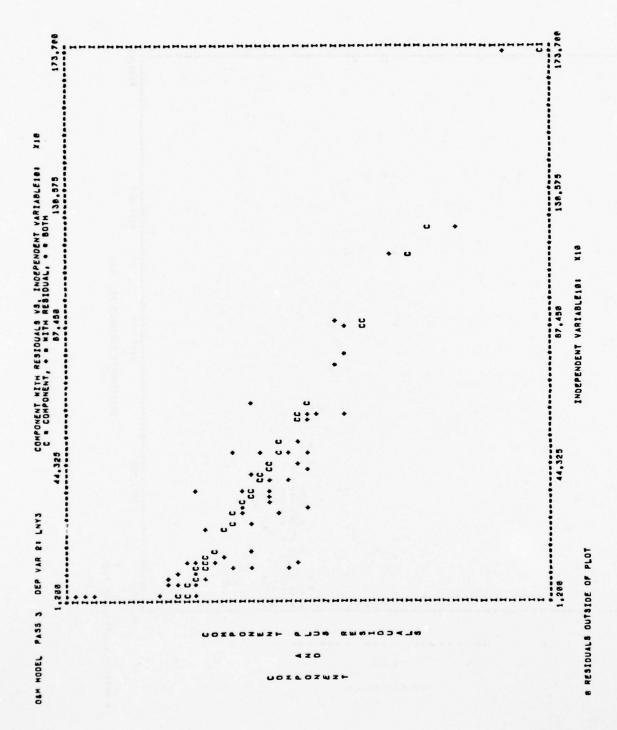




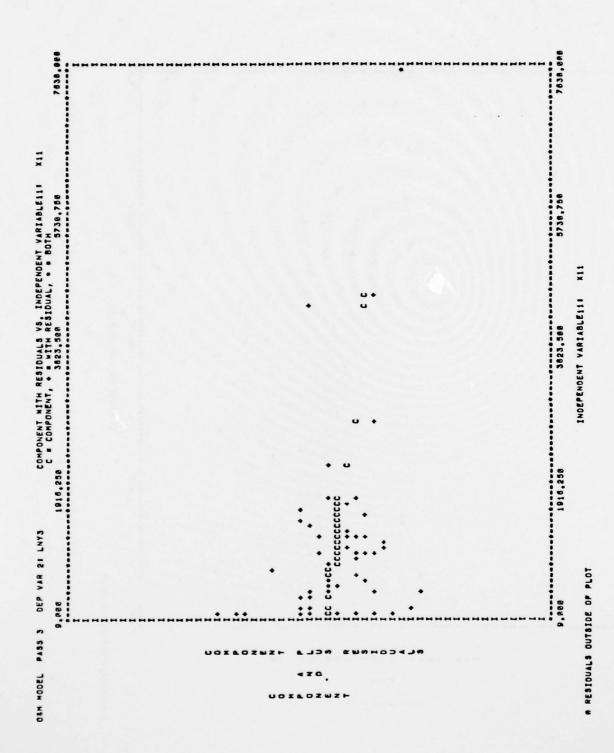




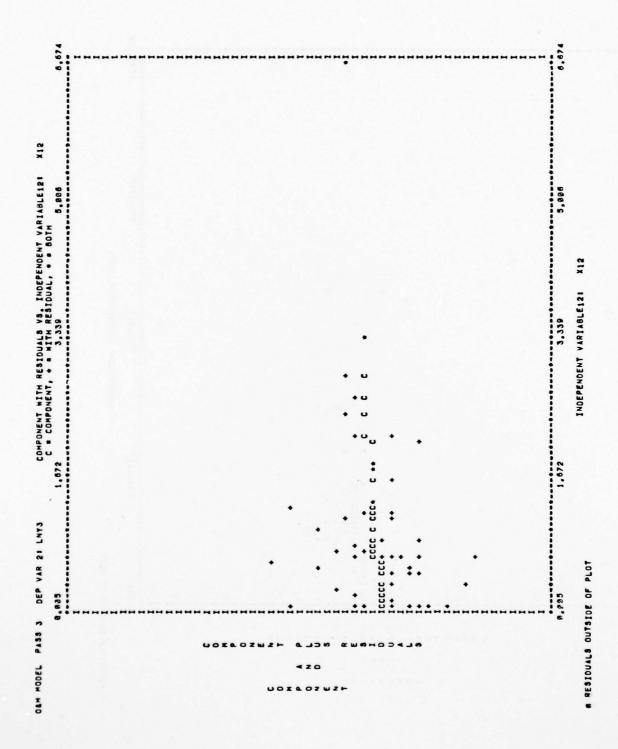












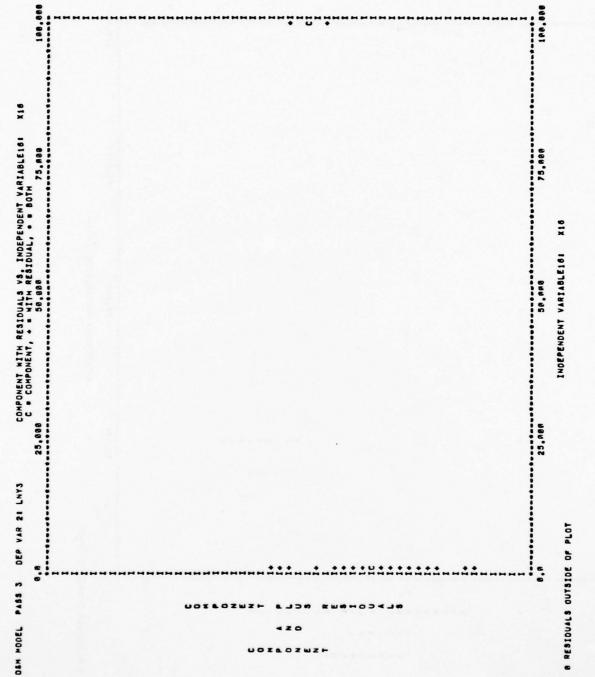
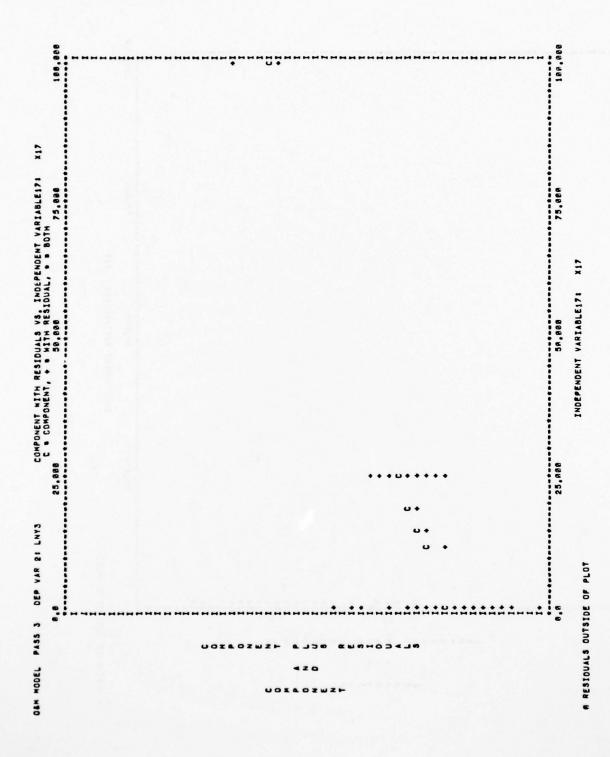


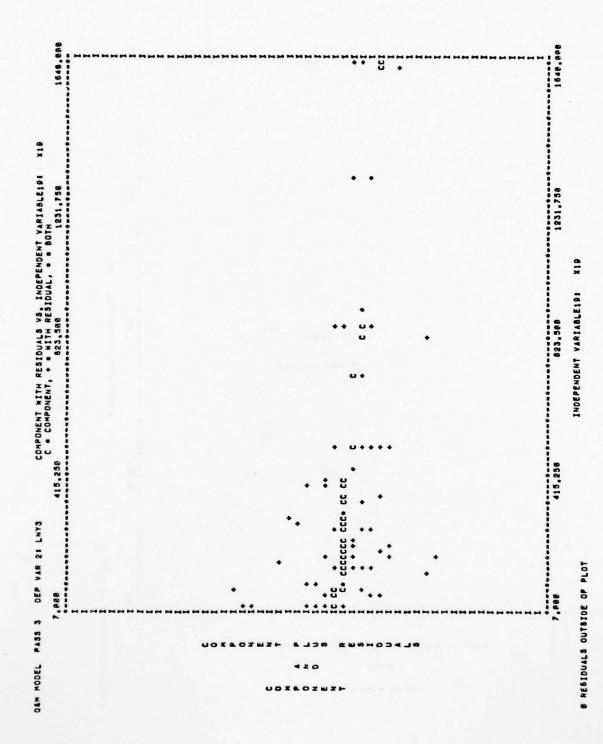
FIGURE 38

UOI402WZF

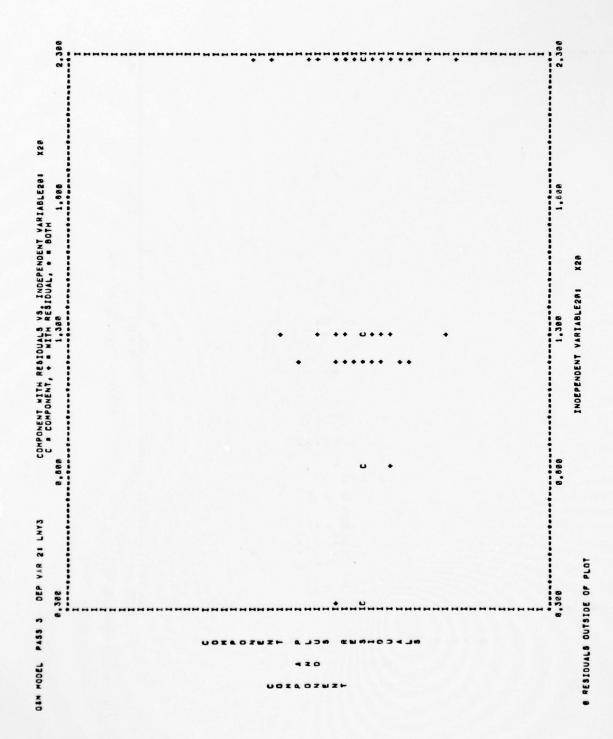


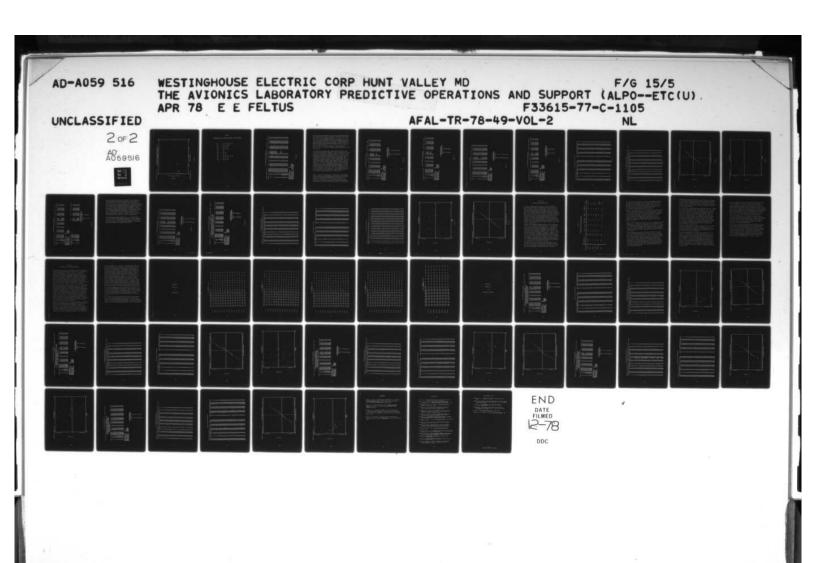




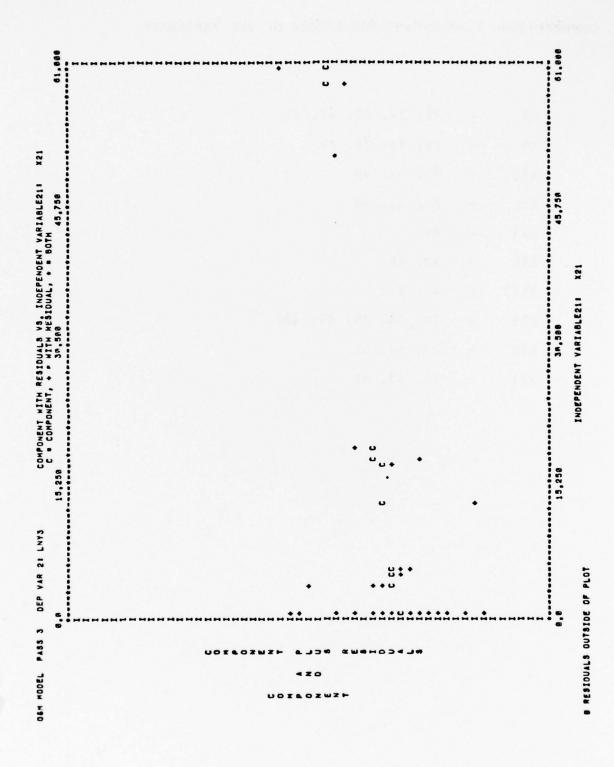












OBSERVATIONS THAT EXTEND THE RANGES OF THE VARIABLES

TABLE 5

x8 - 31, 34, 46, 47, 49 x9 - 18, 22, 28, 47 - 22, 47, 49 X10 - 29, 31, 46 X11 68 X12 - 38, 48 X16 - 42, 47 X17 - 18, 28, 45, 46, 48 X19 - 45, 54, 55 X20 - 46, 47, 48 X21

FIGURE 43

		UEF VAN 18	r.		HIN Y	4.554D B1 MAX Y	. 7.5340 83	RANGE Y . 7	7,4840 03
IND.VAR(I)	NA N	COEF.BC	S.E. COEF.	T-VALUE	R(I)SGRD	MIN X(I)	MAX X(I)	RANGE X(I)	REL, INF, X (1)
s -	**	2 323420 8	92 1.970 93	6	9 0707	7400-01	7 2690-01	66 0666	6
۰ ۵	X SH	5.695550 0		. 6	9.9825	-2.74PD-84	7.2600-01		
	KUX	5.351760 8	32 5.250 92	1:1	9.6476	-2.5800-A1	7.4280-81		26.0
•	H T			9	8.8824	-2.1880-81	7.980-01		90.0
•	×			3.7	8.4584	-2.6330-01	5,3870-31	7.4280-81	6.33
•	8 ×	1.897330 8		-:-	0.5481	-2,1650-R1	5,7350-01	7.9830-81	0.12
	×	-4.16747D @		8.8	0.5796	-2.1550-01	5,7350-61	7.9000-61	9.00
•	8 ×	-7.584970-83		4.0	2,9883	1.5820 92		3,2380 85	0.33
•	6 x	-4.257710-91		4.6	0.9859	3.6900 01			97.0
1.0	x 10	-1.314570 a		4.0	6.9819	1.2080 88		1.7250 82	8.33
=	× :	-4.926573-01		0.0	9.9659	D. 0000.0	7.6380 03		0.50
12	x 12	4.422950-01		5.	9.9876	4.9230-03			9.99
2	× .	2,162440 @	1.670 92	r.	0.000	6.		1. 2000 82	2.92
= :	× :	2.072710 8	1.030	?	9666	8.0			2.77
2	C .	2.199900 0	1.030	201	8 9 9 9 4	8.0			16.5
9	x 1 6	2.852840 3		6.	8666.8	8.0			2.74
12	x17		75 1.610 32		8 6 6 E	0.0			2.72
	8 X				8.8791	0.6			9.15
0.	x 1 9			2.1	8206.9	7.0000 68			9.45
000	XSA			0.1	8.9865	3.8880-81			5.45
2	x 51	-4.43291D @	91 4.630 91		0.9542		6.1980 81	6.19aD PI	9.36
25	X22XB	9.969170-83		6.0	8.9992	8.6			6.43
23	x53x8	4.949140-01	3.470-01	7:1	P.9523	6.0	8.2400 B3	8.2000 83	9.54
24	X24X18	-2.771320 36			0.9727	6.			9.00
52	X25X11			o.	9.0599	6.6			87.8
58	x25x12			9.9	9.8399	6.6			0.10
27	x27x16		1.430		6.9976	0.0			0.11
88	x28x17		91 2.430 81	0.0	0.9151	6.9		1.0000 32	8.31
50	X29X19		1.530 83	6.6	9.95A4	0.0	1.6480 83		8.38
38	X30X23		03 4.170 83	9.0	D.9684	0.0	3880-		9.26
31	x31x21		81 6.120 81	8 .8	0.9723	6.6	6.1380 81	6.1630 61	8.39
NO. OF 085E	RVATION	•	62						
NO. OF IND. VARIABL	VARIAB	ш	31						
RESIDUAL DE	GREES O	F FREEDOM	39						
			3.1						
DESTOUR MEN COLLED	NAME TOOM	SGUARE	983,18789583						
RESTOUT SUM OF SOU	10 F	ABES							
TOTAL SUM OF	F SOUAR		9 Y Y P B C C C C C C C C C C C C C C C C C C						
MULT CORRE	COFF	SOUARE	7697						

used to search these variables is to utilize the fractional replication technique twice, where 12 variables are searched at each stage. First of all, the 12 variables numbered 3, 7, 9, 12, 16, 17, 18, 21, 23, 26, 27, 31 have the smallest t_i -values (see Figure 43). The CP search technique indicated that none of the 12 variables searched are significant enough to remain in the equations. These variables are therefore dropped and the remaining variables are fitted (Figure 44). Next, the twelve variables 1, 2, 3, 5, 6, 8, 9, 10, 12, 13, 16, 19 are put through the C_p search technique. This time only variables 3, 6, 10, 12, and 13 are significant and the results are printed in Figure 45.

We note that the variables X22X8, X24X10, X28X17 and X29X19 remain in the resulting equation. Therefore, curvature in the form of squares and natural logarithms for variables X8, X10, X17 and X19 are introduced into the regressions as shown in Figure 46. Since X17 is the % transmitter, there is no logarithm used. In this pass, there are two variables in the basic set. Since there are obviously some variables of negligible influence (see the T-VALUE and REL. INF. X(I)) a search must be made. Again, using a double fractional factorial search, the Cn-search technique admits only the 12 variables shown in Figure 47. (Pass 39). Additional outputs of Pass 39 are shown in Figures 48-51. Since the residual route mean square = .43, the cumulative estimates of the standard deviation indicates that there is little evidence of lack of fit. There are no observations far from the "centroid" of all observations. The cumulative distribution of residuals plot is now a straight line and the fitted y values have a nice even scatter about the 0 - residual line.

There is, however, one observation, 9, which has a larger (smaller) residual than all the other observations (Figure 48). This observation may be controlling the estimates of the coefficients. To determine if observation 9 is in fact controlling some of the coefficients, a cross verification of coefficients is performed as shown in Figure 52, where the statistics at the top are calculated

LINEAR LEAST-SOUARES CURVE FITTING PROGRAM

IND. VAR(I) NAME COEF.RIT)	OFH HODEL PA	PASS 33	DEP VAR 11	. 43				4.5500 B1 MAX Y .	Y . 7.5380 83	RANGE Y . 7.4840 03	.4840 03
E COFF. R(I)											
748330 82 8.910 82 8.9 9.90 92 8.1 9.90 92 92 9.00 92	ND.VAR(I)	NAN	COEF. B(T)	••	S.E. COEF.	TOVALUE	R(1) SORD	HIN X(1)	HAX X(I)	RANGE X(1)	REL.INF.X(I)
1.2vanto a2 9.830 a2 8.1 1 2.940-p1 7.9993 -2.740-p1 7.9990-p1 1.99p0-p1 1.99p0-p1 7.99p0-p1 7.9	-	H11	7.483330 82		8.910 P2	8.6	P.9154	-2.7 4PD-P1	7.2500-01	1.0000 98	8.18
-2.549780 P2	~	2×	1.243810 02		9.830 02		2.9375	-2.7400-01	7.2600-01	1.3000 00	9.92
3,574770 83 7,350 92 1,8 8,3251 -2,1530-81 5,1870-81 7,4200-81 7,92430-81 7,4200-81 7,4200-81 9,2430-81 7,4200-81 9,2430-81 7,100 62 1,200	•	14	-2.039780 P2		3.830 92	8.8	0.4599	-2.1000-01	7.9000-61	1.8200 23	6.00
7,423450 PP 7,130 62 1,8 8,2440 P2 3,7350-81 7,2360 P2 3,2360 P2 3,2361 P2 3	•	× S	3.574270 03		7.350 92	0.7	0.3251	-2,0330-01	5.3870-01	7.4200-01	B 35
-3,024303-PP 1,210-PP 2 2,5 9,9752 1,5670 82 3,2400 PF 5 3,2380 PF 5 -2,51420-PP 1 1,7370 PP 1,7		×	7.423490 02		7,130 62	1.0	8.2448	-2.1650-#1	5.7350-01	7.9000-61	8.98
-2.451420 A1 3.4650 A1 3.4	•	8 ×	-3.024303-02		1.210-02	2.5	9.9752	1.5600 02	3.2480 85	3,2380 #5	1.31
-2.251420-81 3.860-81 8.7 8.9957 9.9980 88 7.6390 83 7.6590 83 1.6	•	x18	-2.451830 AL		8.130 00	3.3	A.7365	1.2900 00	1.7370 72	1,7250 02	6.57
-4, 172350 PR	•	x 1.1	-2.251420-81	-	3.860-01	0.7	9.9957	9.0960.09	7.6390 83		R.23
-3.451950 00 8 1170 00 8 1770 00 1 1 1 4 8.8120 00 1 1.02.00 00 1 1.02	•	-	84 02KC74. 4-		1. 4.9D A1	4.6	P.7837	8.6	1.0000 02	1.0000 02	90.0
1,274767 31 8,430 80 1,4 8,430 80 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6420 82 1,6330 83 1,6330 83 1,6330 83 1,6420 82 1,5330 83 1,6420 82 1,5330 83 1,6420 82 1,530 83 1,6420 82 1,530 83 1,6420 82 1,530 83 1,6420 82 1,530 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,6420 83 1,642 83 1,6420	19		-3.451950 00		8.170 69	4.6	8.8478	6.0	1.8390 92	1.8940 82	0.05
-1,043490 PP 6,970-01 1,5 0,9482 7,0960 MB 1,6470 R3 1,5330 23 3,1930 PP 8,516 2,0430 MB 2,0430	:	2	1.204760 91	•	8.430 98	•	8.8192	6.6	1.00000 02	1.0000 02	8.15
3,710240 31 8,650 82 8,8 8,976 81 2,3870 87 2,0730 88 2,2730 87 2,27310 87 2,2731	13		-1.443490 PR		6.970-01	1.5	0.8382	7 6960 98	1.6420 83		P.23
2,92940-02 1,190-02 2,4 8,9768 8,8 1,7270 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,7370 62 1,9700 73 1	13		3.719240 91		8.650 02	8.6	0.9425	3.0980-81	2,3830 88		20.0
### 1,242970 81 8,390 cm 1,5 8,7472 8.8 1,7370 82 1,7370	14 X2	2x8	2.9229RD-02		1.190-02	2.4	0.9769	9.6	3.2480 95		1.27
1 2.602130-81 2.740-01 0.9 0.8530 0.8 7.6350 7.5550 7.5	15 x	01x70	1.242070 01		8.390 68	1.5	8.7892	0.0	1.7370 02		6.29
7 1,939,00 01 1,10 01 1,6 9,675 9,2 1,970 02 1,970 P2 9 1,5610 03 1,5620 03	16 x2	5x11	2,692130-01	•	2.890-01	6.0	9.8938	8.6	7.6380 93	7.6380 93	4.27
9 1,30166 08 6,150-P1 2.4 8,709 P.8 1,6400 03 1,640 B3 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	17 X	18x17	1.930000 01	•	1.100 01	9.1	8.6475	6.0	1,0000	1.8000 02	6.26
A 7,77310 82 1.910 83 8.4 8.891 8.8 8.3880-81 8.3880-81 ABLES ABLES ABLES OF FREEDOM 45 9 ARR SOLUME 9292215581297 ARES 12115412.51804668 EF, SQUARE ARES 12115412.51804668		61×60	1.501860 00	_	6.150-01	2.4	9.7989	0.0	1,6480 03	1.6400 93	0.33
ABLES AB		Seres	7,777310 82		1.910 03	4.4	8,8391	9.0	8.3880-81	8.3000-01	6.0
ABLES OF FREEDOM 42 42 42 42 43 43 43 43 44 44 44 44 44 44 44 44 44	0. OF DBSER	ATIONS		.62							
OF FREEDOM 42 AN SQUARE 91.0.4636039 UARE 829.22.15581297 SQUARES	0. OF IND.	ARIABL	63								
AN SQUARE Q13 - 63.56 Q18 - 63.63.63.63.63.63.63.63.63.63.63.63.63.6	ESTOUAL DEGR		FREEDOM	42							
EAN SQUARE 918, A163 A39 OUARE 82922.15561297 SQUARES	-VALUE				•						
OUARE 829222.15581297 SQUARES		ī	SOUARE	613	. A1636939						
SQUARES	PESTOUAL MEAN			19:22	15581297						
UARES 121154112,61844668 OEF, \$QUARED ,7125	ESIDUAL SUM	OF SQU	•								
COEF. SQUARED . 7125	OTAL SUM OF			141,12.	61804668						
	ULT. CORREL	COEF	SOUARED		7125						

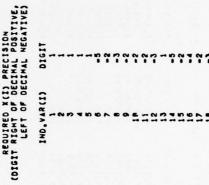


FIGURE 44

FIGURE 45

LINEAR LEAST-SOUARES CURVE FITTING PROGRAM

	. P. C. C. V.	DEL VAN 11 LATS	2			NA	MAX Y = 6.9270 68	MANGE T = 5,1890 88	. 1680 68
		•							
IND.VAR(I)	STAZ	CUEF.B(I)	3.E. COEF.	T-VALUE	R(1) SQRD	HIN X(I)	HAX X(I)	RANGE X(I)	REL.INF.X(I)
	MAX	-5.238320-01	2.190-01	2.4	0.2343	-2.1880-81	7.9000-01	1.0330 93	
~	×	1.435930 09		3.3	P. 1327	-2. P330-91	5.3870-01	7.4200-01	
•	8 ×	-1.941260-05		2.9	9.0539	1.5880 92	3.2490 95	3.2380 05	
•	× 10	-3.571470-02		2.2	7454	1.2040 80	1.7370 82	1.7250 02	
50	X 1 4	9.117490-03		2.2	9,7218	0.0	1.0000 62	1.8480 42	
•	x 1 5	1.847840-92		4.7	N.6173	0.6	1.3600 62	1.3800 62	
^	61 x	-7.418110-84		8.1	P. 7830	7. 8880 88	1.6400 93	1.6330 03	9.24
•	85×	-2,459780-91		1.6	0.1575	3.0900-01	2.3380 83	2.0000 68	6.18
o	x22x8	1.951840-35		3.1	8.9618	6.0	3.2490 93	3.2430 85	1.24
6	x24x10	1.428590-02		2.8	P.7224	6	1,7370 92	1.7370 02	9.40
=	x28x17	2,635640-02		3.0	0.5728	0.0	1.6640 62	1.6000 02	0.52
12	x53x19	1,316130-03		3.8	0.7044	6.0	1,6400 03	1.6400 83	0.42
10. OF 085	DF DBSERVATIONS	•	62						
NO. OF IND. VARIABLE	VARIAB.	LES	12						
RESIDUAL DEGREES	EGREES OF	FREEDOM	64						
F-VALUE			15.9						
PESIDIJAL RI	ROOT MEAN S	SOUARE	9.61431589						
RESIDUAL ME	MEAN SQUARE		9.37738398						
RESIDUAL SUM OF SQUA	UM OF SO	UARES	18.49181129						
TOTAL SUM OF SQUARES	OF SOUAR	ES	98.47129848						
MULT. CORREL	EL COEF	SQUARED	7956						

LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

OEM HODEL PASS 35 DEP VAR 11 LNYS

AIN Y . 3.8180 AM MAX Y . 8.9270 AM RANGE Y . 5.1890 AP

NO.VAR(I)	NAME	COEF. B(I)	8.E. COEF.	T-VALUE	R(I)SORD	HIN K(I)	MAX X(I)	RANGE X(1)	REL.INF.X(I)
	* : *	-1.278570-01	3.270-01		8301	-2.7470-01	7.2600-01	1.0000 33	0.03
~	x 5 m	-2.957250-42	3.890-91		9.88.6	-2.7400-01	7.2690-91	1.6000 23	8.61
•	K 3 H	2,495000-01	2.200-01	1.0	0.6312	-2.5AUD-P1	7.4200-01	1.0000 00	46.6
•	× 4 ×	-1.637500-01	2,530-01	9.0	8.57A1	-2.1030-01	7.9200-01	1.8220 98	8.63
•	K.S	1.733450 66	4.360-01	4.9	8.5578	-2.2339-81	5,3870-01	7.4280-81	9.25
•	5×	5.932030-01	4.140-01	*	9.4347	-2.1650-41	5.7359-81	7.9965-81	00.6
	x7	-2.699470-01	4.050-91	6.7	6.4579	-2.1650-91	5,7350-01	7.9030-01	3.74
	6 ×	2,175770-05	1.450-04	3.2	8.9379	3. apan as	8.2020 #3	8.1720 93	60.4
•	*11	2.125230-86	1.140-34	6.0	9.8252	9.0440 98	7.4380 03	7.6250 03	66.6
10	x12	3.546370-82	1.030-01	5.8	8.7124	4.9230-83	6.6740 93	6. 6690 99	80.6
:	x 1.3	9.675790-69	6.660-22	1.5	3.9945	9.6	1.9990 92	1.30.00 02	1.89
12	x 1.4	9,737530-62	6.410-22	1.5	8 9994	9.6	1.9900 92	1.2000 92	1.98
13	x15	1.175890-81	6.460-92	9:1	6.9992	6.0	1.9390 92	1.0000 02	2.38
	×16	9.451130-02	6.490-82	6.1	A.9974	8.6	1. SPAD B2	1. 8430 92	1.85
13	x 1.6	7.431270-93	3.470-83	2.4	8.78K7	6.0	1.8400 02	1.8000 02	6.15
16	x 2 A	-1.862210-01	3,450-01	9.6	3.9813	3.9930-81	2.3900 88	2. 9990 99	9.97
17	xoı	1.717530-02	1.190-32	•	9.84A3	6.	6.1000 01	6.1940 91	P.21
1.8	x 8 x	1.682670-06	3.740-96	3.4	6.0339	-2.7389 84	2,9650 85	3.2380 85	
	I & I X	1,339110-02	1.220-22	1:1	9.9759	-3.5ag0 A1	1.3750 92	1.7250 P2	6.45
28	X17H	9.657330-92	6.400-02	1.3	8.9978	-1.1100 01	8.8990 91	1.0000 92	1.89
21	x101x	8.200900-05	4.580-94	8.5	8.9938	-3.75AD A2	1.2580 03	1.6330 03	58.8
22	X BOSO	-7.284920-12	2.920-11	9.2	9.8161	5.7750 AB	3.4170 19	3.5660 19	60.6
23	DSOKIX	-1.225820-64	6.930-05	:	9.8672	2.9580 00	1.1790 84	1.1790 84	9.28
24	X170SQ	1.498990-94	1.530-64	1.0	0.7439	2.3410 92	3.5640 93	3.3360 63	6.18
25	8190SQ	4.864520-07	6.880-07	2.0	B. 8173	4.8470 92	8,2990 85	8.2940 #5	86.9
90	LNXA	-3.671710-81	1.570-91	2.3	9.9391	5.8630 98	1.2690 01	7.6260 68	8.55
27	L~x19	-9.152840-91	2.530-01	9.5	0.9491	1.8230-91	5,1570 08	4.9750 98	8.89
	0 . XW					. 0460 .	4 4000		

10-010-1	200	8.46124574 P.21274763	7.72F67195 98.47129848 .9224
18-008-09-11-	OSSERVATIONS IND. VARIABLES	4 SQUARE	SQUARES UARES OEF. SQUARED
	SERVATION OF VARIAB	ROOT HEAN SHARE	SUM OF SOUR
	NO. OF OF	FSIOUAL FSIOUAL	TOTAL SUR

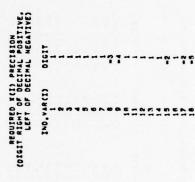


FIGURE 46

FIGURE 47

LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

1300m m80	PASS 39	DEP VAR 11 LNYS	8 24		* × × × ×	3.8180 00 MAX	MAX Y . 8.9270 88	RANGE Y . 5.1890 PR	. 1890 08
IND.VAR(I)	M 4 7	COEF, B(I)	S.E. COEF.	T-VALUE	R(1) SQRD	HIN X(I)	MAX X(I)	RANGE X(I)	REL.INF.X(I)
	DDDIIIOR POCONDIIOR VXXXX VXXOJODOIIIORI VXXXXX VXXIIX VXXXXXXX VXXIIX VXXXXX VXXIIX VXXXXX VXXIIX VXXXXX	2.022270-00 1.0600440 00 1.745460-00 2.660570-00 1.846460-00 1.846960-00 2.726770-00 2.726770-00 1.84560-00 1.84560-00	0.000000000000000000000000000000000000	~ *** *** *** *** *** *** *** *** *** *	\$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$ \$	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4	1,2880 P8 7,4230-91 7,520-91 1,500-01 1,000-01 1,000-01 1,1700-92	608686668888 60-10-640-608 60-4-0869088
NO. OF OBSER NO. OF OBSER RESIDUAL DEG RESIDUAL ROD RESIDUAL WEA RESIDUAL SUM TOTAL SUM OF WUIT, CORREL	NO. OF OBSERVATIONS NO. OF IND. VARIABLE RESTOUAL PEGREES RESTOUAL ROOT HEAN RESTOUAL ROW OF SOUARE TOTAL SUM OF SOUARE MULT, CORREL, COEF.	FREEDOM SOUARE FARES SOUARE	00.00000000000000000000000000000000000						

(DIGIT RIGHT OF DECIMAL POSITIVE, LEFT OF DECIMAL NEGATIVE)

01611	_	-	-5	-5	-5	**	7	•	-	
IND.VARCES	~	•	•	•	•	•	•	61	==	12

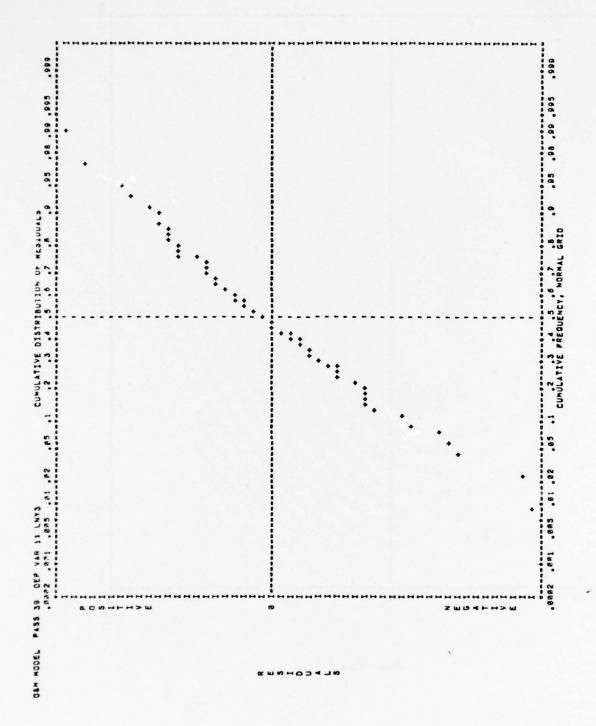
1300m #90

RESIDUALS.... DEP VAR 11 LNYS ANCE

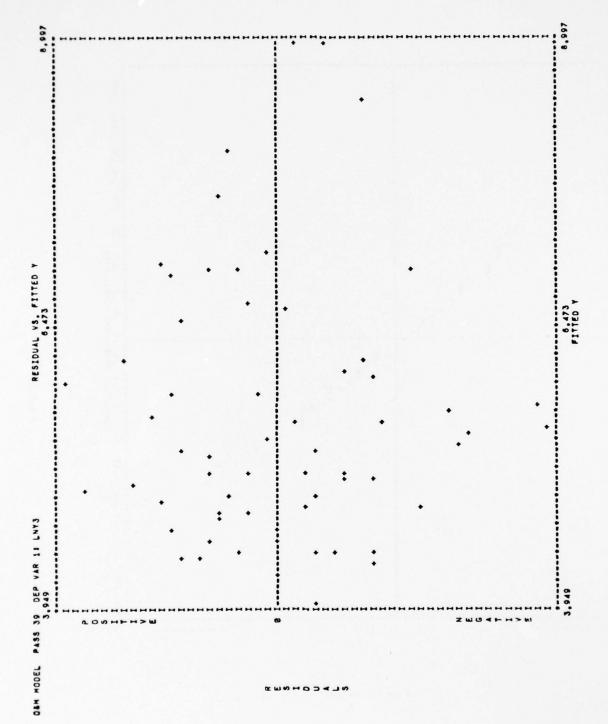
LINEAR LFAST-SOURRES CURVE FITTING PROGRAM

	Y ORDER).
6.43	APART IN FITTED
RESIDUAL ROOT MEAN SOURRE OF FITTED EQUATION: A.43	STINDING DEVIATION ESTIMATED FROM RESIDUALS OF MEIGHBORING DASERVATIONS (OBSERVATIONS : TO 4 APART IN FITTED Y ORDER).
EP VAR 11 LNYS RESIDUA	THATED FROM RESIDUALS OF NEIGHB
DEM MODEL PASS 39 DEP VAR 11 LNYS	STENDARD DEVIATION ES

	CUMULATIVE	05.5	ORSV	صيب		MSS0	ESIDUALS	550	DEL	RESTOUALS	FITTED Y	0884.	SEQ.	
_	1000			٠.			60	7.52		0.15	3.93	22		
			25	•	36		42	8.24		9.39	4.20	31	~ ^	
					6	6	.24	2.01		9.67	4.36	0	,	
		6	27		37	8	.36	6.83		8.65	4.49	20	• •	
	95.0	3.31	65	• •	67	6	.67	00.0		0.57	4.44		, w	
	64	A.21	99		26	6	.63	5.68		00.0	4.40	,	,	
	3.37	8.21	99	•	57		.23	6.0		6.54	4.40			
	200	4	6	-	32	•	.76	2.11		0.52	87.7	0	0 0	
	3.0	30	87	•	5.8	6	54.	2.18		8.12	4.58	10		
		000	26		58	•	.37	6.83		9.22	4.33	2		
		99 6	25		13	•	.22	3.74		6.15	4.68	40	::	
					27	•	.15	45.10		0.41	4.68	0	21	
	200		2		22	6	. 52	36.27		8.24	4.69	47	2	
	0000					-	8	5.95		0.13	4.79	46	:	
	10.0		•;					11.1		9.13	4.93	99	- 2	
	9.41	5	2 6		,			6		3.33	4.85	33	91	
	8.42		,				34	6		0.42	4.88	37	17	
	8.41	66.0			- :					0 0	68.8	5.6	9.	
	0.43	1.0	28		200	•					4.02	13	19	
	6.42	1.97	43		2		25.				4 07	88	8	
50	9.44	1.49	33		6		.70			20.0		63	21	
21	6.43	1.99	33	•	22		.33	200				000	22	
22	8.43	61.1	•		9		.30	7.31		61.0			16	
		1.1	99		33		1.13	0.63		8.03	A	2	2 2	
2 2		2.5	-		8		69.6	7.78		9.11	5.11	77		
			2				1.18	7.35		6.33	5.13	28	52	
2					:			34.84		9.35	5.16	8 0	50	
50	8	10.1	2 5		3			49.34		0.10	5.18	4	22	
27		1.33	0 1					200		4	5.18	4.4	88	
58	8.49	1.00	23		3		200				5.19	24	58	
50	0.43	1.72	2								5.35	~	80	
38	6.43	1.36	37		0		60.00					99	31	
31	0.42	1.87	99		2		60.00					0.00	35	
32	0.43	1.96	22		23		6.53	2.50					33	
33	44.4	1.96	36		2		76.8	1.20		2.0			7	
3.6	8.44	2.03	77		35		9.44	0.0		0.0		:	5	
35	****	2.98	27		69		P.13	11.55		10.8		•		
36	0.43	2.38	27		6			200				44		
3.7	8.42	2.89	4		_		01.0	14.13		20.00			4	
	6. 42	2.11	ē		37		0.27	6.43		10.0	00.0			
0	2 42	2.11	69		37		P. 52	4.42		1.15	20.0	۰.		
, ,	24.0	2.16	89		2		9.62	9,33		9,32	0.70	-		
; ;		2.18			33		0.17	17,32		1.05	5.76	•		
::					-		277	2.24		0.32	5.87	;	42	
42	2.5		-				22	3.87		8.42	5.87	6.	43	
3			5 6				2.0	3.67		1.19	66.6	24	44	
4							120	1.32		6	6.93	32	45	
							. 6	12.22		9.83	6	*	97	
9	20.00	2,50	9 -				99	2.58		9.80	6.16	25	47	
•		20.0					6.57	29.95		89.6	-:	28		
	20.0		3					12.39		8.38	٠.	13	•	
2							4.43	4.96		8.12	٠.	11	80	
0		44.0	2				5.13	7.14		9.28	6.71	52	51	
							20	15.73		86.6	16.9	39	25	
25							25	4.64		9.75	6.08	52	23	
200	200							21.75		9.11	6.00	36	9.6	
		2.30	7 6		**			1.07		8.32	7.80	5	52	
6		10.0	0 6				. 6	23.84		8.43	7.93	21	26	
	2.0				3 8		96.6	2.49		01.0	7.15	4	27	
6								11.78		8.93	7.62	^	2.0	
28	0.42	2.73			:		200	7.7		44	8	23	9.0	
20	27.0	2.73			-					7.0	87.8	•	60	
69		4/.2	2		. 4			6		0.5	84.0	53	01	
9	8.42	2.00					107	11.22		9.24	99.6	91	62	
62	8.42	2.30	22		•									
								FIGURE		649				







LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

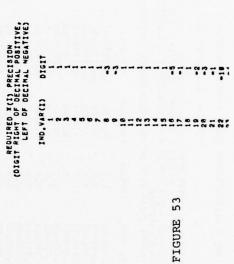
5.1890 99	REL. INF. X (1)	96.8	8.25	9.19	9.34		66.6	87.6	8.29	0.13	9.12	9.35	00.00						REL, INF, X (I)						95.6	9.08	9.36	2.0		5.50										
RANGE Y . 5.	RANGE X(I)	1.9900 88	7.4280-41	7.90.00-01	1.0000 02	.1.0000 02	6.1030 01	1.7250 02	1.1790 84	3,3380 03	8.2940 PS	7.6260 88	4.9750 68						RANGE X(I)			7 4230-91	7 0000-01	1 0000	1.0000 02	6.1940 91	1.7250 82	1.1790 94	8.2940 95	7.6260 98	00 /00	J.O								
8,9270 00	MAX X(I)	4200-0	5.3970-01	7350	COND	0000	1000	3750	1790	5640	2990	1.2690 01	.1570						MAX X(I)			7.4200-81	7445-01	00000	1.9000 02	6.1ceb P1	1.3750 62	1.1790 64	8.0000	1.2690 61	0,010	OF DECIMAL POSITIVE, OF DECIMAL MEGATIVE)	01617	-		- ?	~ ~	27	19	
3.81.80 08 MAX Y	4IN X (I)	-2.5800-01	-2.0330-01	-2.1630-P1	6			-3.5880 PI	2.9580 63	2.3410 62	4.8480 82	5.9630 99	1.8230-01						MIN X(I)			2.5880-91	20.000.00			9.6	-3.594D 91	2.9580 98	4.8470 92	5.9500 98	10-0020-1	COIGIT RIGHT OF DI	IND.VAR(I)	-	~	n •	in ic		. a. <u>e.</u>	
6 • × × × × × × × × × × × × × × × × × ×	R(I)SORD	9.3874	0.2303	A.2117	0.7052	7064	9.5373	8033	9.7486	9 47 RB	9.3479	6.6518	0.9148																			ē								
	T-VALUE	8	200	2.3			2.1	2.6	4.1		2.5	80	e.																											
22	S.F. COEF.	1 600-01	3.270-91	3.320-21	130-27	2.450-33	6.410-33	5.330-93	6. VBD-93	1.010-24	3.410-97	6.260-02	1.830-01	252		R. 43229898	9.18587542	98.47129848																						
DEP VAR 11 LNYS	COEF. R(1)	0 0000000	1.689440 66	6.5294RD-01	744490-02	5 665500.03	7.840460-03	1. 437940-32	8.732620-25	2.022770-04	7 634990-97	-2.377390-61	-1.015570 80	S	FYEEDOR	OUARE	940	SQUARED	COEF.B(1)		9.413860 22	2.498700-81	85 05 50 50 50	10-00-02-00	3.885.420-83	6.652120-23	1.058120-02	-5.55234D-85	2.434320-84	-1.768740-R1	16-09460-61									
08H HODEL PASS 39	R(I) NAME	***	, m	×	*: *	× ×	x21		X 130SO	X17050	05001X	LNXB	NX 10	NO. OF DESERVATIONS	RESIDUAL DEGREES OF	RESIDUAL ROOT MEAN	RESIDUAL MEAN SGUAR	TOTAL SUM OF SOURES MULT, CORREL, COFF.	IND.VAR(I) NAME	*** COEFFICIENTS FR		Inx								LAKB	2									

using the entire set of data and the statistics at the bottom are those calculated with observation 9 omitted. Here we note that the coefficient, for X18 has decreased by nearly 50%, and the coefficient for X21 has decreased by nearly 15%. Sometimes the effects of the observations on a certain coefficient are lying hidden beneath other effects, and hence the effects on the coefficients for X21 may be larger than the statistics indicate. Therefore, in addition to considering curvature for the previous variables X8, X10, X17 and X19, we consider curvature for X18 and X21. If curvature is not needed in variable X21, the C_p search technique should omit it. This exercise is shown in Figure 53. Again a fractional factorial search performed twice, yields the results in Figures 54-59.

The statistics are highly significant, $R_Y^2 = .9183$ and F-VALUE = 41.5, the residuals are now evenly distributed, (indicating constant variance of $\sigma^2(y)$), the cumulative distribution is a straight line (indicating normality) and the cumulative estimates of the standard indicates that there is no evidence of lack of fit.

LINEAR LEAST-SOULARES CURVE FITTING PROGRAM

	3 4 4 2	1.980700 AM	S.E. COEF.	T-VALUE	R(I) SQRD	HIN X(I)	HAX X(I)	RANGE X(1)	REL.INF.X(I)
-	×1×	2.1019RO-02	3.160-01	6	1878 0	20 7490-01	7 3690-91		
~	KON	-8.200570-02	3.630-91		. 6	2 7 7 2 2 2	100000	20000	
n	T C X	1.446340-21	2.230-01		67774	2 5800.91	7 4290-91	00000	7
4	X 4 H	-1.242770-31	2.420-01	S	6.53	6-008-0-	10000 1	00000	200
•	×	1.665340 20	160-01		2000	1300000	100000	2000	74.4
					0 0 0	14-0220-41	2.38/0-61	7.4280-81	6.24
		100000000000000000000000000000000000000	2.20		6.5182	-2.1655-41	5,7350-01	7.9780-01	5.13
	11	2.22.530-01	4.550-61		0.5145	-2.1650-01	5.7350-01	7.9030-01	6.63
•	o ×	-4.969950-95	1.560-04	5.5	0.0521	3.0990 91	8.2PPD #3	6.1700 03	80.0
•	× 1 1	1.642470-94	1.230-04	8.0	8658	9,3930 00	7.6380 93	7.6290 03	0.16
0	x 1 5	-9.957230-03	10-020-1	4.1	0.7242	4.9230-03		6.6690 88	8.81
	x13	1.016210-01	6.330-92	9.1	P. 9985	8.6	1.0000 02		1.99
12	× 1.4	10-0-2250-1	6.120-22	1.7	4000.0	9.6			2.65
2	415	1.219540-01	6.160-02	5.8	8.9992	8.6			2.39
	x 1.6	1.052140-01	6.190-02	1.7	A . 9974	6.0	-		2.86
13	x59	-6.578370-e4	3.800-01	0.0	8688.6	3.0000-01			
•	1	2.379550-97	3.690-96	0.1	0.9392	-2,7380 04			
17	26.7	1.750570-02	1.350-02	1.3	9.9838	-3.5930 P1			
	X17M	9.865670-92	6.080-32	1.6	0.9978	-1.1190 91			101
0	1 0 1 x	6.834960-93	2.940-93	2,3	0.7943				
26	×101	2.679310-94	4.430-84	9.0	8.99.9			1.6330 93	
21	x 51 x	-1.87219h-83	1.390-72	8.1	6.9302				000
22	XBUSG	-1,514550-11	2.790-11	6.5	8.8188				
53	XIVOSO	-1.624500-05	9.650-95	9.1	6.8983				
54	X17050	1.210070-04	1.490-84	8.0	8.7579		7 4640 AN	1 1100 01	
52	x18050	2.217340-04	1.370-44	2.1	6233		20000	-	
96	DSC61x	7,831790-07	6.770-27		8.23				
27	x21050	-6.98P420-04	4 850=84				CB 04.67.0		21.0
88	LNXB	-3.33858D-91	2000			1000000		Ca Orca.	4
58	LNXIS	-8.2669RD-91			0.00		1.2000 01	7,0200 88	9.29
	OLX	-5.28423D-92		? .	4000	1.8230-01			6.91
					*****	00 0016.1	94 0794.	2.43/0 68	66.80
. OF 085E	RVATIONS		25						
NO. OF IND. VARIABLES	VARIABL		33						
RESIDUAL DE	GREES OF		-						
PESTONAL BOOT MEAN SOURCE			14.8						
RESIDUAL MEAN SOURRE	AN SOUAR	344000	9.43089387						
RESIDUAL SUM OF SCUAPES	M OF SEU	IAPES	5.91472622						



			LINEAR LEAS	IT-SOUARES	CURVE FITT	LINEAR LEAST-SOUARES CURVE FITTING PROGRAM			
1300H HT0	-	DEP VAR 11 LNYS	173		MIN Y . 3.8180 88	. 8160 88 HAX Y .	. 8.9270 00	RANGE Y . 5	5,1890 88
		LNY3 = 8(P) + 8	FUR THE REGRESSION ANALYSIS FOR THE TALPOST HODEL B(9) + B(1)XUH + B(2)XAH + B(4)XS + B(4)XG + R(5)X19 + B(6)X11	. VSIS FOR 1	THE "ALPOS"	HODEL R(5) X1P + B(6)	K11		
			+ B(7)x15 + B(8)x184 + B(9)x180s0 + B(19)x190s0 + B(11)x210s0 + B(12)LNx8 + B(13)LNx19	1(9) X18050	+ B(10)×18	050 + B(11) x210	0		
		LNYS . LN(MTBMA)	LY(HIBHA) MIBHA (MEAN TIME BETWEEN MAINTENANCE ACTIONS)	IN TIME BET	THEEN MAINT	ENANCE ACTIONS			
IND.VAR(I)	NAME	COEF.B(I)	S.E. COEF.	T-VALUE	R(1) SORD	HIN X(I)	HAX X(1)	RANGE X(1)	REL. INF. X (I)
-	KON	3,155280-91	1.480-01	2.1	R. 4053	-2.5890-91	7.4280-81	1. PRBD 48	98.88
~	H W	-3,135960-91	1.450-61	2.2	0.2875	-2,1020-01	7.9PAD-01	1.0000 08	94.0
•	ec ×	1.453710 00	2.980-81	6.1	0.2353	-2.0330-01	5.3870-01	7.4200-01	0.21
•	0 ×	6.311760-91	3.940-01	2.7	A.2261	-2.1650-71	5,7350-81	7.9000-01	9.13
•	×10	1.176660-62	3.370-93	3.5	F.7878	1.2000 00	1.7370 82	1,7250 92	9.46
•	x11	6.266690-05	5.060-05	1.0	9.3614	9.0000	7,6380 03	7,6290 83	9.12
	x13	1.758380-02	2.720-03	6.5	0.6786	0.0	1.4980 62	1.0000 02	0.35
•	X 1 8 F	6.500370-03	2.120-03	3.1	9.6748	-6.1880 81			6.13
•	x19050	1.864360-84	6.710-85	8.8	0.2325	7.7260 01			9.18
<u>e</u>	x19050	7.326610-97	2.830-67	2.6	0.2173	4.8400 82	8,2990 05	8.2940 AS	0.12
=	x21050	-4.839340-94	3.240-84	1.5	P.3814	7.8390 81	1,1380 83	1.8590 #3	0.10
12	LNXB	-2.838P50-P1	5.850-82	6.4	0.6602	5.8630 88	1.2690 01	7.6260 88	0.42
13	LNX10	-8.398660-81	1.110-01	7.5	0.0191	1.6230-01	5,1570 00	4.9750 88	0.02
NO. OT 1806 RESIDUAL OF PENIOUS OF RESIDUAL RO RESIDUAL RO RESIDUAL RO ROTAL SULA OUT OF THE SULA OUT	OBSERVATIONS IND. VARIBLE L OEGREES OF L ROOT HEAN L NEAN SOURE L SUN OF SOURE OWNET. COFF	SOUARE SOUARE SOUARE SOUARE SOUARE SOUARE	60 41.5 41.5 7.002.401.51 8.118.40.69 7.004.06916 90.4712.0948						



LINEAR LEAST-SOUARES CURVE FITTING PROGRAM

60.0	D DEVIATION ESTIMATED FROM RESIDUALS OF NEIGHBORING OBSERVATIONS (OBSERVATIONS : TO 4 APART IN FITTED Y ORDER),
. v	2
A 7 10	-
EGU	NOI
TED	VAT
F.1.1	BSE
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OUAR	TION
2	FRVA
ME	088
RESIDUAL ROOT NEAN SQUARE OF FITTED EQUATION:	ING
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V .	ATED
DEP	STIM
EL RUN 3 DEP VAR 11 LNYS	ON
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860	(NP	•	•	6	•	o a	10	=	12	2	1:		-	=	2	8	2 2	200	2 .	2	50	2	8 8	6	32	20	7	8	3	000	; =	7	?;	3	3	3 :	•	4	8	5 6	3	5	6	6	8	•	6
083V.	25	6.0	10				22																	: 4						•	8 0	; -	62	6		25	35	: 3	::	6	en e	9	2	9	:		••	93
FITTED Y	4.03	9 -	4.38	4.38	4.43	16.4		79.7	10.1	4.06	4.74	6.79	20.			4.91	3.94	3.18		200	3.18	9.10		3.0	5.27	20.00	5.39			5.41	14.0	9.00	9.78	6.81		6.12	9.14		9	6.63	9.78	7	7.15	7.16	200	7.68	4 9 7	22.0
EL RESIDUALS																								6.43	98			60	80.6	0.36	٥	200		9.91	00.0		9.16	2.5		6.24	n		0.26	9.12	5.6		70 0	
1880	10.95	20.0	6	9.44	14.11	13.23	00.00	6.14	12.74	6.33	6.25	6.39	2.0		000	18.79	24.38	26.19	13.36	20.5	1.81	1.49	24.72	24.43	11.85		12,82	15.28	3.62	6.49	13.09	9.00	12,61	23.24		7.81	1.76		16.28	29,25	37.11	000	25,15	3.03	10.07	12.84	* .	0000
RESIDUALS	8.80	2.22	8.67	9.36	0.32	50.0	9.40	70.6	00.0	9.66	6.29	7.44	79.67		200	6.13	9.30		7.0		0.20	8.21		6.27	6.25	40	0.00	6.0	8.92	8,23	60.0	20.00	8.93	10.0	0.0		78.8	9.32	12	8.93	1.22		9.00	89.6	9.50	900	**	0/10
#330.																																																
RED BY													<u>.</u>	2 5	-	3	27	- ;	3 -	- •	;	-	25	3.5	99		2	60 F	200	•	5		2	8	6 6	9	5	-;	, 4	2	35	90			9	36		2
083V.	53	6		27	99	21		3 -	9	6	13	2	- ;	2 2		-	11	90			35	4	60 0	~ ~	•	e «		5	~	•		200	•	2:	66	9	69		90	30	•		:	30	::	200		300
1850	8.8	6.6		8.05	9.26	60.0	6		2.0	1.00	1.49	1.63	1.56				1.05	1.94	- 0	2.0	2.39	2.54	5.00		2.00	20.0	3.65	3.19	3.23	3.34	8.38	20.0	3.69	3.71		3.72	3.72	3.7	200	3.96	6.9		4.29	4.27	6.37			
STD DEV	60.0	500		9.36	9,35	16.9	6,00			6.43		9.44	57.8					9.39	6.30			6.43	7 .		6.43		9.40	90					0.42	17.0	2.5	27.5	9.42			6.40	9.42			8.42		77.0		77.6
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LINEAR LEAST-SQUARES CURVE PITTING PROGRAM

DEM MODEL RUN 3 DEP VAR 11 LNYS

100	DENT	088V.	MOS DISTANCE		FITTED	KESTOOME		0890	-	-	,
		-		5.896	5.584	987.6-	•	6.166	5.393	8.772	
		~	•	5.612	5,162	8.459	•	6.686	6.962	0.746	
		•		5.681	4.638	8.111	3	5.153	4.537	9.616	
		•	.0	4.784	2,397	219.8-	25	6.738	6.122	86.6	
			~ .	266.	0.400	215.30					
		•	• • • •	0010	2000	241.6		6.5			
				172	8.263	000	•		4.638	77.0	
11		•		246	5.338	-0.784	50	5.799	5,437	P.362	
11. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0. 0.		1.0	,	6.271	6.883	0.186	9.	4.795	4.434	P.361	-
13. 4. 4. 5. 3. 8. 6. 1. 8. 1.		=	•	6.691	6.486	8,193	30	5.767	2.407	B.358	_
15	_	2		5.356	5.181	9.104	21	7.485	7.153	6.332	-
1			10.	6.997	6.629	8.278	9	7.384	7.825	0.278	-
2		-	•	1.588	1.641	196.60	2	6.097	6.629	0.278	-
		•	:	4.795	4.434	8.361	24	9.458	5.217	0.233	-
2	_	50	53.	8.217	8.672	8.143	3	979.7	4.048	6.296	-
	,	2	24.	7.485	7.153	8,332		4.936	4.738	80.0	-
22		25	20.	9.818	4.835	-8.217	=;	100.0	00.0	500	-
2		2	•	80.0	2.217	8.233	5.			201.0	- '
2		200	25.	6.473	6.787	2:5:4		177.0		_	•
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2			••	27.50		175.91	2.	0000	7 677		•
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2			•		020.	000		200	8.972	201.6	
10.00			•			9 169		2.278	07.1.6		•
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10 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1		: 2		200		174		6.683	4.621	9.652	
100					5.182	. 1.85	93	6.818	6.781	9.937	
100	. 6			4.846	679	9.00	67		5.254	9.821	
10.00		60		9.55	5.181	8.169	•	5.63	5.649	-9.617	
10. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.		36	10.	7,235	7.162	9.973	75	4.765	4.785	-9.929	•
100	0	33	7.	4.648	4.877	-0.229	10	4.397	4.382	-0.625	-,
10. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1.		90		5.687	5.415	8,192	•	7.177	7.227	40.00	-,
4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4		30	13.	1.334	7.025	0.278	80	4.611	4.675	19.00	•
# # # # # # # # # # # # # # # # # # #		9	•. •.	7.177	7.227	800.6	2:	4.588	100.	198.4	
4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4		=:	•		5.263	-8.425		0.827	9.0		
4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4				00.	60.0	25.60		770.			
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2		? :	• • • •	2	1670		• :	200	200.0		
20		::	:.		27.0		3 5		7.291		
4		9	-			416.60	29	4.784	4.987	-6.123	
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2		•	42.	. 683	4.621	200	86	5.646	6.697	-6.161	
20		•		5.276	5.139	6.136	•	6.245	5.414	-0.169	
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	6	67	30.	5.276	5.254	6.921	95	9.529	5.696	-6.176	
25		20	7.	5.846	6.447	-8.161	60	9.836	0.916	-9.188	
22 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	•	5	28.	1.927	9.016	98.80	~	3.010	4.935	-0.217	
25		35		6.738	6.122	0.690	2	4.892	5.111	-P.22B	•
2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2		200	28.	8.636	9.016	-8.18-	6	9.0.0	1.0.		
4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4 4			•	0.830	9.985	9.7.6	33	1.277		6.237	
4 4 4 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9		2			10.0	100.0			200.	407.40	CUST
62 4 4 11 1 2 4 1 1 1 2 1 1 2 1 2 1 1 2 1 2							::		8 186	0/9.00	
68 6. 113 4.382 6.089 6.089 6.189 6.			:.	223		200			787	207.00	
64 64 64 64 64 64 64 64 64 64 64 64 64 6			: 4	3.767	2.497	800		5.835	6.147		
61 6. 6. 6. 6. 6. 6. 6. 6. 6. 6. 6. 6. 6.				4.113	4.382	-0.269	56	4.728	8.899	-P.371	
62 4. 5529 5529 5529 65124 5529 5529 5529 5529 5529 5529 5529 55		9	•	4.337	4.382	-0.P25	=	4.848	5,265	-6.425	-
64 6. 1 5. 10.5 4.557 0.616 1 5.796 0.504 0.506		62		5,528	2.696	-0.176	25	5.667	6,141	-A.474	-
64 2, 5,245 5,414 -7,150 5 4,093 5,485 -7, 65 3, 4,094 4,795 6,186 27 4,289 4,089 6,087 6,087 6,088 6,087 6,088 6,		63	•	5,153	4.537	9.616	-	9.88	9.384	-9.488	
0 4,000 4,000 4,000 6,000 6,000 4,000 6,00		9	3.	5.245	5.414	961.60	•	4.893	5.485		-
56 3, 4.936 4.738 8.198 27 4.243 4.828 89.		80		. 911	4.875	196.0-	90	4.339	4.007	94.948	-
67 5. 4.704 4.109 9.604 4.704 0.507 100		99	,	4.936	4.738	9.198	27	4.243	4.826		
	_	67		704		700					

FIGURE 56

LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

COMPONENT EFFECT OF EACH VARIABLE ON EACH OBSERVATION (IN UNITS OF Y)
(VARIABLES SHOERED BY THEIR RELATIVE INFLUENCE OBSERVATIONS ORDERED BY INFLUENCE OF MOST INFLUENTIAL VARIABLE) OAM MODEL RUN 3 DEP VAR 1: LNYS

,	×																																																				0.07		79.4	96.
																																																								8.50
0.	x18080	9.18	4 .	200	5 6		60.0	6.97	. 1.0	8				-9.29	8.18		6	-6.26	100		6	-0.12	6	62.0	49.0	0.03	9.95	8		5 6	. 6	0.0	9.03	. P. 20	-6.25	4	200	6	9.0	-6.32	00.	-0.17	11.	12.4					9.18	9.18	0.6.6	86.9	9.19		5	
=	x21080	9.0	6		2 6	6	90.0	10.00	10.01	4	8 6	5 6	8	40.01	46.61	40.61	8 50	0.27	20.4			49.0-	6.	4	10.01	80.0	9.0	9.28	8			6	40.00	46.4	. 40		5 6	9	0	40.01	40.01	40.0	44.	6.0				0	40.6	10.00	99.9	-8.81	40.0		5.6	-8.22
-	x19083	8.14	6			2.17	0.10	. 20	6	80.0	22.50	. 6	-8.15	6.0	9.12		.6.13	6	5 6			6	80.0	6.6	80.0	8	00.0		12.00	200	7	66.65	-0.03	19.61	6.6					-0.92		-2.17	11.00	200					-0.17	8.92	-8.21		9.40	9.	20.00	
•	x11	-0.07	20.07	8.8	2 6		94.4	-6.93	96.0	5.0	78.3	2 6	16.97	90.0-	10.01		40.01	10.0-	4.			4	6	6	6.6	9.05	9.85	80.4	10.01	2.0	6	00.0	9.36	-0.05	9.10			25.0		9.85	9.89	8.65					8 8		0.0	P.28	62.4	46.4	-8.PS	00.0	9	5 6
•	X 1 3 I	. 43	2.53	010			0.23	9.24	. 40	64.69	20.63	0 4	4.	-8.16	-9.49	9.25	4.54	.0.18	9.0		9.50	9 . 16	0.25	9.25	0.23	9.24	0.24	0.21	12. A	62.0		8.54	0,24	0.03	0.00	9.5	22.00	200	P. 23	60.00	2.25	-8.25	62.4			200		4	-0.49		9.25	9.24	. P. 23	52.00	9.0	4 6
•	×e	-9.13	6	51.0	5 6		6	8.84	-9.13	-0.13	6			-8.13	0	9.94	40.0	6.	40.			. 6	. 0	-6.13	9. 19	9.47	9.19	40.	6		. 6	0	61.6	46.0	4 6 . 9	9 6	2 6		6.10	49.0	-9.13	6.47	0.47	4.				. 6	40.0	40.0	4 6 . 8	P.84	9.		6	9 6
,	S X	-3.26		18.50			-	-0.29	9.79	0.70		10.0	62.0	-0.26	-0.29	6.11	0.11						000	0 7 8	9.11	-8.25								0.11	9.11	-	5 6	200	6	-0.29	6.79	-9.56	-8.29	-6.56					9.11	9.11	9.11	-6.29	9.11	02.0	20.0	1 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2
1	x15	. P. 29	-0.53	-3.29			-8.29	80.00	1.48	. 48	80.0	22.0-	97.1	-0.29	1.49	-0.27	-0.25	60.	0.0		200		6.50	23.29	-0.20	-0.27	-0.27	-0.50		20.01		50	-0.29	40.0-	0.19	-9.27		200	6.50	64.0	-0.28	-6.50	62.6	.00		200	200	53	-0.29	-2.20	-0.27	-0.29	-0.29	62.0	1.60	18.20
•	K 1 3	-P. 41	4.	000	000			-0.33	. 33	-7.33	-6.32	20.00	200	- 20	-9.28	-9.28	-0.26	-0.25	-0.56	50.00	200			-0.12	84.6	80.	86.	86.00	10.01			0	6	6	66.	9	2 6	. 6	90.	00.0	0.19		6.13				200	28	9.28	9.32	9.46	00.0	60.	20.0	78.8	. 6 6
ARIABLES 12	WXN.	1.18	80.0	00.	200			-0.33	5.63	63.63	9.36	200	0		9	8.98	P . 6 .	-9.25	10.01	0.4.			200		-0.57	9.10	61.6	-0.57	4	. 62		8.85	0.0	6.35	-0.47	. 2.73	2 6	0.01	0	80.00	95.5	0.0	8.50	6 6			3 6	9.24	65.4	40.0	00	61.00	0.21		98.1	0 10
:			4	·	0	n d	0	0	6	•		10 W		6	5		4	v	4	4 .	4 1	0 0		. 6	-	-	-	-	N	~		1	-		2	4 .	4	1 4	4		10	0	0		0	0 1	-	-	-	•	0	6		- (N,	1.04
0834		29	•	52	. :		30	3.6	31	23	30		7	90	22	32	1.0	=	•	9 6	2		. 5	25	28	62	•	13	4	9 :	2 4	96	25	e	~	3	· ·		99	7	2.0		10	63			2 2	52	37	58	31	42	58		2 3	425
550		-	~	n	• •			•	0	10	=	2 :			9	17	80	0	23	2	22	25	52	200	27	58	58	66	5	35	2 2	35	36	37	38	30			5	77	45	9	47		::		200	53	5.4	55	96	57	80	200		62

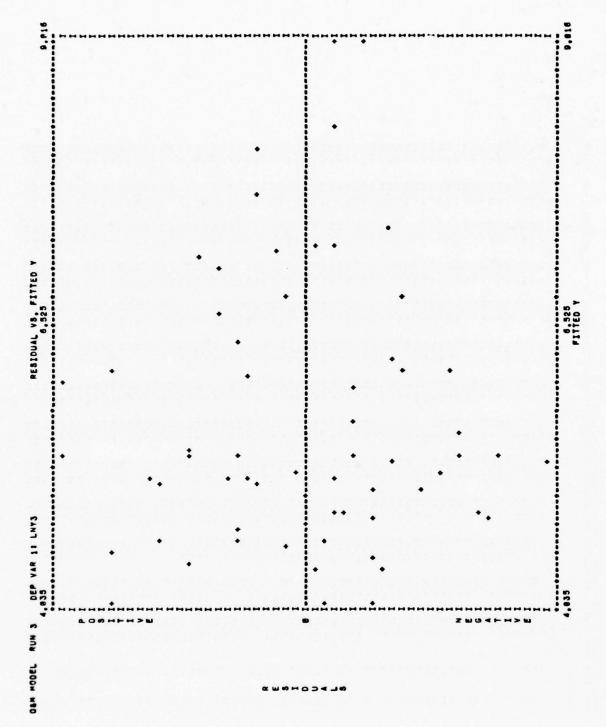
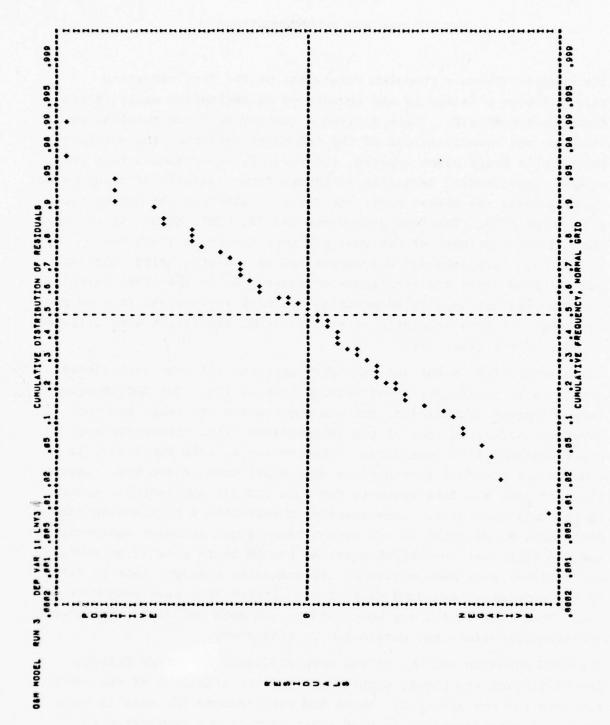


FIGURE 58



SECTION VII

CONCLUSIONS AND RECOMMENDATIONS

The results (Summary Computer Printouts) of the five remaining relationships obtained by the techniques of regression analysis are shown in Appendix B. Table 6 gives a summary of these results, indicating the transformation of the dependent variable, the multiple correlation coefficient squared, the F-VALUE, the residual root mean square, the standard deviation estimated from residuals of neighboring observations, the Normal plot, the Fitted Y plot and the observations with large WSSD. The four parameters MMH/OH, MTBF, MTBMA and LSC/OH, are the major drivers of O&M cost and were therefore given more attention. Each equation was approached as the statistics, tables, computerized plots and techniques directed. As in the MTBMA example, the WSSD was not used in determining the most appropriate form of the equation, but this statistic as well as other statistics were utilized in many of the other fits.

The MMH/OH, MTBF, MTBMA and TRAIN/OH equations all show significant results with no indication of serious lack of fit. The LSC/OH equation, although significant, had one observation, 4, which had the largest residual in each of the intermediate fits, irregardless of the functional form considered. Observation 4, with WUC 71H60, is a Gyroscope platform (navigations equipment) used in the F4E. Investigation into the data elements for this LRU did not indicate errors in the collected data. Consequently, observation 4 remained in the data base, since there is not enough known about avionics equipment and the form that the LSC/OH equation should be to deem it an outlier. The residual root mean square of .52 indicates a slight lack of fit as the cumulative standard deviation estimated from near neighbors is .57. This lack of fit may be caused by the data and other variables and transformations not considered in this study.

The NRTS equation was one of the more difficult equations fitted. The statistics are barely significant and the stability of the coefficients was not attained. Among the many reasons for this is that NRTS is highly dependent on many other factors not considered in

TABLE 6

Summary Results of the Regressions

	WSSD	47, 22	22	None	45	38, 45, 46, 48	None	
	Fitted Values Plot	OK 4	OK 2	OK N	4 high 4	OK 3	20 high N	
	Normal Plot	OK	OK	OK	4 high	OK	20 high	
	Cum. Std. Deviation	.03	.47	.40	.57	.76	7.10	
	RRMS	.03	. 42	.39	. 52	. 75	15.15	
	F-VALUE	20.5	33.5	41.5	25.3	11.7	9.9	
	$\frac{R^2}{Y}$. 9005	.9089	.9183	.9283	.8599	.8200	
	Transf. of. Dep. Var.	None	ln	ln	ln	ln	None	
•		ммн/он	MTBF	MTBMA	rsc/oh	TRAIN/OH	NRTS	
		1		1	09			

this study (most are subjective). Leaving out influential variables can make other collections of variables appear significant when in fact they are not. Although there is a serious lack of fit in the NRTS equation, 15.5 versus 7.1, the results are still useful, since only large differences in NRTS cause significant changes in the total number of spares estimated by the EBO routine.

Because of time constraints, one area not touched upon in this study is that of "prediction intervals." The prediction intervals depend on the standard error or variance of the fitted equation. The formula for computing this variance is rather complicated and is dependent on the residual mean square, the number of observations, the ith diagonal elements of the inverse matrix and the spread of the independent variables. Although the LLSCFP does not compute this variance, we recommend the simple bounds suggested by Daniel and Wood ([1], page 55).

Another area not discussed is that of error in the independent variables (Assumption A3). Some notable contributions on the subject of error in the independent variables has been made for the case of one independent variable (See Bibliography; Acton, Hocking and Leslie, Mandansky). It appears, however, that there are no results now in the statistical literature that lend to practical applications when multiple variables are considered.

As previous experience indicates, many of the logistics, cost and support parameters considered in this study are difficult parameters to estimate, especially Field MTBF, i.e., the value actually achieved. MTBF is usually the major cost and risk driver in resource, warranty and maintenance models. MTBF is estimated from MIL-STD-217B and other reliability documents based on the proposed, detailed system configuration. But the configuration and other parameters which define MTBF are not usually well defined in the early proposal phase. Previous predictions for Field MTBF in the conceptual phase were off by several orders of magnitude, which indicates the risk involved when using these predictions.

The estimating relationships obtained in this study were put through critical statistical examinations and covered a wide range of possible functional forms. Although the results, statistical and validation (Volume I), were quite encouraging, there are still areas for improvement that warrant further study to increase the prediction capability of the equations obtained.

The first and major recommendation is to expand the data base. Although the data base used indicated that relationships did exist, more data would lend to convergence of the "true" functional forms. By expanding the data base we mean more data, more independent variables, extending the ranges of the variables and expanding to newer technology areas. Other variables, not considered in this study, that may have an influential effect on the dependent variable, should be introduced into the regressions, so as to reduce bias and improve the prediction capability of the equations. Some of the variables may not have been experienced over a range adequate enough to display their influence. Extending the ranges of the variables and using newer equipment in the data base will enhance the capability of predicting advanced equipment costs.

The second recommendation is to refine the data base. This includes investigation into other Data Collection Systems to obtain more sound and up-to-date data. Moreover, it is recommended that a panel of qualified experts on the studied equipment be formed, to determine the validity of each data element.

The third recommendation is to consider more transformation of the variables. The transformations considered covered a wide range of possible forms, but there are many other transformations that may better approximate the more complicated cases. For instance, some of the independent variables were percentages which covered a wide range of values. The Inverse Sine transformation can be used to weigh more heavily the small percentages which have small variance. In addition cross products of the variables can also be considered as viable transformations. Again there must be more data available to give the analyst the flexibility needed to consider many different functional forms.

The fourth recommendation is to consider other subset collections of the final collection of variables obtained for each equation. The ${\rm C_p}$ -search technique, in addition to finding those collections of variables which have the smallest total squared error, finds other subsets of these variables and ranks them according to their ${\rm C_p}$ -values. Sometimes these subcollections have approximately the same ${\rm C_p}$ -statistic and variance of prediction as the final collection. This will greatly enhance the flexibility of the use of the equations and the ALPOS model, in that some of the values of the more difficult to obtain variables may not be needed to make satisfactory predictions.

The fifth recommendation is to investigate the possibilities of considering Non-linear Regression Analysis as a means of determining the correct functional form of the equations. Although the relationships considered in this study covered a wide range of possible functional forms, Non-linear Regression Analyses can be used to approximate even more complicated cases.

Although many Logisticians feel that predicting Logistics costs by the techniques of Regression Analysis is not a viable approach to take (usually because of inconsistencies in the data collection systems), the statistics and validation results, however, indicate the great possibilities ahead. We feel that this study has significantly contributed to the art of estimating advanced avionics equipment costs early in the conceptual phase. The stage has been set, the statistics defined, the approach outlined, the results displayed and the recommendations made. There is a light shining across the horizon. The challenge is to reach that light.

SECTION VIII

A TECHNICAL NOTE ON REGRESSION ANALYSIS

The name Regression Analysis is associated with the names Analysis of Variance and Analysis of Covariance. There is not a very acute distinction between these three types of analyses. According to Scheffé (see Bibliography), the distinction can be made that in the analysis of variance all independent variables are qualitative, in regression analysis all independent variables are quantitative, whereas in the Analysis of Covariance the independent variables are both qualitative and quantitative. According to this slight difference, we might say that the analyses presented in this report fall under the realm of analysis of covariance. Regression analysis, however, can be used to consider all three types of problems.

Analysis of variance is used to determine if significant differences exist between the means of different populations. For instance, we may want to know if there is a significant difference between the average MTBF of equipment used in fighters and that of equipment used in bomber or cargo type aircraft. If the statistics indicate that a significant difference exists, the problem is then to find which aircraft type is "more" significantly different from the "baseline." Analysis of variance depends on different methods of comparison to determine the "least" and/or "most" significant differences. Two notable procedures that have been developed are Tukey's method and Scheffe's Method for Multiple Comparisons (See Guenther). Some recent approaches have been the Least Significant Difference and Duncan's Multiple Range Test (See Adler and Roessler).

The problem of finding the "most" significant differences becomes more difficult to disentangle when more categories are used (such as the avionics area) and interactions are considered. However, if an analysis of variance problem is solved using the statistics and techniques presented on regression analysis, much more information can be obtained. The procedure is to consider indicator (independent) variables to represent the qualitative classes (with an assumed baseline) and make a fit on the dependent variable. If the statistics

are not significant (i.e., a bad fit), we can say that there is no significant differences in the means. If the statistics indicate a significant fit, then there are some classes or interactions with significantly different means. If interactions are not among the variables admitted by the C_p -search technique (usually admits those variables with largest t-values and relative influence), then the variable which causes the most significant drop in the $C_{\rm p}$ -values gives a good indication of the class which is "most" different. If interactions prove significant, another fit is to be made with different indicator variables representing the interactive classes (with an assumed interactive baseline), and use the statistics and C_p -search technique to determine which specific interactive classes are significantly different from which others. In addition, useful information can be extracted from other statistics, tables and computerized plots, such as the coefficients of the indicator variables, the "Component Effects" Table and the component-plus-residual plots.

There are also elementary statistical hypothesis type problems, such as testing the hypothesis that the mean of a certain sample is equal to a specified value against the alternative that it is not, that can be handled by the techniques of regression analysis, if the variables are properly defined (again more information can be obtained). Thus, we see the wide range of possible applications of Regression Analysis.

Many examples in standard statistics books as well as some advanced books on statistical hypothesis testing, analysis of variance and analysis of covariance, have been investigated. Using the approaches outlined, above, the results have been similar. The author is planning a paper in the near future showing the results of these investigations.

APPENDIX A

DATA USED

IN

REGRESSIONS

ALPOS MULTIPLE REGRESSION ANALYSIS DATA

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ø) å	8.80	88.	6	6	8	0.0	8	6	6.	8	6	6	00.0	8
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CARCOM	98	73.88	75.93	8.00.	. 66.47	86.98	100,00	97.00	90.00	80.	33.99	100.00	75.99.	93.68
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ALPOS MULTIPLE REGRESSION ANALYSIS DATA (CONTINUED)

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ALPOS MULTIPLE REGGESSION ANALYSIS DATA (CONTINUED)

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POWOIS	265.	8 8 8	851.	88	580.	. 69			212.		439.	335.	1398.	1300.
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×	12.00	6	6	6	25.89	6	6	6	.00	180.88	6	6	6	88.
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SENS	42.00 15.00	39.86	31.98	7.59	104.	14.98	1486.	4.00	43.79	78.59	11.00	15,02	25.23	41.63
CAPGO	1268. 186.	1511.	432.	294.	1734.	399.	358.	88 9 4	1699.	1377.	577.	567.	300	2278. 235.
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ALPOS MULTIPLE REGRESSION ANALYSIS DATA (CONTINUED)

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5	74542	156200.	5330. 1	173,73	1558.	9.156	200.	.00	6.0	6	100.00	06.	270.	2.30	61.80
6	74540	46759.	1866.	35.83	932.	9.58	18.99	66.	88.88	100.00	89.	100.00	1620.	2.38	59.88
0	74500	324003.	3658.	110.00	9.18.	0.00	10.00		100.00	6	00.0	8	400.	2.38	8
80	77500	9791.	1769.	1.00.00	946.	9.54 0.228	11.00.0	8 8 8 8	5.08	6	00.0	94.78	359.	1.30	8
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25	770CA	31698.	959.	26.40	1457.	1,19	9.99	100.00	8.89	00.	60.	100.00	145.	1.30	6.69
20	77080	835. 9.991	8282.	6878	9.00.0	10.00.0	75.00	6 6	188.88	6		6	199.	1.30	8.00
4	73580	24542.	387	99.20	529. 9.863.	1.72	10.00	98.99	12.00	6.00	00.00	98.69	. 898	8.83	6
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22	71600	7191.	1357.	36.98	1674.	1.22	000	75.98	00.0	6	25.00	97.69	256.	2.30	6.
80	63448	9865.	1128.		797.	0.71	.88	75.99.	0.00	6	25.89	97.20	150.	2.38	9.00
6	65448	14271.	377.	323.	982.	.00	17.00	75.88	00.0		8 6 N	100.00		2.30	8.9
8	63844	3845.	1682.	6.6.	1153.	8.00 9.00 9.00 9.00	.00	75.00	6	6	25.00	23.00	. 200	1.30	6.6

ALPOS MULTIPLE REGRESSION ANALYSIS DATA (CONTINUED)

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3846. 0.168	1680.	. 8. 78.	1153.	8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8 8		75.98	8	8	25.89	23.00	599.	1.30	88.
3914.	1844.	20.00	1236.	9.59 9.286	5.70	76.93	. 0 0	8.00	8.00 15.00	97.50	8	1.30	8.0
5864.	1869.	49.00	1378.	8.74 8.295		190.001		8	00.00	79.88	388.	1.30	80.6
3914.	1844.	29.99	1235. 0.825	8 . 8 . 8 . 8 . 8 . 8 . 8 . 8 . 8 . 8 .	98.	76.03	1.00	8	15.00	97.50		1.20	3.38
4883.	1680.	51.88. 55.	1186. 6.142	0.71	88. 88. 98.	75.00	8	9	25.00	23.00	58 20	1.20	80.0
19712.	1120.	13.90	790.	a.71	99.00	75.00	8	0.0	25.98	97.98	150.	1.28	9.69
3946.	1680.	. e e . e	1153.	8.69. 8.439	8 8 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9	75.00	8.99	8	25.00	23.98	6 60	1.20	6.
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APPENDIX B

RESULTS OF

THE

REGRESSION ANALYSES

LINEAR LEAST-SQUARES CURVE PITTING PROGRAM

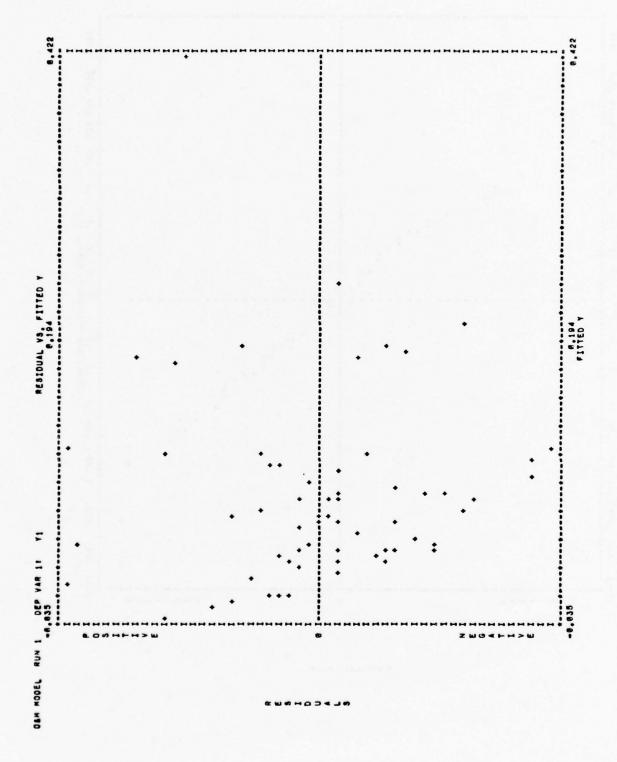
.5540-61	100		78.0		60	8.67	1.19	1.75	.4.	0.26	9.34	0.17	8.8	9.34	9.10	6,22	9.14	•	0.92	6.54							
RANGE Y .		111	7.4680-81	1000	1.9820 82			1.7250 02							1000		33	1.9980 63	5,6110 66	4.9750 86							
4.5630-61			5.3260-61	2000	1.8000 62	1.9400 82	6.7620 P3	1,3870 02	3.73PD B1	8.4490 91	9.6570 81	-			2.5810 P3	2.8730 03		1,1610 63	9.6120 68	5,1570 88							
DEP VAR 1: Y1 HUN Y B B.33GD-F4 MAX Y B 4.5G3D-81 RANGE Y B 4.5G4D-81 HULTIPLE REGRESSION ANALYSIS FOR THE "ALPOS" HODEL B(2) + P(1)X13 + B(2)X15 + B(3)X7 + B(4)X13 + B(5)X18 + B(5)X18 + B(6)X8H + B(7)X13 H + B(9)X14 H + B(9)X13 H + B(10)X15 H + B(11)X15 H + B(6 HOUR)	1114 411	-2.1340-81	20170-01		6.	-1.488D P3	-3.3480 81	-6.27eb e1	-1.6000 81	-3.4330 PB	-4.83AD AE	6.1910 84	1.6440 88	4.0000-82	5.4760 61	1.7500 83	6.2480 81	86 0164.0	1.0230-01							
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ALYSIS FOR B(3)x7 + + B(9)x15 x18057 + B	MANHOURS		a .			2.4	3.4	7.8	3.4	3.5	9.0	2.3	7.8	2.5	3.4	4.5	2.1	2.1	7:7	2.5							
MIN Y B 8,3360-64 MALYSIS FOR THE MALPOS" MODEL 19. ** # # # # # # # # # # # # # # # # # #			2.320-02	20-06-0	3.510-84	1.400-04	1.220-25	6.610-24	3.490-04	3.440-04	4.070-84	5.390-04	2.450-09	5.210-96	92-071-9	7.930-06	3,910-03	2,760-03	1,830-02	1.970-82	80		20.5	8.63915483	6.9999927	0.000000000000000000000000000000000000	.000
DEP VAR 11 HULTIPLE V1 * B(2) + P(1) * B(7) x1 * B(1) x1	Y1 . HHH/OH	1.951150-81	-4.512210-82	29-05-07-0-0	1.622920-03	-3.346750-84	-6,613690-05	4.628710-03	1.872080-03	1.200820-03	1.543680-83	1.249280-03	1.701100-08	-1.292830-45	-0.000000-00	3,568720-85	-6.386870-85	5,781680-65	7.478670-82	-4.981970-82					IAE	S S S S S S S S S S S S S S S S S S S	. SOUARED
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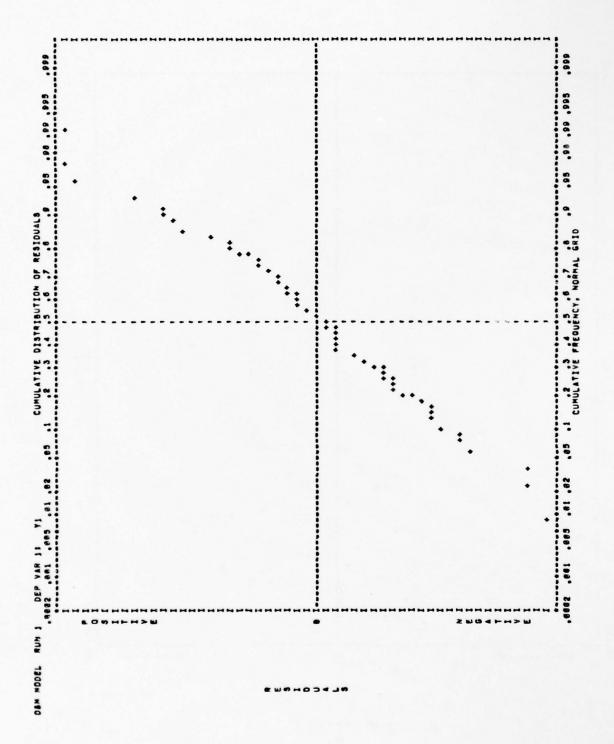
(OIGHT PECUAL POSITIVE, LEFT OF DECIMAL POSITIVE, 100, VAR(1) OIGHT 100, VAR(1) OIGH

SEG	-	CV	,	• •	n •	•	. «	•		=	12	13	-	2	24		10	28	21	25	200	23	58	27	9 0	36	31	35	2		36	2	200	7	7	7 ;	? ;	5	9	•	2 9			35	2		8	25			5	82
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FITTED Y ORDERED	B.187	-9.463	8.933	8.170	-8.93	0.00	0.170	22.0	9.756	913.6	9.139	9.782	9.650	8.196	800	500	199.87	6.423	189.8-	713.0		6.6	998	8.888	8.831	200	9.765	P. 854	6.867	200	8.951	8.971	766.6	8.9.8	6.642	9.170	6.160	0.4.0	6.827	8.198	5000	870.0	9.186	6.936	600	9.631		8.974	200	9.857	6.169	9 987
088. 4	8.168	8.88	8.838	F.224	8.693	9.143	1.211	000	8.077	500.0	8.206	6.919	B. 975	0.123	436.			6.832	8.691	8.024	6.636	8.00	8.8.8	8.883	8.834	8 6 6 6	6.653	0.052	8.065	5 6	8.6.6	6.867	50.0	100.00	400.0	0.171		4	6.811	0.174	80.00		8.168	6.015			F. 883	9.616	6.176	6.631	6.931	
0994	10	•	32	27	1	90	23	22		15		:=	63	04	8 6		200	•	16	25	7 ;				-		2	95	\$			76	7,		6.	24	200	2.4	1.3	67	a :		58	10	D		30	60	2	> ~		**
RESTOUAL	8.663	076	648.6-	8.803	-6.324	-8.83	8.938	-8.926	200	2.6	- 6	100.4	-8.818	9.735	6.633	-8.621		706.0	8.821	416.6-	.894	76.6			6.057	610.6		-8.883		6.863	3.9.50	568.84	800.0	410.61	280.61	198.	564.64	2.6	.6.628	8.63.8	9.637	760	6.663	8.638	-9.912	710.61	-6.835	190.6	28.62	6.662	504.60	
FITTED	8.631	0 0 0	55.5	6.693	6.873	190.0	-8.93	6.028	0.953	200	200.0		0.027	8.176	6.239	9.836	200	8.422	905	6.823	9.89	0.179	90.00		9.033	564.6	780	9.9.0	6.179	8.629	20.00	700.0	0.644	9.0	8.831	0.040	0.242	80 60	6.97	68.89	8.014	18.87	20.00	8.184	6.193	200	8.197	8.197	450.0	900	8.695	
COMPUTER		100		9.00	9.049	0.026	6.663	8.902	9.032			5 6	3.311	0.211	9.176				8.977	0.0.0	9.878	R. 224	0.100	200	660	5.193	100.0	2.21	0.171	8.936	5 6	500	9.958	400.0	2 6	9.636	6.237	200	. 6	6.001	A. 424	.00.0	200	8.143	469.	9.10	8.932	6.158	9.935	2 4 5	000	
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LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

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17 27 25 25 26 26 26 26 26 26 26 26 26 26 26 26 26	14	17 27 25 25 26 26 27 27 27 27 27 27 27 27 27 27 27 27 27				7.51	6.0	80.6	40	31
17. 27. 28. 28. 28. 28. 28. 28. 28. 28. 28. 28	17. 27. 28. 28. 28. 28. 28. 28. 28. 28. 28. 28	17. 27. 28. 28. 28. 28. 28. 28. 28. 28. 28. 28				55.94	9.92	80.6	62	32
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2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2				14.77	69.63	96.8	63	**
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67 17 8.62 183.57 8.76 9.76 9.18 9.18 9.18 9.18 9.10 9.10 9.10 9.10 9.10 9.10 9.10 9.10	67 17 672 15357 6.76 918 918 918 918 918 918 918 918 918 918	67 17 8.92 153.57 8.76 8.18 8.18 8.18 8.18 8.18 8.10 8.10 8.10			6.33	8.6	60.0	9.18	37	90
7 26 9,83 116.17 8.64 8.19 61 17 8.63 81.93 8.73 81.9 68 17 8.30 123.98 8.02 8.19 48 28 8.81 241.93 8.03 8.21 37 31 8.93 287.78 8.84 8.24	7 26 9.93 116.17 0.04 0.10 61 17 9.93 123.99 0.02 0.19 68 17 0.90 123.9 0.02 0.19 48 28 0.91 241.03 0.03 0.21 49 20 0.91 241.03 0.04 0.21	7 20 610 116.17 5.04 60.19 61.9 61.9 61.9 61.9 62.19 6			8.82	155.57	8.66	8.0	27	37
62 17 6.02 61.03 6.03 6.09 6.19 6.09 6.019	61 17 6.93 81.93 8.93 8.19 6.19 6.19 6.19 6.19 6.19 6.19 6.19 6	61 17 9.93 61.93 6.63 6.19 6.19 62 62 62 62 63 63 63 63 63 63 63 63 63 63 63 63 63			26.6	116.17	8.84	6.10	28	96
68 17 0.30 123.98 0.02 0.19 6.21 48 28 0.93 0.93 0.93 0.21 3.41.93 0.93 0.24 0.24	62 17 67 69 69 69 69 69 69 69 69 69 69 69 69 69	60 17 0.00 123.00 0.02 0.19 0.19 0.02 0.19 0.03 0.03 0.03 0.03 0.03 0.03 0.03 0.0						61.6	31	30
49 29 6.91 241.93 6.93 6.24 6.24 8.73 8.84 6.24	20 20 0.00 241.00 0.00 0.00 0.00 0.00 0.00 0.00 0.00	49 29 6.91 241.93 6.93 6.24 6.24 6.24 6.24 6.24 6.24 6.24 6.24							40	6.0
37 31 8.93 287.78 8.84 6.24	27 28 8.04 8.04 8.04 8.24 0.24 0.24 0.24 0.24	40 KG 60,01 CA1.00 C.00 C.00 C.00 C.00 C.00 C.00 C.00			20.00	163.30			•	
37 31 8.83 287.78 8.84 6.24	27 21 8.93 287.78 8.94	37 31 8.63 287.78 8.84 6.24 27 31 8.63 5.02 8.82 9.42			18.8	541.03	20.0			
	67 8	27 31 8,83 5,92 8,82 6,42			8.83	287.78	10.0	0.24	•	20
200										





LINEAR LEAST-BOUARES CURVE PITTING PROGRAM

			MULTIPLE REGRESSION ANALYSIS FOR THE "ALPOS" MODEL 8(a) + 8(1)x3M + 8(2)x4M + 8(3)x5 + 8(4)x6 + 8(5)x15 + 8(6)x9M + 8(7)x16M + 8(9)x90x6 + 8(9)x130x6 + 8(19)x180x0 + 8(11)x210x0 + 8(12)LNx8 + 8(13)LNx9 + 8(14)LNx19	LYSIS FOR	THE "ALPOS" + B(4)X6	HODEL B(S)X15 + B(6	ХОН		
		N 2 .	10H + 8(9) X 9 C 1 C 1 C 1 C 1 C 1 C 1 C 1 C 1 C 1 C	H + B(3) X3	+ B(19)	8(5) X15 + B(6) XOM		
			2) LNX8 + B(13) LN						
		LNY2 . LNCHT	BF) HTBF (MEAN	1 TIME BETW	EEN FAILURE	0			
IND, VAR(I)	NAME	COEF, B(T)	8.E. COEF.	T-VALUE	R(I) SORD	HIN K(I)	MAX X(I)	RANGE X(I)	REL. INF.X(I)
	XXX	3.462810-81	1.550-91	2.1	P.4524	-2.5800-41	7.4280-81	1. 5040 98	4.97
۰.	I T	-4.582430-61	1.590-01	5.6	0,3118	-2.1880-01	7.9300-81	1.6940 98	0.6.
,,,	K.X	1,153740 00	3.240-61	3.6	0.2530	-2.0330-01	5.3870-81	7.4200-01	9.17
•	×	6.347630-A1	3.270-91	6.1	0.2250	-2.1650-01	5.7350-01	7.9000-01	9.19
•	X15	1.724340-02	2,850-93		9.6568	9.6	1.0000 02		9.35
•	16×	3.791880-84	9.660-05	3.9	0.8823	-1.4540 93	6,7150 03		8.63
	KIBI	9.887450-03	2.370-03	4.2	0.6977	-6.19PD B1	3.0000 81	1.0000 82	9.20
•	2808×	-6.188970-88	2.670-08	2.3	8.7498	7.6730 84	2,3650 87		0,39
0	x13050	2.897050-04	1.180-94	9.1	P.2238	8.4540-P1	3,2440 83		9.14
10	X18080	1.883490-94	7.710-95	2.4	0.3245	7.7260 81	2,7260 03		6.19
11	X21050	-5.826640-04	3.430-84	1.7	0.2768	7.839n B1	1.1300 83		R.13
25	C vx	-2.385660-81	6.070-02	3.9	9,6335	5,6630 88	1.2690 01	7,6260 88	0.37
22	LNX9	-6.250550-01	2.160-01	2.9	0.9522	3.4810 98	9.8120 83	5,6110 88	0.72
=	LNX10	.4.698980-01	1.620-01	5.0	6.8953	1,6230-81	5.1570 64		0.47
NO. OF 140. VARIABLE RESIDUAL DEGREES OF FEVELUE ROST HEAN SCHOOL ROST RESIDUAL NUM OF SOUR RESIDUAL SUM OF SOUR RULT, CORREL, CORREL,	OBSERVATIONS TOOSE CARTIONS TO OCCREES OF THE ROOT MEEN THE MEEN SOURS TO SOURS TO SOURS	US DT FREEDOM N SQUARE DOARES	02. 03.5 03.5 04.200007 04.2000007 07.2000000 08.2000000000000000000000000000						

(DIGIT RIGHT OF DECIMAL POSITIVE, LEFT OF DECIMAL MEGATIVE) IND.VAR(1) DIGIT

LINEAR LEAST-SQUARES CURVE PITTING PROGRAM

8.42	PART IN FITTED Y ORDER).
OF FITTED EQUATIONS	(OBSERVATIONS 1 TO 4 A
RESIDUAL ROOT MEAN SQUARE OF FITTED EQUATIONS 6.42	NEIGHBORING ORSERVATIONS
	STANDARD, DEVIATION ESTINATED FROM RESIDUALS OF NEIGHBORING DESERVATIONS (OBSERVATIONS 1 TO 4 APART IN FITTED Y ORDER).
DEM HODEL RUN 2 DEP VAR 11 LNY2	STANDARD, DEVIATION

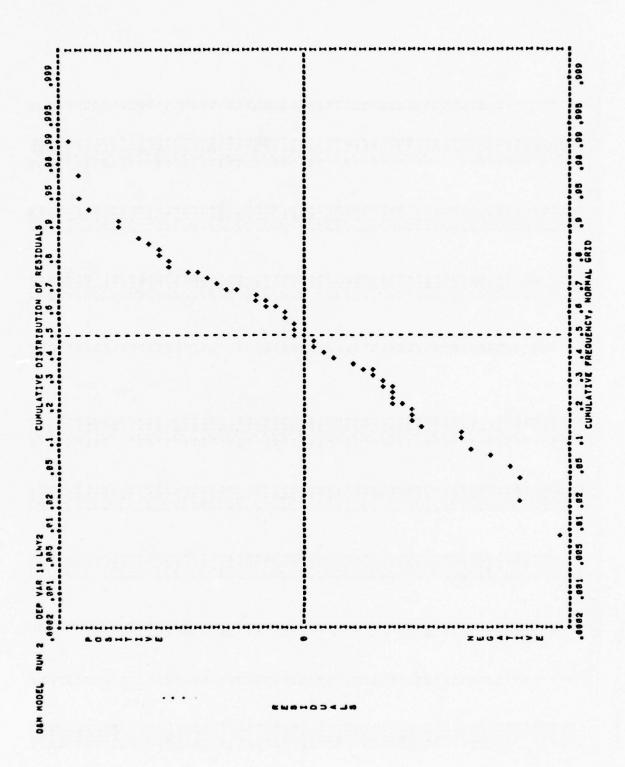
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: 8	55	92	67	•			•	5	? ~	•			: 2		33	37	36	88	56	42	œ	,	67	•	8	9	2	8	₹,	• ;	3 *	. 46		6	62	•	8	80	- ;		25	35	3.4	=	::	60			900	2	90	?	•		• :		2.
FITTED Y	4.30	4.35	4.38	4.59	4.30										5.27	5.28	5.28	5,29	5,37	5.43	67.0	9.38	3.33	8.63		9.71	5.73	5.75	3.76	2.0	0 · ·			8	9.00	60.0	6.03	46.0	9.12	6.43		6.53	6.78	6.79	9.0	9.00			7.20	7.44	7.48	7.50	18.0	6.33	20.0		
RESIDUALS								9.50	44.0							9.52	9.39	0.10	9.84	6.83	88.8	0.11	4.52	9.88	9.04	9.47	4.24	76.0	1.22	20.0	8.52			9	97.6	6.52	8.43	00.	9.00		51.1	10.6	9.84	9.42	6.53	9.10	8.33	2.6		9.57	11.0	6.36	6.63	76.0	200	20.00	20.0
130 DEL	•	99	86		3.43	60.		900							2.91	6.6	32.69	96.61	12,55	4.77	3.84	2.23	11.68	27.44	10.01	7.17	5,42	13.88	80.0	3.03	2.83	20.00		36.20	16.65	29,48	9.49	1.00	2.69		8.32	2.77	19,65	13.51	25.59	~	9.0	75.57	:`	-	25.48	-	1.78	~	12.68	n	
ESIDUALS	. 42	9.52	9.21	07.0	.34		2.5								76	•	19-1	1.29	9.69	6.47	61.0	0.05	H.57	6.55	6.32	76.0	0.11	0.42	20.03	9.10			24.0		17.6	8.99	10.6	8.68				8.48	8.75	6.29	10.0	0.27	1.22		20.0	6.37	6.19	8.71	9.95	6.97	0.25	10.0	0,0
DEL	_	_	_	_	_																																																				
RED BY 0634.											•	2 6	•	: `	27	9	37	•	89	_	_	5	4	4	•	•	7	2	2	2	200	, 4	2	? :	2	2	3	9	0	2			70	9	•	96	• :	::		42	42	•	^	3	? :	2:	?
OBSV.	23	25	91	53	5	9	86	8.	2 -	-:	2			2 4	-	75		•	67	67	63	93	7	33	2	•	70	2	-	5		-	- :		69	80	35	8	2	3*	, ;	~	2	63	4	C	=:		- E	21	90	~	60	•	200	•	2
1000	0.0	0.0	8.8	80.0	10.0		0.25	9.53		2		2 .		7.			183	1.98	1.90	1.99	1.98	1.90	1.96	1.99	2.03	8.25	2.23	2.27	2.47	2.63	200		20.0	2.75	2.75	2.78	2.77	2.86	2.0			30.0	3.86	3.11	3.18	3.23	80.0	20.0		3.48	3.48	3.53	3.64	3.66	000	00.0	
STO DEV	8.62	9.28	8.25	. 8.32	. 0.32	80.0	4.32	6.50		07.0	***	200	000			9.52	6.55	6.39	69.9	0.50	. 0.57	6.59	6.55	6.55	9.54	.52	16.91	16.91	16.8	9.28				27.0	. 0.47	0.40	6.47	8.47	4.0				6.43	. 0.45	. e.45	6.45	4.47	4.6		6.45	6.63	8,45	6.45	9.46			20.0
											=:	.:	2:			::			50	21	22	2	24	52	50	22	5	50		5	25	3 2	;;	200	34	38	30	•	7:	2:	:			11	•	9			7 5		23	90	21	80			-

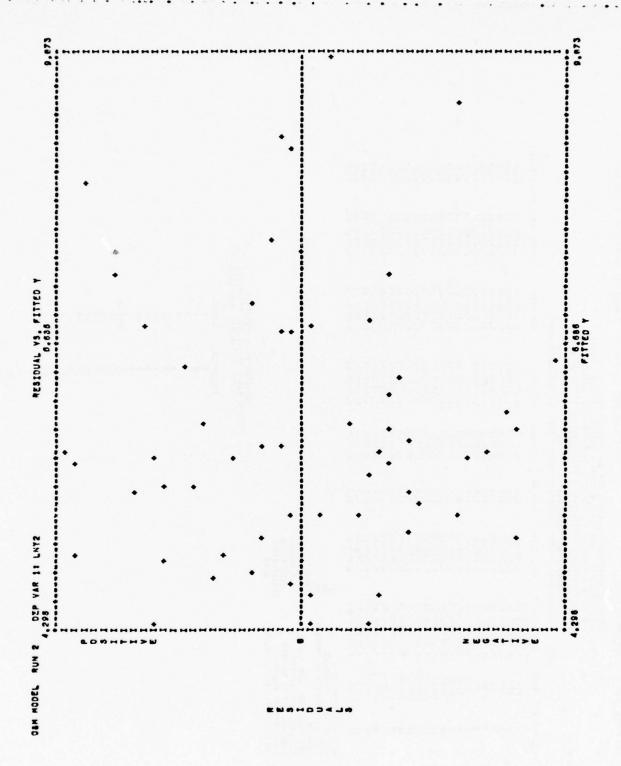
LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

DEP VAR 11 LNY2

DEM MODEL RUN 2

ESID. SEG		0.683							-	-		2	•	-	-	113		20	15 22	92 23	55	20	27	.949 20	. 827 29	8.918 38	35	142 33	143	25	37	.186 38	42 39	193	P.5	107 43	32		19	65	64	24 51	26 52	34		479	92 57	564 58	628 59	
T ORDERED H	7.0	6	6			6	4.0	4.6	4.6	4.				9.2	6.8	9.21					406.0			6	6		200.00	P. 042	-9.943		67.677	- 4-		20.00	-8.293	-0.207	-0.232		- 6.5	P. 25	18.87	-6.32	.P.32	50.00			-6.49	6.0	0.6-	
FITTED	5.779	5.685	800		7.280	5.483	6.837	5.747	4.383	2.497	266.	200	90.0	4.736	4.948	5.738	000.	2000	5.678	7,562	6.788		4.718	5.282	6.761	466.8		4.352	6.861	5.278	0.978	0.073	6.828	20.50	6.891	4.298	4.588	3.714	3.986	6.251	1.208	808.8	3.449	5.141	2000	5.282	5.756	9,816	6.117	
088. Y	6,510	6.369	5.684	0.00	7.853	5.992	7,322	6.199	4.836	5,922	2.315		6.322	4.998	5.177	3.943		5.078	5,193	7.664	6.852		4.751	5.322	0.00.0	6.352	200.7	4.389	6.616	5.226	6.993	8.973	9.89	5.116	6.688	4.891	4.337		5.722	5.987		9.000	5,123	7.00.1		200.4	5.264	5,252	9.496	
088V.	•	50	n (36	~	;	30	29		200	20	9 6	•	17	2;			9	?	=:		1.1	9.	76	- ;	0 0		50	200	200	31	2	9 0	5	55	6 5	. 4	29	5	0.0	4	42	6	2 6	96	=	•	-	
RESIDUAL	-0.629	9.300	6.673		600	9.9.6	9.055	-9.198	-8.283	90.00	612.8		8.254	-9.467	8,568	-8.287	200	19.30	-8.669	-0.183	8.663	20.00		0.425	0.285	786.6	20.00	6.563	8.663	20.80	0.102	6.483	-8.263	00.00	9,316	-0.324	66	9.366	. 6.97	0.027	20.00		-8.654	6,432		.0.204	A.42B	-0.243	-4.842	
FITTED	6.117	5.483	4.038		2,179	8.334	8.398	5.626	6.891	6.788	9.1.9	5.0.7	7.736	8.684	7,443	4.298	0000	3.372	5.190	5.292	500.5		5.278	5.497	6.638	7.456	100	7.289	8.610	0.7.0	7.562	6.837	6.251	20.0	5,532	5,893	000	9 9 9	9.673	192.9	0.00	5.282	5.257	5.747		3.000	4.803	5,782	4.352	
088° Y	5.496	5.005	9000		8.50	8.352	8.445	5.428	6.698	6,852		101.	000	8.217	8.011	100.4		200	4.431	5.119	6.369	2.5	5.226	5.922	6.322	7.452	2000	7.853	8.675	2000	7.664	7.322	5.987	5.177	9.93	5.560	6.897	9999	8.995	6.800		5.322	5.202	6.199		5.722	5,312	5,539	998.	
HSS DISTANCE			•	• •		26.	22.	3.	12.	•	•••	•	33.	45.	20.	28.	•	•		33.	.:	•			÷.	<i>:</i> •		::	.01				÷:		3.	31.	• •		25.		•		2.				5.	•		
0837	-	~		• •			•	•	6	=:	2			20	21	22		28	27	28	2 2	200	::	*	33	9 :	38	30	2:	- 3	3	::	•	•	•	9 :	e -	25	93	3:		33	86	200		. 2	63	*	6	
TOENT.	71820	73538	711.00	000	71699	71719	72460	71659	717 43	2011	2011	7110	73084	73680	73CEN	73654	1 1 1 1 1 1	73696	71549	72EAA	72804	21749	71649	72048	72400	1710	7246	72488	13938	1486	74810	6419	706AP	84.44	BH 47 4	74500	77660	17007	77083	73543	2 1 N A B	1000	83443	8446	83644	65844	61884	658AA	83468	





6.352D 68		REL, INF. X (1)	8.40	77.6	200	9.10	4.87	9.13	4.27		10.	200	75.0	87.4	9.28	8,93	9.11	1.86	8,22	2.42	01.0						
RANGE Y . 0.		RANGE X(I)	1.26.00 88	000000000000000000000000000000000000000	7 4690-84	7.9430-81	7,9400-01	1.0000 02			2 0044			2.7450 83				1.7640 88	1.6780 83	99 OBCO. /	2,0110 00						
2,0400	5	HAX X(1)	7.1480-61	7.3000-01	1 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	5.6690-91	5.7960-91	1. POND P2	3.6790 81	3.8990 61	6.6880-81	3.0230 16	2 6 0 0 0 0	2. 8400 93	3.5760 03		6.4270 05	_	_		8.6120 68						
NG PROGRAM	SXXO M	MIN X(I)	-2.8680-91	-2.7880-81	-2.54m0-m1	-2.2710-01	-2.144D-P1	6.6	-6,3380 A1	-6.1190 01	-1.3490 98	5.1930 MB	2000000	4990	2,3100 62		2,2500 02	1.4590-01	6,6230 01	5.6380 98	3.4410 68						
MARAN LEADING CORRES TO THE TANKE DROGGES TO A MINITED TO SECOND SECTION SECTI	THE "ALPOS"	R(I) 80RD	0.9483		9.6622	2000	8 000	P. 4192	8.7676	0.6928	6.9327	69.69	8.328Z	0/00-0	97.6	9.4652	6.3924	0.8716	0.4010	9.7431	P.732B						
A81-50UARE	12 X 19 B C 3 X 18 B C	T-VALUE	8.8	8.8	o .			.0	8.9	3.8	9.5	8.	5.0				8.7		3.6	9.0	.2						
	AN II LNY4 HULTIPLE REGRESSION ANALYSIS FOR THE "ALPOS" HODEL 8(2) + 8(1)XIR + 8(2)XXR + 8(3)XXR + 8(4)XX + 8(1)XXR + 8(1)XX	9.E. COEF.	6.410-01	6.530-01	2.600-91	4 740-01	1000	4.930-93	3,750-03	2.500-03	4.580-01	1.840-11	5.250-43	000.1	20000	1080-04	4.290-07	7.670-81	4.780-84	0.280-02	1.080-01	63		25,3	8.27342179	11,21029323	
	DEF VAR 11 LNYA NULTIPLE R NULTIPLE R NULTIPLE R NULTIPLE R NULTIPLE R H 0 (1) 0	<u>=</u>	3.861110 89	3.665330 88	-4.852710-01	-2.566630 88	19 020200 1	1.273560-02	2.259670-92	-7.429990-03	2,385030 00	-9.263840-11	-1.52864D-84	-1.6710=0-13	20010000	10-01-01-01-01-01-01-01-01-01-01-01-01-0	98-089	5. 930cAD 99	1.798420-03	4.602930-01	2,355830-01	92	F FREEDOM		PE SOUANE	UARES	
		NA N	*	×2×	# 3 H	en :			1 4 1 %	X1.9H	X 28 H	x8080	x19089	X14080	00000	200	0 % CO . X	×290.80	x21050	LNXB	C X N	OF OBSERVATIONS	EGREES		EAN SOU	UN OF 30	
	1	IND.VARCE)	• •	. ~	•	•	n •	D .			19	==	12	-	=:			::		36		NO. OF OBS	RESIDUAL DEGREES OF FREEDOM	F-VALUE	REGIOUAL WOOT MEAN SQUAME	RESIDUAL SUM OF SQUARES	

REQUIRED X(1) PRECISION (DIGIT RIGHT OF DECIMAL POSITIVE, LEFT OF DECIMAL NEGATIVE) THO.VAR(1)

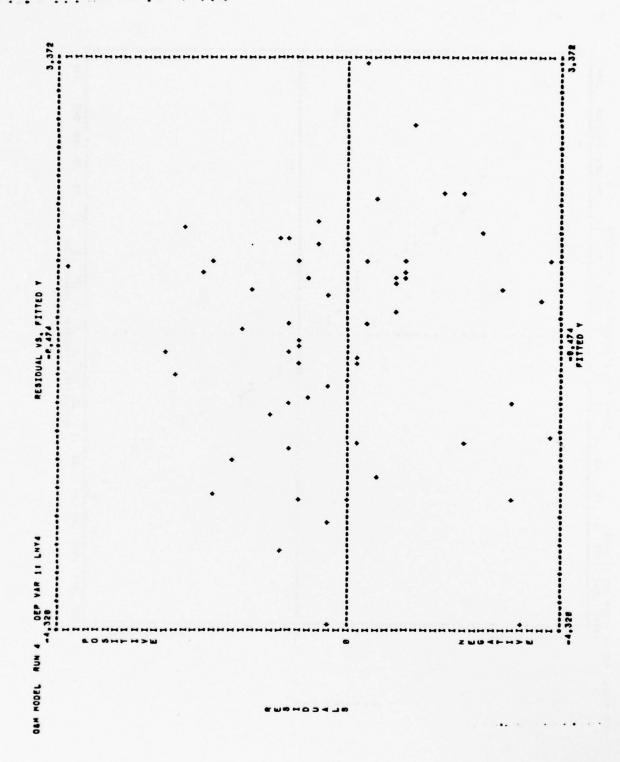
LINEAR LEAST-SOUARES CURVE FITTING PROGRAM

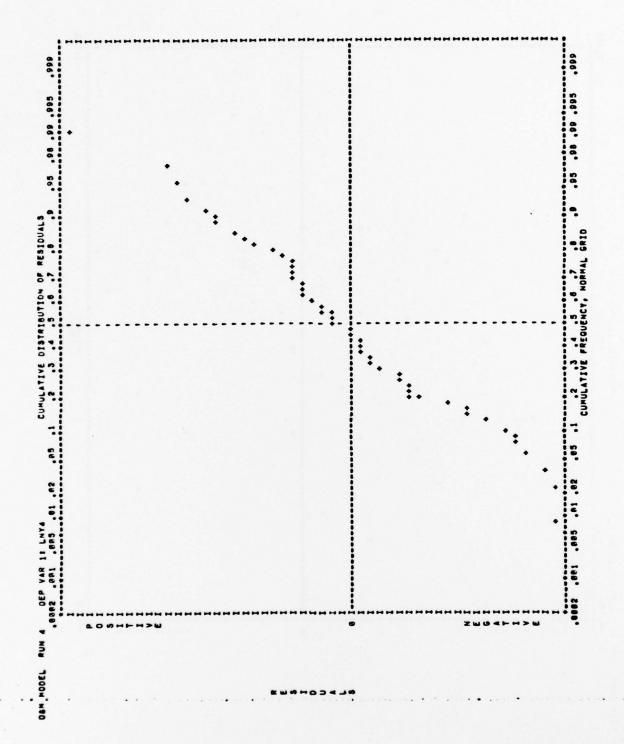
APART IN FITTED T ORDERS																																																	
RT IN FI		SEO.	-	~ .		•	•	•	•		=:	2 -	=	£ .	2:		-	58	22	53	3.5	2 2	2.2	88	2 5	50	35	3 3		38	2	90	::	: 3	? ;		9	•	•	8.	2 2	33	3	en en	9	86	a (5
TO 4 APA		0837	53	16	- ;	5.0	• •	2 9	9	•	2:	60	2	9	2 5		99	•	2 2	;-	=	N 0	•	8 6	2 5	=	7	n ;	8	58	e .	.5	?:	2.5	9 6		86	À .	-	" ;	\$2	6	5	5 6	. 6	4:	•	,	4.1
COBSERVATIONS 1	,	FITTED Y	-4.32	.4.32	25.00	-2.55	.2.33	2.43	-2.23	-1.97			-1.74	-1.78		91.19	-1.16	80.6	0 40	-0.72	12.00	000	-8.54	-8.58	200	-0.22	-6.16			0.23		8.42	7.7		e. e	.0.0	99.0		1.7.	2.71	20.0	0.0	1.02	1.92	==		70.1		
		DEL RESIDUALS	.88	50.	22.		00.	0.00	63	.25	15.		. 83	.28		200		•		24	25	250	33	96.		33	.33		. 6.	.33	0 17	85	90			27	61	37		0.0	28	20	22	36	57	88			
OBSERVATIONS																50															- 6		8 6	•	5 6				.1		2	6	•••						
NEIGHBORING		UALS ×S		188	121	6			60	113,	29	87.	136,	30		143	100	80	98	20.	128	32.	123,	23.		. 69	7.5		661	166	126.	119.16	127.7	-		23		103.7	104.3	329		1001	0	2 4	157.5	107.7			110.0
9		DEL RESIDUALS	0.21	2.4.	900	0.50	8.58	00.0	9.17		1.0		6.00	1.0		0.12	1.37	7.5	000	6.83	0.12		10.5		200	6.28		11.17	6.87	60.0		1.83	1.27	0.12	90.0	1.01	3.5				9.37	6.20			8.93	6.6			0.0
RESIDUALS	4	. NS 8	6	0 .	• en	27	9 6	, •	8	,,	? \$, -	67	2 3	55	-	9	a ;	2	58	÷ :	9	53	55	22	33	6.5	33		•	50	8 9	. 4	6 :	2 6	90	ê°		• !	? ?	2	6	27	3	9	2 9			
ESTIMATED FROM	000	088V	5	2	3 6	2	0 6		•	25	00	38	2	25	9	•	•	- :	. •	99	~ ~	2	33	, 0	0	70	2 5	6	*	<u>.</u>		2,		•		•	• 0	30	2	- 2	2	= "	D 10	67	-;	9 6	3		
		WS 80		. 6		8.95	6.17	1.52	2.84	2.55	2.27	2.37	2.44	2.43	3.33	3.37	3.83	6.3	3.66	6.99	90.0	7.27	7.94	0.53	9.38	11,51	27.	13.55	16.98	19.91	21.15	23.12	23.31	23.71	23.63	24.23	24.23	25,25	25,35	27.69	27.59	30.72	32.66	33,61	35.43	35.86	37.46		
STANDARD DEVIATION	THE ATTVE	310 054	6.23		6.73	.0.	000		6.33	20.00		6.63	8.65		6.62	6.30	6.99		9.0	19.9	0.0	6.94	6.97	6,00	6.56	6.55		6.57	6.55		90.0	20.0	6.97	9.36	6.57	6.39	60.0	6.63		6.63	6.52					800			
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LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

DEM MODEL RUN 4 DEP VAR 11 LNY4

1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1, 1,			× 041414	ALC: SAN	0834	000			
1, 10, 10, 10, 10, 10, 10, 10, 10, 10,	24.		-P.720	9.198	•	1.883		1.192	-
1, 10, 10, 10, 10, 10, 10, 10, 10, 10,		366	8.293	.9.661	86	8.266		9.766	~
1.02	24.	-0.174	9.714	-9.997	32	-8.195		6.735	0
2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2		. 89.5	6.611	1.192	69	1.015		4.69.	•
2.237	24.	9.199	-0.113	9,222	88	1.135		4.684	•
2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2	25.	-9.08	-8.081	6.881	26	-1.896	-2,453	8.537	•
1, 200	31.	-2.918	-3.189	0.298	13	1.253	8.786	9.547	•
2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2	26.	-3.237	-2.529	-A.708	9.	-1.479	-1.972	567.6	
2. 324	.92	-P.659	-0.621	-6.938	7	0.201	6.223		• :
2007	33.	-0.749	-9.711	000.60	3	6.000	25.6	10.0	-
7, 200 1, 200	29.	-2.547	-2.528	0.00	2	819.10	200	2000	::
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1. 201 1. 201	.86	-1.479		20.0		80.0		217	-
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1	45.	2,185	~	-11.01	5	166.50	66.70		
1. 100	. 69	P.268	9.459	-4.201	-	-8.525	-0.720	8.1.0	2
2	.67	-1.895	-1.192	20.79	62	-0.238	-6.427		2
2		9.287	0.230	0.937	26	198.0	9.682	6.178	22
1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1. 1		1.27	8.085	9.286	9	-0.192	-6.378	0.177	23
1.139 1.139	•		17.4	-	99	-1.831	-1.161	9.138	2
10.12.00 10.	::	000			2	205	8.467	9.139	25
1.000 1.000	•	1				34	1.222	8.119	28
1.0771 1.0772 1.0773 1.	25.	207.6	2000	414	; >		80.0	9.038	27
1.3224 1.		00.00	1.8.	200	•	000		9.00	28
10.000 10.0		1.139	1.57			22.75	20.00	796	58
11. 20.0	54.	1.366	1,524	001.00		1000	810	A. 957	
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10.000 10.0000 10.000 10.000 10.000 10.000 10.000 10.000 10.000 10.000 10.0	• 66	015.6	081.1.	222.5	. •			800	
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10.000 10.0000 10.0		100.00	2000	767 8	•	000	-0.621	. P. 938	30
1	•	1000	22.0			740	-8.711	68.69	6
1. 20.000 1. 20.0000 1. 20.0000 1. 20.000		802. W.	ACT	1100		3 4 3	740	. 9.942	3
10.000 10.0000 10.0000 10.0000 10.0000 10.0000 10.0000 10.0000 10.0000 10.0000 10.00		000		100	2 2	726		-9.062	30
10.020 10.020		102.20			: :		-1.788	486	;
10.040 10	204.	9.813	D. B. B.	A	2	700.1			
2	=:	9.320	20.80	12.4	90		3.72	600	•
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20.000 20.000 20.00000000	.3.	H. 184	B. 835	102080	2 :	007.00			•
200	. 66	788.6	8.878	900		100.34	7000		•
2		4.200	880.84	00/.				20.00	•
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1.2022 9.202 9.202 9.202 9.203 9.203 9.203 9.203 9.203 9.203 9.203 1	22.	92/00		200.0		2000	57.6	-A. 237	7
7		/2.	900		2.6	430	6.682	-0.242	4
2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2. 2		16/02	2000	700			215	.0.244	3
7.400 7.505 7.	• 60	302.1	00/01-		3		6.435	-0.251	•
7.000 1 1.000		100.	200.0	242	22	2.185	2.562	-0.317	6
1.231 1.231	•		7 6	5	33	1.139	1.577	-8.438	ń
1.221 1.921 1.921 1.921 1.023			500	107.6	:	-2.281	-1.776	-8.584	n
1.010 1.021 2.004 07 2.005 1.070 1.0		500.	100	8.218	•	181	1.628	-0.519	•
10.000 1 10.			1.921	100.00	69	0.472	1.079	-0.697	ñ
9.103 0.24 0.17 0.177 0.177 0.177 0.177 0.177 0.177 0.177 0.171 0.		87.0	-0.427	081.6	~	P. 366	6.295	-8.661	6
2.102	•		607	510.85	50	.1.895	-1.192	.9.783	ñ
1.015 1.121 0.094 01 -5,000 -4,020 -9,000 0.0268 0.034		000		44.		-3.237	-2.329	-0.798	ň
0.266 8.17 68.24 R0 69.17 8.134 68.134 7.266 8.247 8.34 8.34 8.34 8.34 8.34 8.34 8.34 8.34	.2.	261.00				0 0 0	4.320	-0.730	9
000.11 000.01 000 000.01 000.01 000.01		010.1	1.161			21.4	134		•
	•	9000				280		-6.871	
		2/15	8/4-1	140.4	80	2000			





LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

TEGUIRED X(1) PRECISION

TREAT OF DECIMAL POSITIVE,

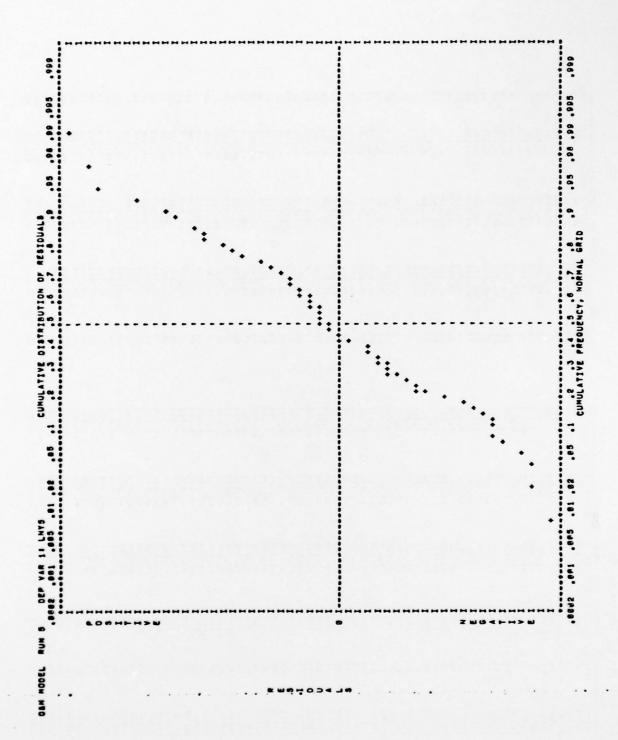
THOUGH OF DECIMAL NEGATIVE)

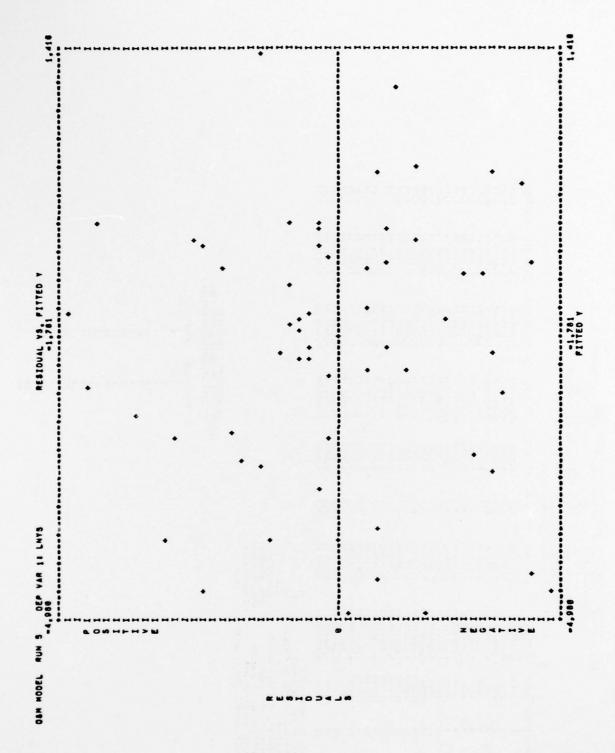
LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

	\$E0.	-	~ ~	. •	10	•			. 5	=	12	2	1:			•	9	9 .		23	7.	25	520	200	88	9,		32	**	32	000	36	30		. 2	?	::	. 9		7		91	25	26		300	25	28	20		
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FITTED V	FITTED Y	-4.98		-4.62							-3.31		97.5	2.03	-2.89	-2.88	-2.64	-2.41	20.00	-2.17	-2.15	-2.02	-2.69	10.10	16.1-			.1.63	-1.60	-1.57	26.1.	-1.43	91.10			-8.97	0.0		-0.75	-8.74	69.	-0.67	-8.38	.0.22		. 4.47	96.99	9.87	6.0		
A CONTRACTOR OF THE PROPERTY O	DEL RESIDUALS	6.49	6.77			1.93	80.0	8.32		6.63	1,23	9.11	1.25	70.0	0.56	9.52	1.98	2.21		8.22	6.33	4	5.03	200		9.25	200	.0.0	6.63	1.25	1.32	9.40	1.00	20.12		1.64			6.03	80.0	1.17	6.53	9.30	96.	A	1.26	10.0	99.0	6.43	• • •	
	.330	9.0	12.14	2.99	6.62	7.78	22.13	200	25.49	56.23	94,26		1143.73	11.69	496.47	556.87	789.78	4.19	1481.74	1624.03	1532,84	685.25	9	1547.73	1512,08	1268.68	22.87	79.83	9.17	5.79	31.58	78.75	9.28	20.000	0.31	36.35	1334.20	379.83	4.65		1.35	1385,01	1483.27	900	9 5	6.74	9.14	184.46	40.70		
	DEL RESIDU	9.13	6.0		6.83	1:1	1.00	5.		9.14	0.12	8.18	9.0	0.00	9.01	9.68	2.86	6.97		9.37	8.89	1.36	9.70	9.74	6.97	6.97	9 6	1.10	96.98	6.63	100.0	1.74	1.03		0.57	1.25	2.08	11.1	1.26	9.00	24.0	9.38	6.45	B 6	20.0	90.1	1.30	9.62	62.8		
E0 87	. AS80	69		. 2	89	•	90	6		12	25	2	200	26	-	25	2	•			62	•	- ;	200	9	25		•	96	6	2	8	6:	•	88	50	2 :		•	5	200	9	2:	250	0 0	50	35	8			
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	4550		0.0							1.78	2.25	2.72	2.63	00.0	3.14	3.15	3.23	3.23		20.0	3.63	3.83	3.97					.63	4.65		7.7	10.4	3.10	5.23	3,39	5.79	56.6	6.74	6.74	7.55	7.97	9.91	.0.	10.0	9.46	9.29	9.30	99.0	4/.01		
CUMULATIVE	STO 0EV	6.17	2.5	8.38	9.38	8.69	9.00			6.03	9.79	9.74		50.5	90.0	6.65	6.92				6.00	86.8	60.0		.0.	9.0		16.9	96.9	6.67		86.8	6.0		6.87	9.00	6.92	6.92	6.93	8.02		6.03	8.05	60.00	9.05	6.03	.0.0	50.6	24.0		62
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OLM MODEL RUN S DEP VAR 11 LNYS

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10	1820	-		-1.427	-1.652	0,225	28	-6.965	-1.531	1.486	-
1	11.88	,	. 25.	.1.	.0.078	-1.012	32	-1.031	-2.350	1.319	~
11	601	• 1	418.	8.012	-0.469	1.282	•	6.812	469	1.282	n
1				1.000	-0.674	-8.412	= :	-1.545	-2.635	1.001	•
	90.			200	606.	200.1-	2:	30.00	693.	000	0
1	2469			200	200.	87.0	:6	24.6	200	45.00	•
100	. 649			22 . 74	20.01	0.0	•		619	200	•
2	1649		. 56	3.878	3.124	13.734	98	9.0	9.739	60.0	•
1	1689	:=	288	-1.343	-2,635	1.691	12	-6.362	. B. 972	6.619	
15. 10. 10. 10. 10. 10. 10. 10. 10. 10. 10	1549	13	146.	-5.487	4.437	-1,831	•	-2.302	-2,877	9.575	:
1	1043	2	97.	-1.805	-2.003	861.6	3	-2.691	-3,182	16.591	15
	1485			-2.875	-2.892	9.6.	8	-2.866	-3.261	200.4	13
	100	2		-8.362	-8.972	8.0.6		•	1.418	0.379	= :
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20	3585	50	146.	-8.935	-1.188	9,253	13	-1.805	-2.003	0.100	55
20	1048	27	.88	8.065	-8.591	9,736	47	-	-2,015	9.158	23
1.77 1.79	2511	58	.96	-9.003	-1,531	1.468	25		-1.918	0.150	54
10.00	ZECA	50	147.	-1.893	-1.003	-0. A.D	3	-1.366	-1.513	0.147	53
10	526			1.797		9.379	25	-8.636	-8.754	9.116	50
100		25	136.	100.10	2000	1:31		107.10	900	6.119	200
10.00	2002	300	200	7.00	:	0.6	2 2	27.8		200.6	000
144 1144 1144 1144 1144 1144 1144 1144	2469	33	.36.	-2.159	-1.003	-0.265	97	P. 658	-6.739	. 800	80
11.00 11	1716	36	145.	.5.703	-4.624	-1.169	•	-2.174	-2.233	6.628	31
144	220	2	. 55	-0.638	-8.734	9.110	2:	-8.813	.6.653	8.638	35
11. 1	203	200		2000	-3.261	66.60	22	-2.873	-2.692	0.00	3
### ### ### ### ### ### ### ### ### ##	2033			2.0		101		7. 7.0			;
### ### ### ### ### ### ### ### ### ##	4809	-	62.	100	8.977	878.60		20.00	860.4	4.6	
### ### ### ### ### ### ### ### ### ##	0 161		923.	-1.182	-6.916	-0.266	65	-1.596	-1.433	-8.147	200
### ### ### ### ### ### ### ### ### ##	8197	?	862.	-2,661	-3,182	6.501	2	-2.327	-2.154	-0.173	3.8
111.12. 4.0 4.0 4.0 4.0 4.0 4.0 4.0 4	8118	::	143.	-2.076	-2.933	6.857	69	-1.221	-1.027	-9.193	30
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665. 665. 666. 677.	E H 4		1132.	20.00	2.159	507		200		200	**
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### ### ### ### ### ### ### ### ### ##	EC.	20	.00	-1.479	-1.691	0,212	67	-0.623	975-0-	-0.275	43
11.000 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	0 2 2	16	. 998	40.0		-8.83	55	8.701	1.928	-9.327	9
000 000 000 000 000 000 000 000 000 00		**		000.1		190.0			201.20	00.00	•
99 99 12 22 11 1 1 2 2 3 1	200	7		79	9 4 9	0 6		200	64.5	20.00	•
94 94 94 94 94 94 94 94 94 94 94 94 94 9	8536	98	663.	-1.769	-1.918	8.158	25	-5.438	4.988	-6.458	
97 997 11 1 00 0 10 0 10 0 10 0 10 0 10		96	57.		-6.739	6.692	98	-2,211	-1.691	-6.618	31
10 12.2.11	800	37	57.	-0.658	-6.739	9.989	4	-1.755	-1.869	-8.485	35
0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0 0.0	6776			-2.211	-1.691		6 1	-3.678	-3.124	-8.754	93
62 500 100 100 100 100 100 100 100 100 100	44.		• 00	2000	0.00	20.0		26.	200	9.618	0
62 46. 1.988 1.147 6.193 6.193 6.194 6.195 6.196 6.197	3644			9.375		208		45.1.50			0 6
63 1491.221 -1.727 -6.193 23 -53,378 -2.413 -6 64 451.356 -1.513 7.147 3 -1.998 -9.78 -1 66 361.358 -1.571 0.217 12 -5.487 -1 67 971.971 0.217 12 -5.487 -1.57 67 971.971 0.217 12 -5.487 -1.57 68 971.972 -1.973 -1.972 -1.973	2844	95	. 9	-1.588	-1.433	-6.147	=	-8.793	8.077	87	32
64 45, -1,356 -1,513 P,147 3 -1,998 -9,978 -1,678 -	1881	63	149.	-1.221	-1.627	.0.193	52	-3.300	-2.413	-0.887	96
00 00. 10.000 11.001 00.017 120 50.600 00.00	2844			-1.366	-1.513	6.147	•	-1.898	-8.878	-1.912	30
TOTAL NAME OF CALLES				2	1,5,1	0.217	~ •	-5.487	4.457	-1.031	9
				20.00	0.0	6/7.80	• ;	24.	904.91	1.032	6





	20000		REL.INF.X(1)														20.0																															
	MANGE 4 1		RANGE X(1)	1.99.0 88	1.0000 90	1.8980 68	7.5860-11	1.0000	6.6690 99	1.8800 82	1.0000 02	20 0000	2.3740 67	2.1830 67	18 0010 1	2.5500 93	2.9940 A3	3.2230 83	1.7680 88	1.0760 93	7.6580 88	5.0110.6	5.4570 00									ve,																
	1.4880 82		MAX X(1)		7.2660-91	7.5860-81	5.5630-91	1.0000 92	6.7250 83	0.2320 91	3,6500 91	16 0504.8	2.3690 67	2.1870 97	1,7350 01	2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2 2	2.9480 93	3,4810 93	20100.0	1.1430 93	1,2690 #1	9.6120 88	8.0440								707 81 43 60 11	DECIMAL POSITI	1 01611		-	- 0	:?	- •	~ 9		1	••	-	??	?	? ?		•
ING PROGRAM	F X 4 4 6 . E	MODEL 6(3)x16 + 6(6)x9M (15)x13080 + 6(19)x19080 + 6(23)L4x9	HIN K(I)		2 7 4 8 0 - 8 1	-2 423D-81	-2.8770-81	6.6	-1.4450 83	18-0100	10 000	-1.1170 01	1.3580 88	4.7520 84	9.3650-83	8.4640-01	2.5860-61	2,5600 82	2.8920 82	1.4/30-61	8.0300	3.4810 99	2.1970 00	1.9460 00								COIGIT RIGHT OF DECIMAL POSITIVE,	IND.VAR(I)	- 0	••	•	n e	•	••	• •	:=	25	2 2		-	•	: 2	•
S CURVE FITTING	. > 111	THE "ALPOS" + 8(4)x5 + 8(10)x17W + 4)x1205Q + 8 6(10)x17D5Q + 8(22)LWXB	0808113080	2000	1880.0	8.9002	8.3412	A.7738	9,9317	2000	6.50	0.7276	0.9528	1001.0	6.7423	0.9230		8.979.8	6.5611	6.9463	6.3013	9.9768	6.9963	6.6782																								
LEAST-SQUARES		1 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	911 140	-													2.0																															
LINEAR LEA	9,	HULTIPLE REGRESSION ANALYSIS FOR THE "LIPOS" HODE. (69) + P(1)xiH + B(2)xXH + B(3)xSH + B(1)xSH + B(1)xXH + B(1)XH + B(1																						2.540 88	25	25		15.14933888	224,58222307	45919.9691935																		
	0EP VAR 11			COEF.B(I)	6.354140 81	4.136390 91	-1.455300 21	2,787270 81	1.525170 48	1. 000 001	-1.41 FORD PR	-1.652970 80	3,114550 81	-2.919790-66	3,129250-06	3.171.50	5.327760-62	-3.357580-62	-5.156210-02	1.941890 92	-9.8AØ560-02	6.981490 99	6.344820 81	6.636610 79	87	BLES COFFEEDIN	חו באבנותה	N SOUARE	IARE	TOTAL SUM OF SQUARES	F. SOUARED																	
	8 NO.			NAME	* . *	101	I O X	s ×		1	I SI X	X 1 4 M	X178	0806×	x11050	X12090	X14050	x15050	x17050	3608CX	x21030	LNXB	Ox N	LAXII	BSERVATIO	NO. VARIA	DEGREES	ROOT MEA	HEAN SOL	100 10 H	PREL. CO																	
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LINEAR LEAST-SQUARES CURVE FITTING PROGRAM

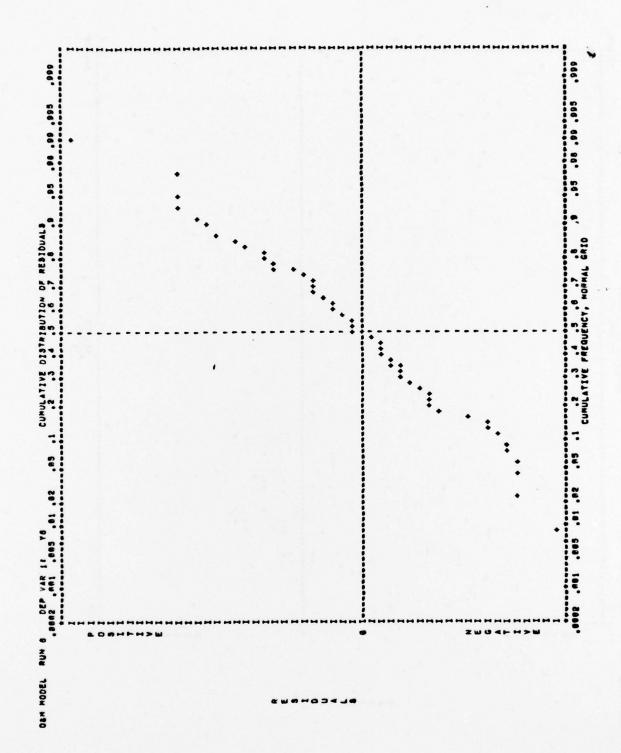
Y ORDER).		
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OF FITTED EQUATIONS (OBSERVATIONS 1 TO		86.98
8 8	w O	3.04
RESIDUAL ROOT MEAN SQUARE NEIGHBORING OBSERVATIONS	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	360,61
RESIDUAL ROOTS OF NEIGHBORING		3.04
RESIDUAL	8 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	13
11 Y6		20
DEP VAR 1: Y6 ESTIMATED FROM RESIDUAL	$\begin{array}{c} a & a & b & a & b & a & b & a & b & a & b & a & b & b$	88.78
STANDARD DEVIATION	0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0	
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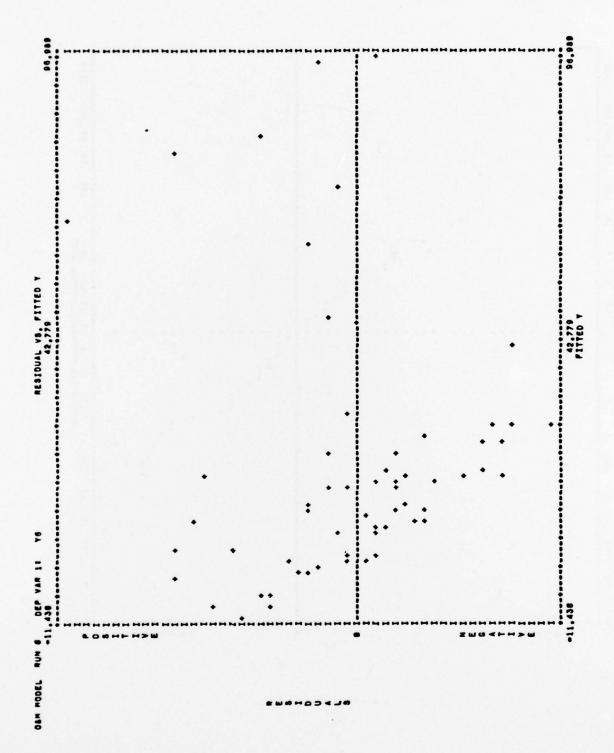
4.6

DEP VAR 11

9 NOSEL RUN 6

SEO	-	~	•	•	•	••		• •		:=	12	2	=	2	= :	2:		:		23	23	2	25	50	2			3 5	32	33	3	6			30	;	= :		:	•	•	::	; ;		91	25	20		96	33	:	30	:	5 6
ORDERED RESID.	34,288	21,253	20,837	20,686	18,621	17.234	20.01	13.439	11.306	10.676	16,365	16,055	7,214	6.471	5,643	1/1.6	0000		3.7.5	- F	3,131	2,683	2,453	2,345	1.376	000.0	9 6		-1.345	-1.443	-2.186	100.20	20.00	-2.989	-3.156	-3.593	4.798	4.973	-5.171	-5.938	-6.367	200.40	-6.322	.6.453	.6.453	-9.769	-13.126	19.10	-16.407	-17.482	-17.758	-18.214	-10.402	-18.969
FITTED Y	65.868	2.747	79,163	-2,686	0.379	16.766	246	-11.438	5.376	82,324	-7.345	-5,855	6.666	-1,471	19.357	620.10	701.70	110.00	200.00	21.442	14.869	47.317	72.547	6.655	1.224	20.132		100.00	1.345	9.443	2.146	100.0	870 51	96.96	7.956	10.595	21.798	19.373	15.171	16.939	12.967	200.4	84.322	18.453	19.453	16.466	17.128	10.10	26.407	17,482	23,758	42.814	27.402	26.968
OPS. Y	100.001	24.886	186.088	16.000	27.848	34.099	. 7 . 2 . 2	88.0	6.8	93.93	3,988	5.884	8.199	21000	16.009	67, 888	864.	900	5 6 6	860	18.08	50.000	75.898	996.0	2.668	55.50	986	46.998		9.88		200.	966	94.49	4.668	15.000	17.090	2.4	10.00	11.000	5.700	900	16.888	2.698	2.000	6.740	999.7	9 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6 6	10.00		6.000	84.688	0.99	880.
0884	59	56	12	54	•	÷,		35			22	7	67	37	80	66	3:				•	9	66	20	80	6	* 3			7	1:	3:	À -	:=	96	•	2 2	. 69	36	2	25		•		19	6	95	2.5	•	13	8	=	90	62
RESIDUAL	-2,948	0.756	18.621	-6,322	4.917	11,366	10.01	-18.214	-2.080	20.637	-5.938	-17,482	4.911	-15.215	-15,107	34,200	66.00	00000	900	21.053	4.268	5,643	-17,758	17.234	13,438	2000		-13.128	6.471	-2,661	-18.482	266.7	66.01	10.676	-2,196	8.00	2.683	3.131	-16.407	2,345	4.700	6.53	3.171	-3.136	-2,756	-0.766	14.137		-6.367	-4.973	9.590	1,376	-7.113	12.
FITTED Y	15.948	15,244	9.379	24,322	15.917	-5.386	8 C . C C	42.214	000	79.163	16.938	17.482	-0.811	23,215	18.197	63.808	200	000	900.70	2 7 7 7	-8.268	10.337	23,750	16.766	-11.438	12.102	20.00	17.128	-1.471	6.661	27.482	260.	20.40	92.324	2.186	96.07	47.317	690.41	26.487	6.635	21.798	14.19.	61.00	7.936	7.956	16.466	3.043	20.01	12.967	16.373	8.500	1.224	6.913	
OBS. Y	13.429	16.90	27.999	16,999	11.999	6.666		24.999	85.6	196.000	11.999		604.7	8.069	3.939	888.88	900		5 6 6	808 70	666	16.888	6.498	34.469	2.400	99.7		806.7	5.866	4.00	606.0			93.666	6.4	96.07	6.6		19.99	0.00	17.99		67.000	4.888	9.238	6.784	17.208		5.786	3.400	3.600	2.649	1.884	
MSS DISTANCE	26.	57.	35.	76.	37.	24.	.212.	. 62		95.	.68	58.	61.	155.	161.	284.			• • • • • • • • • • • • • • • • • • • •			150.	136.	116.		•			. 97	177.	•	. 661	. 20	120	38	307.	268.	187	189.	45.			113.		.97	32.				67.	43.		35.	
0834	-	~	•	•	•	•	•	•	::		2	13	12	•	6.	50	25	*			22	58	58	31	35	2		90	35	38	8		2	12	:	\$	2:	:	•	8	25	200	60	96	97	80	86		95	63	:	69	9	6
IDENT	71829	73536	711.80	21469	71PK9	71988	2070	71649	71583	71549	71049	71 ABE	71404	7308A	7116	73080	7305	1000	73684	TIERE	71549	72544	72ECA	72673	71748	71049	7245	71716	71310	72860	72888	8060	7486	74818	78110	76649	74578	746 118	74500	77EC9	77004	23589	73060	71MAR	11000	83448	65446	4364	65FAA	79919	44859	8316	63446	63121





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